

DISCUSSION PAPER SERIES

No. 4746

INTERNATIONAL RISK SHARING AND THE TRANSMISSION OF PRODUCTIVITY SHOCKS

Giancarlo Corsetti, Luca Dedola
and Sylvain Leduc

INTERNATIONAL MACROECONOMICS



Centre for **E**conomic **P**olicy **R**esearch

www.cepr.org

Available online at:

www.cepr.org/pubs/dps/DP4746.asp

INTERNATIONAL RISK SHARING AND THE TRANSMISSION OF PRODUCTIVITY SHOCKS

Giancarlo Corsetti, European University Institute and CEPR
Luca Dedola, European Central Bank (ECB)
Sylvain Leduc, Board of Governors of the Federal Reserve System

Discussion Paper No. 4746
November 2004

Centre for Economic Policy Research
90–98 Goswell Rd, London EC1V 7RR, UK
Tel: (44 20) 7878 2900, Fax: (44 20) 7878 2999
Email: cepr@cepr.org, Website: www.cepr.org

This Discussion Paper is issued under the auspices of the Centre's research programme in **INTERNATIONAL MACROECONOMICS**. Any opinions expressed here are those of the author(s) and not those of the Centre for Economic Policy Research. Research disseminated by CEPR may include views on policy, but the Centre itself takes no institutional policy positions.

The Centre for Economic Policy Research was established in 1983 as a private educational charity, to promote independent analysis and public discussion of open economies and the relations among them. It is pluralist and non-partisan, bringing economic research to bear on the analysis of medium- and long-run policy questions. Institutional (core) finance for the Centre has been provided through major grants from the Economic and Social Research Council, under which an ESRC Resource Centre operates within CEPR; the Esmée Fairbairn Charitable Trust; and the Bank of England. These organizations do not give prior review to the Centre's publications, nor do they necessarily endorse the views expressed therein.

These Discussion Papers often represent preliminary or incomplete work, circulated to encourage discussion and comment. Citation and use of such a paper should take account of its provisional character.

Copyright: Giancarlo Corsetti, Luca Dedola and Sylvain Leduc

ABSTRACT

International Risk Sharing and the Transmission of Productivity Shocks*

A central puzzle in international finance is that real exchange rates are volatile and, in stark contradiction to efficient risk sharing, negatively correlated with cross-country consumption ratios. This Paper shows that a standard international business cycle model with incomplete asset markets augmented with distribution services can account quantitatively for these properties of real exchange rates. Distribution services, intensive in local inputs, drive a wedge between producer and consumer prices, thus lowering the impact of terms-of-trade changes on optimal agents' decisions. This reduces the price elasticity of tradables separately from assumptions on preferences. Two very different patterns of the international transmission of positive technology shocks generate the observed degree of risk sharing: one associated with improving, the other with deteriorating terms of trade and real exchange rate. In both cases, large equilibrium swings in international relative prices magnify consumption risk due to country-specific shock, running counter to risk sharing. Suggestive evidence on the effect of productivity changes in US manufacturing is found in support of the first transmission pattern, questioning the presumption that terms-of-trade movements in response to supply shocks invariably foster international risk pooling.

JEL Classification: F32, F33 and F41

Keywords: consumption-real exchange rate correlation puzzle, distribution cost and incomplete asset markets

Giancarlo Corsetti
Robert Schuman Centre
European University Institute
Via dei Roccettini 9
San Domenico di Fiesole (FI) I 50016
ITALY
Tel: (39 055) 468 5760
Fax: (39 055) 468 5770/776
Email: giancarlo.corsetti@iue.it

Luca Dedola
General Economic Research Division
European Central Bank
66-68 Neue Mainzer St.
D - 60311 Frankfurt
GERMANY
Tel: (49 69) 1344 8668
Fax: (49 69) 1344 8553
Email: luca.dedola@ecb.int

Sylvain Leduc
Board of Governors of the Federal
Reserve System
20th and C Streets NW, Stop 23
Washington DC 20551
USA
Tel: (1 215) 574 6440
Fax: (1 215) 574 4364
Email: sylvain.leduc@phil.frb.org

For further Discussion Papers by this author see:
www.cepr.org/pubs/new-dps/dplist.asp?authorid=148438

*We thank our discussants Larry Christiano, Mick Devereux, Fabrizio Perri, Cedric Tille, and VV Chari, Marty Eichenbaum, Peter Ireland, Paolo Pesenti, Morten Ravn, Sergio Rebelo, Stephanie Schmitt-Grohe, Alan Stockman, Martin Uribe, along with seminar participants at the AEA meetings, the SED meetings, Boston College, the Canadian Macro Study Group, Duke University, the Ente Einaudi, the European Central Bank, the European University Institute, the Federal Reserve Bank of San Francisco, IGIER, the IMF, New York University, Northwestern University, the University of Pennsylvania, the University of Rochester, the University of Toulouse, the Wharton Macro Lunch group, and the workshop 'Exchange rates, Prices and the International Transmission Mechanism' hosted by the Bank of Italy, for many helpful comments and criticism. Corsetti's work on this Paper is part of a research network on 'The Analysis of International Capital Markets: Understanding Europe's Role in the Global Economy', funded by the European Commission under the Research Training Network Programme (Contract No. HPRN-CT-1999-00067). Part of Dedola's work on this Paper was carried out while he was visiting the Department of Economics of the University of Pennsylvania, whose hospitality is gratefully acknowledged. The views expressed here are those of the authors and do not necessarily reflect the positions of the ECB, the Board of Governors of the Federal Reserve System, or any other institution with which the authors are affiliated.

Submitted 02 November 2004

1 Introduction

Is consumption risk optimally hedged across countries? Despite the development of international financial markets in the last decades, the answer from a large body of financial and macroeconomic research appears to be “no.”¹ While the literature has analyzed many different facets of (the lack of) international risk sharing, a crucial testable implication is that, in a world economy characterized by large deviations from purchasing power parity, domestic households should consume more when their consumption basket is relatively cheap.² As first shown by Backus and Smith [1993], this is clearly at odds with the data. For most OECD countries, the correlation between relative consumption and the real exchange rate (i.e., the relative price of consumption across countries) is generally low, and even negative. A striking illustration of such finding is presented in Figure 1, plotting (the log of) quarterly U.S. consumption relative to the other OECD countries and the U.S. real trade-weighted exchange rate in the period 1973-2001. The swings in the dollar in real terms are not associated with movements of the consumption ratio in the same direction; on the contrary the two variables tend to comove negatively.

An obvious element in the explanation is that international financial markets are not developed enough. Yet, theory offers convincing arguments to doubt that incomplete asset markets *per se* be sufficient to bring models in line with the Backus-Smith evidence on (the lack of) risk sharing. Several contributions have shown that the equilibrium allocation in economies that only trade in international, uncontingent bonds may be quite close to the first best (e.g. see Baxter and Crucini [1995]). Indeed, trade in bonds insures that the real rate of currency depreciation and the growth rate of relative consumption are highly and positively correlated in expectations — although not necessarily period by period *ex-post*, as is the case when markets are complete. In addition, Cole and Obstfeld [1991] has called attention on the role of movements in the terms of trade in potentially insuring against production risk independently of trade in assets.

But do the observed swings in international relative prices necessarily foster risk sharing — up to the point of partly offsetting the frictions and limitations of the asset markets? Could the Backus-Smith evidence instead indicate that in equilibrium these swings magnify (rather than reduce) consumption risk due to country specific shocks? In this paper, we study the link between high exchange rate volatility (the exchange rate volatility puzzle) and international consumption

¹See the surveys by Lewis [1999] and Obstfeld and Rogoff [2001].

²As discussed in the next section, under standard assumptions on the utility function this is the main implication of efficient risk-sharing in the presence of real exchange rate (PPP) fluctuations — as opposed to a high cross-country correlation of consumption. See also Engel’s [2000] discussion of Obstfeld and Rogoff [2001].

risk sharing (the Backus-Smith puzzle) in a two-country model of international business cycles. We adopt a model with traded and non traded goods similar to Stockman and Tesar [1995], except that we assume incomplete asset markets. Moreover, as in Burstein, Neves and Rebelo [2003], we introduce distribution services produced with the intensive use of local inputs. Combined with standard preferences across tradables, distribution services contribute to generate a realistically low price elasticity of tradables. In this setting, the terms of trade and the real exchange rate are highly volatile in response to productivity shocks and have large, uninsurable effects on relative wealth. Specifically, we show that large movements in international prices can actually hinder risk sharing, making the set of international assets available to agents less effective as instruments to hedge consumption risk against country-specific shocks. We conclude our paper by complementing our quantitative analysis of the model with statistical evidence on the responses of the real exchange rate and the terms of trade to innovations in productivity in the U.S. economy.

The theoretical and quantitative analysis in this paper yields two novel and important results. First, when we calibrate our model to match the U.S. real exchange rate volatility, we find that it generates large departures from efficient risk sharing. The predicted correlation between the real exchange rate and relative consumption is negative, while the comovements in aggregates across countries are broadly in line with those in the data. The main predictions of the model are reasonably robust to extensive sensitivity analysis.

Second, depending on the value of the price elasticity of tradables, our model predicts a low degree of risk-sharing for two very different patterns of the international transmission of productivity shocks, each corresponding to a plausible set of parameters values for preferences and technology. In our benchmark calibration, for a price elasticity slightly above $1/2$, international spillovers in equilibrium are large and positive. A *positive transmission* is a standard prediction of the international business cycle literature: an increase in the productivity of the domestic tradable sector leads to a deterioration of the terms of trade and a depreciation of the real exchange rate. However, in our baseline economy the deterioration is so large on impact that relative domestic wealth decreases, driving foreign consumption above domestic consumption. For a price elasticity slightly below $1/2$, instead, international spillovers are still large but — strikingly — negative. With a *negative transmission*, following a productivity increase, the home terms of trade and the real exchange rate appreciate, reducing relative wealth and consumption abroad. In both cases, large swings in international relative prices run counter to efficient risk sharing. But our numerical results suggest that the overall performance of the model is best under the negative transmission mechanism.

The latter pattern of international transmission is due to a combination of an unconventionally sloped demand curve, and nontrivial general equilibrium effects arising from market incompleteness. Because of home bias in consumption, domestic tradables are mainly demanded by domestic households. With a low price elasticity, a terms-of-trade depreciation that reduces domestic wealth relative to the rest of the world would actually result in a drop of the world demand for domestic goods — the negative wealth effect in the home country would more than offset any global positive substitution and wealth effect. Therefore, for the world markets to clear, a larger supply of domestic tradables must be matched by an increase in their relative price, that is, an appreciation of the terms of trade — driving up domestic wealth and demand.

To investigate whether the international transmission of productivity shocks to tradables in the U.S. data bear any resemblance to the above patterns, we close our paper with some suggestive evidence. Two findings stand out. First, we provide novel evidence in support of the prediction of a negative conditional correlation between relative consumption and the real exchange rate. Following a shock that increases permanently U.S. labor productivity in manufacturing (our measure of tradables) relative to the rest of the world, U.S. relative output and consumption increase, while the real exchange rate appreciates.³ Second, the same increase in productivity improves the terms of trade, as suggested by our model under the negative transmission.

In light of the results in this paper, the Backus-Smith evidence appears less puzzling yet more consequential for the construction of open-economy general-equilibrium models, with potentially strong implications for welfare and policy analysis. In fact, if the international transmission mechanism is such that a positive shock to productivity translates into a higher, rather than lower, international price of exports, foreign consumers will be negatively affected. Terms-of-trade movements will not contribute at all to consumption risk-sharing. Thus gains from international portfolio diversification may well be large relative to the predictions of standard open-economy models.

The paper is organized as follows. After providing a brief summary of the evidence on the correlation between relative consumption and the real exchange rate for industrialized countries, in the following section we derive some key implications for the link between these two variables in standard two-goods open-economy models. In Section 3 we introduce the model, whose calibration is presented in Section 4. Section 5 explores the quantitative predictions of the model in numerical

³Conditional on a productivity increase in tradables, an appreciation of the real exchange rate and an increase in domestic consumption are also predicted by the Balassa-Samuelson model with no terms-of-trade effect (because of perfect substitutability of domestic and foreign tradables). Yet, as shown by our numerical experiments, a model with a high price elasticity of tradables cannot generate either enough volatility of the real exchange rate and terms of trade or replicate the negative Backus-Smith unconditional correlation.

experiments. Section 6 presents suggestive evidence on the effects of shocks to productivity in the open economy. Section 7 summarizes and qualifies the paper’s results, suggesting directions for further research.

2 International consumption risk-sharing: reconsidering the Backus-Smith puzzle

In this section, we first restate the Backus and Smith [1993] puzzle, looking at the data for most OECD countries. Second, we reconsider the general equilibrium link between relative consumption and the real exchange rate in the framework of a simple endowment economy with incomplete markets and tradable goods only, in the spirit of Cole and Obstfeld [1991]. The goal of our exercise is to provide an intuitive yet analytical account of the determinants of the comovements between the real exchange rate and relative consumption conditional on endowment (supply) shocks. Using our framework, we will show that the link between these variables can have either sign depending on the price elasticity of tradables: a low elasticity can generate the negative pattern observed in the data. But since a low price elasticity also means that quantities are not very sensitive to price movements, a negative correlation between the real exchange rate and relative consumption will be associated with a high volatility of the real exchange rate and the terms of trade relative to fundamentals and other endogenous macroeconomic variables — in accord with an important set of stylized facts of the international economy. These results shed light on the main mechanisms driving our quantitative results in the second part of our paper.

2.1 Stating the puzzle

As pointed out by Backus and Smith [1993], an internationally efficient allocation implies that the marginal utility of consumption, weighted by the real exchange rate, should be equalized across countries:

$$\frac{P_t^*}{P_t} U_{c,t} = U_{c^*,t}^*, \tag{1}$$

where the real exchange rate (RER) is customarily defined as the ratio of foreign (P_t^*) to domestic (P_t) price level, expressed in the same currency units (via the nominal exchange rate), $U_{c,t}$ ($U_{c^*,t}^*$) denotes the marginal utility of consumption, and C_t and C_t^* denote domestic and foreign consumption, respectively. Intuitively, a benevolent social planner would allocate consumption across countries such that the marginal benefits from an extra unit of foreign consumption equal

its marginal costs, given by the domestic marginal utility of consumption times the real exchange rate $\frac{P_t^*}{P_t}$, i.e., the relative price of C_t^* in terms of C_t .

If a complete set of state-contingent securities is available, the above condition holds in a decentralized equilibrium independently of trade frictions and goods market imperfections (including shipping and trade costs, as well as sticky prices or wages) that can cause large deviations from the law of one price and purchasing power parity (PPP). It is only when PPP holds (i.e., $RER = 1$) that efficient risk-sharing implies equalization of the *ex-post* marginal utility of consumption — corresponding to the simple notion that complete markets imply a high cross-country correlation of consumption.

Under the additional assumption that agents have preferences represented by a time-separable, constant-relative-risk-aversion utility function of the form $\frac{C^{1-\sigma} - 1}{1-\sigma}$, with $\sigma > 0$, (1) translates into a condition on the correlation between the (logarithm of the) ratio of domestic to foreign consumption and the (logarithm of the) real exchange rate.⁴ Against the hypothesis of perfect risk-sharing, many empirical studies have found this correlation to be significantly below one, or even negative (in addition to Backus and Smith [1993], see for instance Kollman [1995] and Ravn [2001]).

Table 1 reports the correlation between real exchange rates and relative consumption for OECD countries relative to the U.S. and to an aggregate of the OECD countries, respectively. Since we use annual data, we report the correlations for both the HP-filtered and first-differenced series. As shown in the table, real exchange rates and relative consumption are negatively correlated for most OECD countries. The highest correlation is as low as 0.53 (Switzerland vis-à-vis the rest of the OECD countries), and most correlations are in fact negative — the median of the table entries in the first two columns are -0.30 and -0.27, respectively.

Consistent with other studies, Table 1 presents strong *prima facie* evidence at odds with open-economy models with a complete set of state-contingent securities. Given that debt and equity trade, the most transparent means of consumption-smoothing, are far less operative across borders than within them, a natural first step to account for the apparent lack of risk-sharing is to assume that financial assets exist only on a limited number of securities. Restricting the set of assets that agents can use to hedge country-specific risk breaks the tight link between real exchange rates and the marginal utility of consumption implied by (1). It should therefore be an essential feature of models trying to account for the stylized facts summarized in Table 1.

⁴Lewis [1996] rejects nonseparability of preferences between consumption and leisure as an empirical explanation of the low correlation of consumption across countries.

2.2 Into the puzzle: supply shocks, international transmission and risk sharing in a simple endowment economy

2.2.1 International transmission and the volatility of relative prices

Building on a simple setting similar to Cole and Obstfeld [1991], we now analyze the Backus-Smith correlation in a two-country, two-good endowment economy under the extreme case of financial autarky. We will refer to the two countries as ‘Home’ and ‘Foreign’, denoted H and F. For the Home representative consumer, consumption is given by the following CES aggregator

$$C = C_T = \left[a_H^{1-\rho} C_H^\rho + a_F^{1-\rho} C_F^\rho \right]^{\frac{1}{\rho}}, \quad \rho < 1, \quad (2)$$

where $C_{H,t}$ ($C_{F,t}$) is the domestic consumption of Home (Foreign) produced good, a_H is the share of the domestically produced good in the Home consumption expenditure, a_F is the corresponding share of imported goods, with $a_F = 1 - a_H$. Let $P_{H,t}$ ($P_{F,t}$) denote the price of the Home (Foreign) good, and $\tau = \frac{P_F}{P_H}$ the terms of trade, the relative price of Foreign goods in terms of Home goods. Therefore, an increase in τ implies a depreciation of the terms of trade. The consumption-based price index P is

$$P = P_T = \left[a_H P_H^{\frac{\rho}{\rho-1}} + (1 - a_H) P_F^{\frac{\rho}{\rho-1}} \right]^{\frac{\rho-1}{\rho}}. \quad (3)$$

Let Y_H denote Home (tradable) output. In financial autarky, consumption expenditure has to equal current income, i.e., $\frac{PC}{P_H} = Y_H$. Domestic demand for Home goods can then be written:

$$C_H = a_H \left(\frac{P_H}{P} \right)^{-\omega} C = \frac{a_H}{a_H + (1 - a_H) \tau^{1-\omega}} Y_H$$

where the demand’s price elasticity coincides with the elasticity of substitution across the two goods, $\omega = (1 - \rho)^{-1}$. Analogous expressions can be derived for the Foreign country.

What is the link between international relative prices and the world demand implied by this simple general equilibrium model? Taking the derivative of C_H with respect to τ :

$$\frac{\partial C_H}{\partial \tau} = \underbrace{\frac{\omega}{[a_H + (1 - a_H) \tau^{1-\omega}]^2} Y_H}_{SE} - \underbrace{\frac{a_H (1 - a_H) \tau^{-\omega}}{[a_H + (1 - a_H) \tau^{1-\omega}]^2} Y_H}_{IE} > 0 \iff \omega > 1, \quad (4)$$

makes it clear that the Home demand for the Home good C_H can be either increasing or decreasing in the terms of trade τ , depending on ω . When $\omega > 1$, a fall in the relative price of the domestic tradable — an increase in τ — will raise its domestic demand. This is the case when the positive substitution effect (SE) from lower prices is larger in absolute value than the negative income effect

(*IE*) from a lower valuation of Y_H .⁵ Conversely, when $\omega < 1$ the negative income effect will more than offset the substitution effect. Thus, a terms-of-trade depreciation will reduce the domestic demand for the Home tradable. The foreign demand for Home tradables C_H^* , instead, will always be increasing in τ : independently of ω , the substitution and income effects in this case are both positive.⁶

As long as the negative income effects in the Home country is not too strong, the world demand for Home goods $C_H + C_H^*$ will be decreasing in their relative price, i.e. increasing in τ .⁷ For ω sufficiently high, then, the equilibrium Home terms of trade needs to depreciate in response to a positive shock to Home output Y_H . The international transmission through terms of trade adjustment is therefore *positive*: foreign consumption of Home tradables will rise, responding to the fall in the relative price of imports. However, when ω is sufficiently below 1 and the Home bias in consumption is sufficiently high (i.e., a_H is large relative to a_H^*), the response of world demand for the Home goods to relative price movements will be dominated by the strong negative income effects of its domestic component: world demand will be falling in τ . In other words, the negative income effect of worsening terms of trade on Home demand will more than offset any positive substitution effects worldwide and income effects abroad.⁸ For a positive supply shock to Y_H to be matched by an increase in world demand for the Home goods, the Home terms of trade needs to *appreciate*. The international transmission in this case is *negative*: a positive domestic supply shock has a negative impact on consumption and welfare abroad.

To analyze the relation between international transmission and price volatility, we take a log-linear approximation of the market clearing condition for Home tradables ($Y_H = C_H + C_H^*$) around a symmetric equilibrium (with $a_H = 1 - a_H^*$ and $Y_H = Y_F^*$). The equilibrium link between relative

⁵Formally, by a straightforward derivation of the Slutsky equation, the substitution effect is obtained from the compensated demand function x_H :

$$\frac{\partial x_H}{\partial \tau} = \omega \frac{a_H (1 - a_H) \tau^{-\omega}}{[a_H + (1 - a_H) \tau^{1-\omega}]^2} Y_H.$$

⁶Using self-explanatory notation:

$$\frac{\partial C_H^*}{\partial \tau} = \underbrace{\omega (1 - a_H^*) \tau^{1-\omega} \frac{a_H^*}{[(1 - a_H^*) \tau^{1-\omega} + a_H^*]^2} Y_F^*}_{SE} + \underbrace{+ a_H^* \frac{a_H^*}{[(1 - a_H^*) \tau^{1-\omega} + a_H^*]^2} Y_F^*}_{IE} > 0.$$

⁷We are grateful to Fabrizio Perri for suggesting this line of exposition.

⁸In this simple setting, strong income effects raise the possibility of multiple steady states (e.g., see the discussion in Corsetti and Dedola [2002]). It is worth stressing, however, that the specification of preferences in the model we use in our numerical exercises below always ensure a unique steady state.

output (endowment) changes, and the terms of trade/real exchange rate can be expressed as follows:

$$\widehat{\tau} = \frac{\widehat{Y}_H - \widehat{Y}_F^*}{1 - 2a_H(1 - \omega)}, \quad (5)$$

$$\widehat{REER} = \frac{2a_H - 1}{1 - 2a_H(1 - \omega)} (\widehat{Y}_H - \widehat{Y}_F^*), \quad (6)$$

where a “ $\widehat{}$ ” represents a variable’s percentage deviation from the symmetric values. Consistent with our analysis above, these expressions show that, for given movements in relative output, the sign of the response of international relative price changes depending on ω . In addition, they suggest that the volatility of the terms of trade and the real exchange rate follows a hump-shaped pattern in ω .

To see this, assume home bias in consumption ($a_H > 1/2$). For a sufficiently high elasticity of substitution, i.e. $\omega > \frac{2a_H - 1}{2a_H}$, the real exchange rate and the terms of trade both depreciate in response to a positive Home supply shock. This is the region of parameters’ values in which the world demand schedule is conventionally sloped, and the international transmission is positive. In this region, higher values of ω reduce the coefficient relating $\widehat{Y}_H - \widehat{Y}_F^*$ to \widehat{REER} and $\widehat{\tau}$: the larger the price elasticity, the lower the volatility of the real exchange rate and the terms of trade relative in response to shocks to relative output.

Conversely, for a sufficiently low price elasticity of imports, that is, for $0 < \omega < \frac{2a_H - 1}{2a_H} < 1/2$, a Home supply shock cause both the real exchange rate and the terms of trade to appreciate in equilibrium. As shown above, underlying this result is a weak substitution effect relative to the income effect of changes in relative prices, so that the domestic and world demand schedules for Home tradables are negatively sloped. In this region of parameters’ values, a higher elasticity of substitution tends to raise the volatility of international prices.

Note that the response of international relative prices to output shocks tend to become stronger as ω approaches $\frac{2a_H - 1}{2a_H}$ from either side, whereas the slope of the demand function becomes flatter and flatter before changing sign. For ω around the cutoff point, the coefficient relating $\widehat{Y}_H - \widehat{Y}_F^*$ to \widehat{REER} and $\widehat{\tau}$ in the above expressions becomes quite high in absolute value, driving up the volatility of the real exchange rate and the terms of trade in response to shocks to relative output. An important implication of our analysis is that there will be two values of ω (below and above $\frac{2a_H - 1}{2a_H}$) that yield the same volatility of the terms of trade and real exchange rate: one associated with positive, the other associated with negative international transmission.

2.2.2 Exchange rates, consumption and risk sharing

So far, we have shown that there can be different patterns of relative price movements in response to supply shocks, shaping the sign and magnitude of the international transmission mechanism. We can now derive the implications of our results for risk sharing and the equilibrium comovements between the real exchange rate and relative consumption. With incomplete markets the scope for insurance against country-specific shocks is limited, and equilibrium movements in international relative prices will expose consumers to potentially strong relative wealth shocks.

In our simple model with financial autarky, we can use the balanced-trade condition to derive an expression for relative consumption as a function of the terms of trade:

$$\tau C_F = C_H^* \iff \frac{C}{C^*} = \left[\frac{(1 - a_H^*) \tau^{1-\omega} + a_H^*}{a_H \tau^\omega + (1 - a_H) \tau} \right]^{\frac{\omega}{1-\omega}}; \quad (7)$$

from this, we can then derive the following log-linearized relationship between the real exchange rate and relative consumption:

$$\widehat{REER} = \frac{2a_H - 1}{2a_H\omega - 1} (\widehat{C} - \widehat{C}^*). \quad (8)$$

The relation between real exchange rates and relative consumption can have either sign, depending on the values of a_H and ω . Specifically, with home bias in consumption, it will be negative when $\omega < \frac{1}{2a_H} < 1$.

We have seen above that, for a given change in the terms of trade and the real exchange rate, the international transmission of shocks can be positive or negative, depending on whether ω is above or below $\frac{2a_H - 1}{2a_H}$. But this cutoff point is smaller than $\frac{1}{2a_H}$. Hence, a negative correlation between the real exchange rate and relative consumption can correspond to different patterns of the international transmission. Specifically, in response to a Home supply shock, the Home terms of trade improves and the real exchange rate appreciates, while Home consumption rises relative to Foreign consumption, when $\omega < \frac{2a_H - 1}{2a_H}$; the Home terms of trade and exchange rate depreciates, driving Foreign consumption above domestic consumption, when $\frac{2a_H - 1}{2a_H} < \omega < \frac{1}{2a_H}$. Depending on the size of equilibrium movements in prices, consumption at Home may or may not fall — i.e., accounting for the Backus Smith evidence in this case does not necessarily imply ‘immiserizing growth.’

Contrast these results with the benchmark economy constructed by Cole and Obstfeld [1991], which is a special case with $\omega = 1$ and $a_H = a_H^* = 1/2$. This contribution — as well as Corsetti and Pesenti [2001a] — build examples where productivity shocks to tradables bring about relative price movements that *exactly* offset changes in output, leaving cross-country relative wealth unchanged.

Even under financial autarky, agents can achieve the optimal degree of international risk sharing, under the additional assumption of logarithmic utility ($\sigma = 1$) in (1).

But optimal risk sharing via terms-of-trade movements is likely to be an extreme case, since according to the evidence, both the magnitude of relative price movements and especially the sign of the transmission appear to be different from what is required to support an efficient allocation. Even when the international transmission is positive — as should be in the examples by Cole and Obstfeld [1991] and Corsetti and Pesenti [2001a] — equilibrium fluctuations in real exchange rates and the terms of trade of the magnitude of those observed in the data may be *excessive* relative to the benchmark case of optimal transmission and hinder international risk sharing, as is the case when $\frac{2a_H - 1}{2a_H} < \omega < \frac{1}{2a_H}$. Our analysis above unveils that an “*excessively positive*” international transmission of productivity shock generates an empirical pattern of low risk-sharing and can therefore rationalize the Backus-Smith anomaly: large terms-of-trade and real exchange rate depreciations will be reflected in a reduction in relative consumptions.

Risk-sharing is clearly hindered by a negative transmission, which prevails when $\omega < \frac{2a_H - 1}{2a_H}$. In this case, a terms of trade appreciation in response to a productivity shock raises domestic real import and consumption, while reducing wealth abroad — again in line with the Backus-Smith evidence, but at odds with risk-sharing via relative price movements.

2.3 Implications for open economy models

The above stylized two-country, two-good model with financial autarky and endowment shocks shows that, depending on the price elasticity of tradables, the correlation between relative consumption and the real exchange rate can have either sign. These results emphasize a low price elasticity as a promising modelling strategy to address the Backus-Smith anomaly. In what follows, we pursue this strategy by developing a fully-fledged dynamic model with capital accumulation and international trade in uncontingent bonds. Different from standard models, however, a low price elasticity of tradables will not be exclusively related to the elasticity of substitution ω , but will be an equilibrium implication of assuming a realistic structure of the goods market, whereas we introduce distribution services.

It is worth stressing that our explanation of the Backus-Smith puzzle abstracts from nominal rigidities and demand shocks — consistent with previous results from leading contributions. One may argue that the standard Mundell-Fleming-Dornbusch model also suggest a way to rationalize the Backus-Smith observation as a consequence of demand shocks. In this model, shocks to demand that drive domestic expenditure and consumption up appreciate the currency in real terms. Some

external demand needs to be crowded out in order to make “more room” for domestic demand. This model thus appears to be consistent with the above evidence, but only to the extent that international business cycles and real exchange rate fluctuations can be described as mainly driven by demand shocks. Moreover, allowing for demand shocks (monetary and government spending shocks) in a two-country model with sticky prices (set by producers in the currency of the market of destination), Chari, Kehoe and McGrattan [2002] emphasize that the correlation between relative consumption and the real exchange rate remains close to 1 even when the only internationally traded asset is a nominal bond. In light of this result, in what follows we abstract from nominal rigidities altogether.⁹

3 The model

In this and the next section, we develop our model, which will then be solved by employing standard numerical techniques. Our world economy consists of two countries of equal size, as before denoted H and F, each specialized in the production of an intermediate, perfectly tradable good. In addition, each country produces a nontradable good. This good is either consumed or used to make intermediate tradable goods H and F available to domestic consumers. In what follows, we describe our setup focusing on the Home country, with the understanding that similar expressions also characterize the Foreign economy — whereas starred variables refer to Foreign firms and households.

3.1 The Firms’ Problem

Firms producing Home tradables (H) and Home nontradables (N) are perfectly competitive and employ a technology that combines domestic labor and capital inputs, according to the following Cobb-Douglas functions:

$$\begin{aligned} Y_H &= Z_H K_H^{1-\xi} L_H^\xi \\ Y_N &= Z_N K_N^{1-\zeta} L_N^\zeta, \end{aligned}$$

⁹Nonetheless, one can still envision shocks, e.g., taste shocks to utility, that move the level of consumption and the marginal utility of consumption in opposite directions, thus mechanically attenuating the link between the real exchange rate and relative consumption in (1) also in models assuming complete asset markets. While Chari, Kehoe and McGrattan [2002] cast doubts on this perfunctionary rationalization of the Backus-Smith puzzle by showing that traditional demand shocks can hardly match the effects of those unobservable taste shocks, in Section 5 we will conduct sensitivity analysis to the inclusion of taste shocks, following Stockman and Tesar [1995].

where Z_H and Z_N are exogenous random disturbance following a statistical process to be determined below. We assume that capital and labor are freely mobile across sectors. The problem of these firms is standard: they hire labor and capital from households to maximize their profits:

$$\begin{aligned}\pi_H &= \bar{P}_{H,t}Y_{H,t} - W_tL_{H,t} - R_tK_{H,t} \\ \pi_N &= P_{N,t}Y_{N,t} - W_tL_{N,t} - R_tK_{N,t},\end{aligned}$$

where $\bar{P}_{H,t}$ is the *wholesale* price of the Home traded good and $P_{N,t}$ is the price of the nontraded good. W_t denote the wage rate, while R_t represents the capital rental rate.

Firms in the distribution sector are also perfectly competitive. They buy tradable goods and distribute them to consumers using nontraded goods as the only input in production. As in Burstein, Neves and Rebelo [2003] and Corsetti and Dedola [2002], we assume that bringing one unit of traded goods to Home (Foreign) consumers requires η units of the Home (Foreign) nontraded goods.

3.2 The Household's Problem

3.2.1 Preferences

The representative Home agent in the model maximizes the expected value of her lifetime utility, given by:

$$E \left\{ \sum_{t=0}^{\infty} U [C_t, \ell_t] \exp \left[\sum_{\tau=0}^{t-1} -\nu (U [C_t, \ell_t]) \right] \right\} \quad (9)$$

where instantaneous utility U is a function of a consumption index, C , and leisure, $(1 - \ell)$. Foreign agents' preferences are symmetrically defined. It can be shown that, for all parameter values used in the quantitative analysis below, these preferences guarantee the presence of a locally unique symmetric steady state, independent of initial conditions.¹⁰

The full consumption basket, C_t , in each country is defined by the following CES aggregator

$$C_t \equiv \left[a_T^{1-\phi} C_{T,t}^\phi + a_N^{1-\phi} C_{N,t}^\phi \right]^{\frac{1}{\phi}}, \quad \phi < 1, \quad (10)$$

where a_T and a_N are the weights on the consumption of traded and nontraded goods, respectively and $\frac{1}{1-\phi}$ is the constant elasticity of substitution between $C_{N,t}$ and $C_{T,t}$. As in Section 2, the consumption index of traded goods $C_{T,t}$ is given by (2).

¹⁰A unique invariant distribution of wealth under these preferences will allow us to use standard numerical techniques to solve the model around a stable nonstochastic steady state when only a non-contingent bond is traded internationally (see Obstfeld [1990], Mendoza [1991], and Schmitt-Grohe and Uribe [2001]).

3.2.2 Price indexes

A notable feature of our specification is that, because of distribution costs, there is a wedge between the producer price and the consumer price of each good. Let $\bar{P}_{H,t}$ and $P_{H,t}$ denote the price of the Home traded good at the *producer* and *consumer* level, respectively. Let $P_{N,t}$ denote the price of the nontraded good that is necessary to distribute the tradable one. With competitive firms in the distribution sector, the consumer price of the traded good is simply

$$P_{H,t} = \bar{P}_{H,t} + \eta P_{N,t}. \quad (11)$$

We hereafter write the utility-based CPIs, whereas the price index of tradables is given by (3):

$$P_t = \left[a_T P_{T,t}^{\frac{\phi}{\phi-1}} + a_N P_{N,t}^{\frac{\phi}{\phi-1}} \right]^{\frac{\phi-1}{\phi}}. \quad (12)$$

Foreign prices, denoted with an asterisk and expressed in the same currency as Home prices, are similarly defined. Observe that the law of one price holds at the wholesale level but not at the consumer level, so that $\bar{P}_{H,t} = \bar{P}_{H,t}^*$ but $P_{H,t} \neq P_{H,t}^*$. In the remainder of the paper, the price of Home aggregate consumption P_t will be taken as the numeraire. Hence, the real exchange rate will be given by the price of Foreign aggregate consumption P_t^* in terms of P_t .

3.2.3 Budget constraints and asset markets

Home and Foreign agents hold an international bond, B_H , which pays in units of Home aggregate consumption and is zero in net supply. Agents derive income from working, $W_t \ell_t$, from renting capital to firms, $R_t K_t$, and from interest payments, $(1 + r_t) B_{H,t}$, where r_t is the real bond's yield, paid at the beginning of period t but known at time $t - 1$. The individual flow budget constraint for the representative agent in the Home country is therefore:¹¹

$$P_{H,t} C_{H,t} + P_{F,t} C_{F,t} + P_{N,t} C_{N,t} + B_{H,t+1} + \bar{P}_{H,t} I_{H,t} \leq W_t \ell_t + R_t K_t + (1 + r_t) B_{H,t}. \quad (13)$$

We assume that investment is carried out in Home tradable goods and that the capital stock, K , can be freely reallocated between the traded (K_H) and nontraded (K_N) sectors:¹²

$$K = K_H + K_N.$$

¹¹ $B_{H,t}$ denotes the Home agent's bonds accumulated during period $t - 1$ and carried over into period t .

¹² We also conduct sensitivity analysis on our specification of the investment process, below.

As opposed to consumption goods, we assume that investment goods do not require distribution services. The price of investment is therefore equal to the wholesale price of the domestic traded good, $\bar{P}_{H,t}$. The law of motion for the aggregate capital stock is given by:

$$K_{t+1} = I_{H,t} + (1 - \delta)K_t \quad (14)$$

The household's problem then consists of maximizing lifetime utility, defined by (9), subject to the constraints (13) and (14).

3.3 Competitive Equilibrium

Let $s_t = \{B_H; \mathbf{Z}\}$ denote the state of the world at time t , where $\mathbf{Z} = \{Z_H, Z_F, Z_N, Z_N^*\}$. A competitive equilibrium is a set of Home agent's decision rules $C_H(s), C_F(s), C_N(s), I_H(s), l(s), B_H(s)$; a set of Foreign agent's decision rules $C_H^*(s), C_F^*(s), C_N^*(s), I_H^*(s), l^*(s), B_H^*(s)$; a set of Home firms' decision rules $K_H(s), K_N(s), L_H(s), L_N(s)$; a set of Foreign firms' decision rules $K_H^*(s), K_N^*(s), L_H^*(s), L_N^*(s)$; a set of pricing functions $P_H(s), P_F(s), \bar{P}_H(s), \bar{P}_F(s), P_N(s), P_N^*(s), W(s), W^*(s), R(s), R^*(s), r(s)$ such that (i) the agents' decision rules solve the households' problems; (ii) the firms' decision rules solve the firms' problems; and (iii) the appropriate market-clearing conditions (for the labor market, the capital market and the bond market) hold.

4 Model calibration

Table 2 reports our benchmark calibration, which we assume symmetric across countries. Several parameter values are similar to those adopted by Stockman and Tesar [1995], who calibrate their models to the United States relative to a set of OECD countries on annual data. Throughout the exercise, we will carry out sensitivity analysis and assess the robustness of our results under the benchmark calibration.

Productivity shocks We previously defined the exogenous state vector as $\mathbf{Z} \equiv \{Z_H, Z_F, Z_N, Z_N^*\}'$. We assume that disturbances to technology follow a trend-stationary AR(1) process

$$\mathbf{Z}' = \lambda \mathbf{Z} + \mathbf{u}, \quad (15)$$

whereas $\mathbf{u} \equiv (u_H, u_F, u_N, u_N^*)$ has variance-covariance matrix $V(\mathbf{u})$, and λ is a 4×4 matrix of coefficients describing the autocorrelation properties of the shocks. Since we assume a symmetric economic structure across countries, we also impose cross-country symmetry on the autocorrelation and variance-covariance matrices of the above sectoral process.

Consistent with our model and other open-economy studies (e.g., Backus, Kehoe and Kydland [1995]), we identify technology shocks with Solow residuals in each sector, using annual data in manufacturing and services from the OECD STAN database. Since hours are not available for most other OECD countries, we use sectoral data on employment. An appendix describes our data in more detail.

The bottom panel of Table 2 reports our estimates of the parameters describing the process driving productivity. As found by previous studies, our estimated technology shocks are fairly persistent. On the other hand, in line with empirical studies, we find that spillovers across countries and sectors are not negligible.¹³

Preferences and production Consider first the preference parameters. Assuming a utility function of the form:

$$U [C_t, \ell_t] = \frac{[\varkappa_t C_t^\alpha (1 - \ell_t)^{1-\alpha}]^{1-\sigma} - 1}{1 - \sigma}, \quad 0 < \alpha < 1, \quad \sigma > 0, \quad (16)$$

where \varkappa_t is a taste shock, we set α so that in steady state, one third of the time endowment is spent working; σ (risk aversion) is set equal to 2. Following Schmitt-Grohe and Uribe [2001], we assume that the endogenous discount factor depends on the average per capita level of consumption, C_t , and hours worked, ℓ_t , and has the following form:

$$\nu (U [C_t, \ell_t]) = \begin{cases} \ln (1 + \psi [C_t^\alpha (1 - \ell_t)^{1-\alpha}]) & \sigma \neq 1 \\ \ln (1 + \psi [\alpha \ln C_t + (1 - \alpha) \ln(1 - \ell_t)]) & \sigma = 1 \end{cases},$$

whereas ψ is chosen such that the steady-state real interest rate is 4 percent per annum, equal to 0.08. This parameter also determines the speed of convergence to the unique nonstochastic steady state.

The value of ϕ is selected based on the available estimates for the elasticity of substitution between traded and nontraded goods. We use the estimate by Mendoza [1991] referred to a sample of industrialized countries and set that elasticity equal to 0.74. Stockman and Tesar [1995] estimate a lower elasticity (0.44), but their sample includes both developed and developing countries.

As regards the weights of domestic and foreign tradables in the tradables consumption basket (C_T), a_H and a_F (normalized to $a_H + a_F = 1$) are chosen such that imports are 5 percent of aggregate output in steady state. This corresponds to the average ratio of U.S. imports from Europe, Canada

¹³See Costello [1993]. The persistence of the estimated shocks, though in line with estimates both in the closed (e.g., Cooley and Prescott [1995]) and open-economy (Heathcote and Perri [2002]) literature, is higher than that reported by Stockman and Tesar [1995]. The difference can be attributed to the fact that they compute their Solow residuals from HP-filtered data - while we and most of the literature compute them using data in (log) levels.

and Japan to U.S. GDP between 1960 and 2002. The weights of traded and nontraded goods, a_T and a_N , are chosen as to match the share of nontradables in the U.S. consumption basket. Over the period 1967-2002, this share is equal to 53 percent on average. Consistently, Stockman and Tesar [1995] suggest that the share of nontradables in the consumption basket of the seven largest OECD countries is roughly 50 percent.

We calibrate ξ and ζ , the labor shares in the production of tradables and nontradables, based on the work of Stockman and Tesar [1995]. They calculate these shares to be equal to 61 percent and 56 percent, respectively. Finally, we set the depreciation rate of capital equal to 10 percent annually.

Distribution costs and the price elasticity of tradables The introduction of a distribution sector in our model is a novel feature relative to standard business cycle models in the literature. Before delving into numerical analysis, it is appropriate to discuss an important implication of this feature regarding the price elasticity of tradables. From the representative consumer's first-order conditions (regardless of frictions in the asset and goods markets), optimality requires that the relative price of the imported good in terms of the domestic tradable at consumer level be equal to the ratio of marginal utilities:

$$\frac{P_{F,t}}{P_{H,t}} = \frac{\bar{P}_{F,t} + \eta P_{N,t}}{\bar{P}_{H,t} + \eta P_{N,t}} = \frac{1 - a_H}{a_H} \left(\frac{C_{H,t}}{C_{F,t}} \right)^{\frac{1}{\omega}}, \quad (17)$$

where $\omega = (1 - \rho)^{-1}$ is equal to the elasticity of substitution between Home and Foreign tradables in the consumption aggregator $C_{T,t}$, and thus to the consumer price elasticity of these goods. Note that $C_{H,t}/C_{F,t}$ is the inverse of the ratio of real imports to nonexported tradable output net of investment. In analogy to the literature, we can refer to this ratio as the (tradable) import ratio.

Because of distribution costs, the relative price of imports in terms of Home exports at the consumer level does not coincide with the terms of trade $\bar{P}_{F,t}/\bar{P}_{H,t}$ — as in most standard models (e.g. Lucas [1982]). Let μ denote the size of the distribution margin in steady state, i.e., $\mu = \eta \frac{P_N}{P_H}$. By log-linearizing (17), we get:

$$\hat{\tau}_t = \frac{1}{\omega(1 - \mu)} \left(\widehat{C}_{H,t} - \widehat{C}_{F,t} \right). \quad (18)$$

where the terms of trade τ is measured at the producer-price level so that $\omega(1 - \mu)$ can be thought of as the producer price elasticity of tradables. Clearly, both ω and μ impinge on the *magnitude* of the international transmission of country-specific shocks through the equilibrium changes in the terms of trade. It is well known that for any given change in $\widehat{C}_{H,t} - \widehat{C}_{F,t}$, a lower ω transpires

into larger changes in the terms of trade. In our model, a larger distribution margin μ (i.e., a larger η) has a similar effect. Accounting for distributive trade thus results into an amplification of fluctuations in international relative prices for any given variability in real quantities. So, for given ω and μ , large movements in the difference between the real consumption of domestic and imported tradables $\widehat{C}_{H,t} - \widehat{C}_{F,t}$ (the inverse of the import ratio) will be reflected in highly volatile terms of trade and deviations from the law of one price.¹⁴ Remarkably, it will be shown below that in the U.S. data the absolute standard deviation of this ratio is very close to that of the terms of trade (4.13 and 3.68 per cent, respectively).¹⁵

There is considerable uncertainty regarding trade price elasticities. Using time series data, empirical researchers have found estimates that range from about 0.1 to 2 (see the comprehensive study on G-7 countries by Hooper, Johnson and Marquez [2000]). For instance, for the U.S. Taylor [1993] estimates a value of 0.39, while Whalley [1985] finds it to be 1.5. For European countries most empirical studies suggest values below 1.¹⁶ Correspondingly, there are differences in the quantitative literature. For instance, in a model with traded and nontraded goods similar to ours, Stockman and Tesar [1995] set the parameter ω — directly related to the price elasticity with no distribution costs — equal to 1. Following Whalley [1985], in a model with only tradable goods Backus, Kydland, and Kehoe [1995] set it equal to 1.5, whereas Heathcote and Perri [2002] estimate it as low as 0.9. However, these authors also report sensitivity analysis suggesting that much lower values, in the range of 0.5, can improve their model performance in accounting for features of the international business cycle like the volatility of the terms of trade.

Given the uncertainty surrounding the appropriate parameter value, and the key role this elasticity plays in open economy models, we choose to follow a different route. First, we rely on estimates in the trade literature on distribution costs to pick a value for μ . According to the evidence for the U.S. economy in Burstein, Neves and Rebelo [2003], the share of the retail price

¹⁴In particular, the tradable import ratio will display more variability, *ceteris paribus*, when changes in absorption of domestic and imported tradables have opposite signs.

¹⁵Note that under financial autarky the counterpart of condition (4) in our fully-specified model with distribution services is:

$$\frac{\partial C_H}{\partial \tau} > 0 \iff \underbrace{\omega(1-\mu)(1-a_H)\left(\frac{P_F}{P_H}\right)^{1-\omega}}_{SE} - \underbrace{(1-a_H)\left(\frac{P_F}{P_H}\right)^{1-\omega} - a_H\mu}_{IE} > 0.$$

A positive distribution margin μ reduces the substitution effect (*SE*) from a deterioration in the terms of trade, while making the income effect (*IE*) more negative, as the presence of distributive trade causes the consumer price to fall less than one-to-one relative to the relative price of domestic tradables.

¹⁶Ruhl [2003] shows a way to reconcile these time series estimates with the contrasting evidence on the large growth in trade volumes resulting from changes in tariffs.

of traded goods accounted for by local distribution services ranges between 40 percent and 50 percent, depending on the industrial sector. In their exhaustive survey on trade costs, Anderson and van Wincoop [2004] report that in industrialized countries representative local distribution costs account for over 55 percent of retail prices. Thus, we follow the calibration in Burstein, Neves and Rebelo [2003] and set distribution costs to 50 percent.

Second, we set the elasticity of substitution ω to match the volatility of the U.S. real exchange rate relative to that of U.S. output, equal to 3.28 (see Table 3 below). Therefore, our quantitative analysis below can be interpreted as investigating the link among international price movements, risk sharing and the international transmission conditional on the model being consistent with the observed volatility in real exchange rates. In Section 2.2, we have used a simplified setup to show that the volatility of international prices is hump-shaped in ω , and discussed at length the mechanism underlying this pattern. Consistently, we find two values for the elasticity ω such that the model matches the volatility of the U.S. real exchange rate. In our benchmark calibration these two values are $\omega = 0.95$ and $\omega = 1.14$. Strikingly both values are quite close to that assumed by Stockman and Tesar [1995], implying similar consumers preferences across tradables for given consumer prices. Most important, when combined with the calibrated value for μ , the implied trade price elasticities are well in the range of available estimates. While apparently close to each other, however, the two possible values for ω imply quite different dynamics and international transmission patterns for shocks to tradables productivity. These differences will become central to our discussion of the evidence discussed at the end of the paper.

5 Real exchange rate volatility and the international transmission of productivity shocks

Our goal in this section is to verify whether our model can match the empirical evidence on the unconditional correlation between international prices and quantities, as well as the their relative volatilities. The evidence is summarized by the statistics reported in the first column of Tables 3 and 4. The statistics for the data — all filtered using the Hodrick and Prescott filter — are computed with the United States as the home country and an aggregate of the OECD comprising the European Union, Japan and Canada as the foreign country.¹⁷ Notably, the Backus Smith correlation between relative consumption and the real exchange rate is equal to -0.45.

In what follows, we will show that, different from standard open-economy models, our artificial

¹⁷Here we follow Heathcote and Perri [2002]. See the Data Appendix for details.

economy performs quite well in this dimension. Throughout our exercises, we take a first-order Taylor series expansion around the deterministic steady state and simulate our model economy using King and Watson [1998]’s algorithm. We compute the model’s statistics by logging and filtering the model’s artificial time series using the Hodrick and Prescott filter and averaging moments across 100 simulations. Consistently with our dataset, in our simulations changes in aggregate GDP and consumption are computed using constant prices (precisely, we use relative steady state prices). The results for our baseline model and some variations on it are also shown in Tables 3 and 4.

5.1 Volatilities and correlation properties

The real exchange rate and the terms of trade Using our framework, we can write the real exchange rate ($REER$) in the following log-linear form, reflecting movements in the terms of trade as well as in the relative price of non-traded goods:

$$\widehat{REER}_t = (1 - \mu)(2a_H - 1)\widehat{\tau}_t + \mu\left(\widehat{P}_{N,t}^* - \widehat{P}_{N,t}\right) + \Omega\left(\widehat{q}_t^* - \widehat{q}_t\right), \quad (19)$$

where $0 < \Omega < 1$ and \widehat{q}_t represents the relative price of nontradables.¹⁸ In our numerical results, the first two components, arising from home bias in consumption and deviations of the law of one price for the CPI of tradables, dominate real exchange-rate movements.

In our baseline economy the real exchange rate and the terms of trade are tightly related. Their correlation is positive (and equal to 0.97 and 0.98, depending on our calibrated value of ω), though higher than in the data (0.6). A positive sign for this correlation is an important result relative to alternative models that — like ours — allow for deviations from the law of one price but do so by assuming sticky prices in the buyer’s currency. As argued by Obstfeld and Rogoff [2001], these models can generate high exchange rate volatility as well, but at the cost of inducing a counterfactual negative correlation between the real exchange rate and the terms of trade.

The terms of trade is very volatile, even more than in the data. The volatility of the terms of trade relative to output is 2.84 with $\omega = 0.95$, and 4.47 with $\omega = 1.14$, compared to 1.79 in the data. In this sense, our model suggests that high volatility of the international prices *per se* is not a measure of their ‘disconnect’ from fundamentals. To stress this point, consider the volatility of the import ratio (IR), defined as the ratio of real imports to nonexported tradable output net of investment (empirically, we compute this ratio using manufacturing output). As shown in Table 4, the standard deviation of the import ratio is 4.13 percent in the data. In our

¹⁸Namely, $\Omega = a_N \bar{q}^{\frac{\phi}{\phi-1}} / (a_T + a_N \bar{q}^{\frac{\phi}{\phi-1}}) > 0$, where \bar{q} denotes a steady-state value and $\frac{1}{1-\phi}$ is the elasticity of substitution between tradables and nontradables.

benchmark parametrization, it is equal to 2.26 for the smaller ω , but increases to 4.33 percent for the larger ω .¹⁹

Moreover, with $\omega = 0.95$ the model is consistent with the ranking of variability in international prices observed in the data: the real exchange rate is more volatile than the terms of trade. The difference may be due either to the volatility of deviations from the law of one price (which drives a wedge between the terms of trade and relative prices at consumer levels) or to the volatility of nontradable prices, or a combination of the two. For this reason, the correct ranking of volatility is very hard to replicate using models that abstract from the features above (see Heathcoate and Perri [2002]).

We find that the relative price of nontradables across countries is not the main force driving the high volatility of the model's real exchange rate. Table 3 shows that the volatility of the relative price of nontradables predicted by our model is quite in line with that in the data: depending on ω , this volatility is 1.41 and 1.13, against an empirical estimate of 1.73. When we compute the ratio between the standard deviation of the relative price of nontradables across countries, and the standard deviation of the real exchange rate, this ratio is 40 percent and 33 percent, for the low and high ω respectively. This figure is in line with that estimated by Betts and Kehoe [2001], who find this ratio to be between 35 and 44 percent using a weighted average of U.S. bilateral real exchange rates.²⁰

The Backus-Smith correlation An important novel result of our baseline model shown in Table 3 is that the correlation between relative consumption and the real exchange rate is not only negative, but also quite close to its empirical counterpart. With $\omega = 0.95$ and $\omega = 1.14$, the correlation generated by the model is respectively -0.48 and -0.45, against our empirical estimate of -0.45. A similar pattern emerges for the terms of trade: its correlation with relative consumption is -0.63 and -0.66 in the model, against an empirical estimate of -0.53.

Since our two values of ω are set so that the model replicates the volatility of the real exchange rate in the data, our results show that the price elasticity that is consistent with a realistic volatility in international prices also implies a realistic pattern of risk-sharing.²¹ What generates a negative

¹⁹Remarkably, the data supports the tight and negative link between the terms of trade and the real exchange rate, on the one hand, and the import ratio, on the other hand, predicted by the theory. In the data these correlations stand at -0.68 and -0.41, respectively, against -1 predicted by the model for either value of ω .

²⁰Following a different procedure, Engel [1999] finds that deviations from the law of one price in traded goods virtually account for all of the volatility of the U.S. real exchange rate.

²¹The model can also generate a negative Backus-Smith correlation when we calibrate ω as to match the empirical volatility of the terms of trade (rather than the real exchange rate) relative to volatility of output. Following this

Backus-Smith correlation is the mechanism linking volatility and risk-sharing discussed in Section 2 in a very simple setting under financial autarky. However, note that the simple model predicts a perfectly negative correlation between relative consumption and the real exchange rate. In our baseline economy with capital accumulation and international borrowing and lending, the same mechanism accounts for the quantitative result of a negative but less than perfect correlation. It is instructive to inspect the reason underlying this difference in detail. When international asset trade is limited to uncontingent bonds, the relation between the real exchange rate and marginal utilities of consumption only holds in expected first-differences — the log-linearized Euler equations for the bond yield (abstracting from the time-varying discount factor):

$$E_t \left(\widehat{RE}R_{t+1} - \widehat{RE}R_t \right) \approx E_t \left[\left(\widehat{U}_{c,t+1}^* - \widehat{U}_{c,t}^* \right) - \left(\widehat{U}_{c,t+1} - \widehat{U}_{c,t} \right) \right]. \quad (20)$$

To the extent that the tight link between growth rates of variables is inherited by their levels, this expression suggests a mechanism that may prevent standard models allowing for borrowing and lending at international level to fit the Backus-Smith evidence. But in a stochastic environment, the international bond is traded only after the resolution of uncertainty, and does not provide households with ex-ante insurance against country-specific income shocks — it only makes it possible to reallocate wealth and smooth consumption over time. The *impact effect* of a shock to tradables in a bond economy will thus be roughly the same as under financial autarky, moving relative consumption and the real exchange rate in a direction that will depend on the value of the price elasticity. Under our calibration, the Backus-Smith correlation will therefore be negative on impact, but positive in the aftermath of a shock, when the dynamics of relative consumption and the real exchange rate is determined by the above equation. For this reason, the Backus-Smith correlation in a bond economy will be less negative than under financial autarky.²² It will also become higher and closer to that implied by complete markets, the weaker the impact response (in absolute value) of the real exchange rate — i.e., the less volatile the real exchange rate and the terms of trade on impact.²³

approach, we obtain a value of ω equal to 0.89, corresponding to a Backus-Smith correlation equal to -0.23. In this exercise the predicted volatility of the real exchange rate is about 74 percent of what is found in the data.

²²Interestingly, the model can also replicate the Backus-Smith correlation even when we look at first-differenced data. In our economy this correlation *ex-post* is -0.47 (-0.56) when ω equals 0.95 (1.14).

²³The same mechanism holds in an economy in which the consumption share of nontradables is set to zero, so that they are used only in distribution, and their production function is not subject to technology shocks. In this case, we find that the Backus-Smith correlation is around -0.90.

International relative prices and business cycles Consider now the rest of the statistics for the baseline economy in Tables 3 and 4. As is well known, most open-economy models — including those allowing for nominal rigidities and monetary shocks — predict a strong and positive link between relative output and real exchange rates. As Stockman [1998] points out, this prediction is at variance with the data: the empirical correlation shown in Table 3 is -0.23. A similar shortcoming concerns the correlation between relative output and the terms of trade, which is negative in the data (and equal to -0.20), while it tends to be positive in quantitative models.

Our baseline economy yields contrasting results on this issue. The correlation between relative output and the real exchange rate (the terms of trade) is high and positive — equal to 0.93 and 0.98 respectively — with $\omega = 1.14$, but becomes strongly negative with $\omega = 0.95$. This is because, with the lower ω , positive productivity shocks in the tradable sector appreciate the terms of trade and the real exchange rate — a result that we will discuss in greater detail below. We observe here that this very mechanism also accounts for the ability of the model to match the observed positive correlation between international relative prices and net exports, also shown at the bottom of the table.

In Table 4, we see that the cross-country correlation of GDP in the model (0.42 and 0.40 depending on ω) is very close to that in the data (0.49), and higher than that of consumption. The cross-correlation of consumption is lower than in the data (0.10 versus 0.32), while the cross-correlations of investment and employment are higher. While positive comovements in production mainly arise because of the positive cross-country correlation of the shock innovations, the baseline model still does relatively better in this dimension than the standard international real business cycle model. It is well known that this class of models predicts that consumption should be more correlated across countries than output, largely independent of the shocks' cross-country correlations, and that the correlation across countries of investment and employment is negative, even under financial autarky — in Backus, Kehoe and Kydland [1995] this empirical incongruity is dubbed the ‘quantity anomaly’.

Finally, a discrepancy between the benchmark model and the data is that — relative to output — consumption, investment, and employment are slightly less volatile than in data; net exports are about half as volatile in the model as in the data (0.25/0.38 against 0.63). However, note that our results with $\omega = 0.95$ account for countercyclical net exports. Their correlation with GDP is -0.51 in the model, very much in line with the data.

The Arrow-Debreu Economy The fourth column of Tables 3 and 4 reports results for an economy with a complete set of Arrow-Debreu securities. Since in such an economy the volatility of the real exchange rate is to a large extent independent of the price elasticity of imports, we only show numerical results for the lower value of ω — basically replicating the parameterization in Stockman and Tesar [1995]. As expected, including distribution services in such an environment is not enough to account for the Backus-Smith anomaly. The correlation between the real exchange rate and relative consumption is approximately equal to one. Moreover, the volatility of the real exchange rate, the terms of trade, the import ratio and net exports is several times lower than that in the data.

Nevertheless, this model generates a negative correlation between the real exchange rate and relative output, in line with the observed one. This is because productivity gains in the Home tradable sector raise relative output, worsen the Home terms of trade, but appreciate the real exchange rate — the real appreciation reflecting a higher relative price of nontradables and a fall in relative consumption in the period following the shock, driven by a drop in the consumption of nontradables. On the other hand, contrary to the data, the correlation between the terms of trade and relative output is positive, while the real exchange rate and the terms of trade are basically uncorrelated (0.02).

5.2 Sensitivity analysis

We now assess the sensitivity of our results to (a) removing the distribution sector from our baseline economy; (b) setting a very high elasticity of substitution of tradables — as to check whether the Backus-Smith correlation could be explained by a Balassa-Samuelson effect of productivity shocks on consumption and the real exchange rate; (c) removing cross-country spillovers from the process driving productivity shocks; (d) using different specifications of investment; and (e) introducing taste shocks. Results from these exercises are also shown in Tables 3 and 4.

Changing the distribution margin and the elasticity of substitution When we abstract from distributive trade and set $\eta = 0$, the two values of ω for which the relative volatility of the real exchange rate in the model is the same as in the data are 0.27 and 0.44, a good deal lower than in our benchmark economy. As discussed in Section 4, the need to combine tradables with retailing in our baseline economy makes the price elasticity of imports lower than the value implied by the preference parameter ω . Without retailing, we need to assume a lower elasticity of substitution between Home and Foreign goods in agents' preferences for the model to fit the volatility of the

real exchange rate.

With a lower elasticity of substitution but no distribution services, the model still performs remarkably well with respect to the Backus-Smith anomaly: the correlation between the real exchange rate and relative consumption is negative and equal to -0.22 (-0.81) for $\omega = 0.27$ (0.44). The underlying mechanism has already been thoroughly discussed in sections 2.2 and 2.3.

With $\eta = 0$, however, there are no deviations from the law of one price, contradicting an important stylized fact of the international economy (e.g., see Engel [1999]). As a consequence, movements in the relative price of nontradables across countries contribute to real exchange-rate fluctuations much more than in our benchmark economy. The standard deviation of the relative price of nontradables across countries is now 84 percent of that of the real exchange rate when $\omega = 0.27$ and 56 percent when $\omega = 0.44$, a significantly higher fraction than in the data. Moreover, the relative price of nontradables is more volatile than in the data and substantially more volatile than under our benchmark economy (2.83 and 2.07 against 1.41 and 1.13). These results suggest that introducing a distribution sector improves the performance of the model independently of contributing to lower the trade elasticities relative to what would be implied by preferences.

Balassa-Samuelson effects An interesting issue is whether the Backus-Smith anomaly could be accounted for by Balassa-Samuelson effects, linking exchange-rate fluctuations to movements in the relative price of nontradables. The idea is as follows. Consider a model in which domestic and foreign tradables are highly substitutable. A positive productivity shock to the tradable sector should appreciate the real exchange rate (terms of trade movements are tiny), *and* drive up domestic relative to foreign consumption. Is the Backus-Smith correlation driven mainly by this effect?

Our answer is no. In a numerical experiment, we abstract from distributive trade ($\eta = 0$) and assume a rather high value of ω , equal to 10 — so as to make tradables more homogeneous across countries and reduce the role of the terms of trade in exchange-rate fluctuations (results are the same for higher ω). With such a high elasticity of substitution, it is the correlation between the real exchange rate and relative output that becomes very negative (-0.54), but the corresponding correlation with relative consumption remains close to one, i.e. as high as 0.84. In addition, both the real exchange rate and the terms of trade are a great deal less volatile than output (1.06 and 0.26), while their cross-correlation is substantially lower than in the data (0.13).

Absence of Spillovers As shown in Table 2, the process driving productivity that we estimate and use in our model displays substantial cross-country spillovers. How much of our results can be attributed to the magnitude of such spillovers? It turns out that removing them altogether in our

numerical exercises does not substantially affect our main conclusions. Adopting the productivity process without spillovers, we again calibrate our economy such that the real exchange rate is as volatile as in the data, obtaining $\omega = 0.93$ and $\omega = 1.17$. The Backus-Smith correlation remains close to the one in our baseline economy: -0.65 and -0.34. However, one significant implication of removing spillovers is that consumption becomes negatively correlated across countries for $\omega = 0.93$.

Changing the investment specification In our baseline economy investment is carried out solely in domestically produced tradable goods. In our last exercise, we allow for a more general specification in which investment is a composite good comprising both Home and Foreign tradables. We assume that investment goods are given by the following CES aggregator

$$I_{T,t}(j) \equiv \left[a_H^{1-\rho} I_{H,t}(j)^\rho + a_F^{1-\rho} I_{F,t}(j)^\rho \right]^{\frac{1}{\rho}},$$

where $I_{H,t}$ ($I_{F,t}$) is the level of investment in terms of the domestic (imported) traded good. As in our baseline calibration, we set a_H and a_F such that imports (which now also include investment) are 5 percent of aggregate output in steady state. Since in this exercise investment goods also include imports, we also introduce distribution services in the price of investment to differentiate between trade-price elasticities and preferences. But following Burnstein, Neves, and Rebelo [2004], we set the share of distribution services in the price of investment to be 16.7 percent. In Tables 3 and 4 results are shown under the heading ‘‘CES Investment.’’

With the more general CES specification for investment, the values of ω needed to reproduce the volatility of the real exchange rate relative to that of output are smaller than under our benchmark calibration. This is because investment goods can now be imported from abroad, and investment does not require as much distribution services as consumer goods do. Thus, any given price elasticity of imports corresponds to a lower elasticity of substitution relative to our baseline specification. Nonetheless, the model still succeeds in generating a significant departure from the complete markets outcome. Although the real exchange rate and relative consumption are not as negatively correlated as in our previous experiments, their correlation remains well below zero. When $\omega = 0.5$, the model predicts a slightly negative correlation of -0.34.²⁴

Taste shocks The last two columns of Tables 3 and 4 report results from introducing taste shocks in an economy with a complete set of Arrow-Debreu securities and in our baseline economy, respec-

²⁴We also analyzed an economy without capital accumulation, whose results are not reported for the sake of space. Excluding capital does not substantially change the match of the model with the data along most dimensions. However, consumption becomes more volatile than output, while the volatility and cross-country correlation of employment tend to be very low.

tively. We follow Stockman and Tesar [1995] and calibrate the taste shocks \varkappa_t in (16) assuming that they are uncorrelated across countries and have a standard deviation and serial correlation equal to the largest between the two productivity shocks, 0.0089 and 0.961, respectively (see Table 2).

The results in the last column in Tables 3 and 4 show the performance of our baseline economy accounting for these shocks. Interestingly, our results are broadly unchanged. The only significant difference is that, these shocks mainly affecting consumption, its cross-country correlation becomes too low, basically zero.

The results shown in the column before the last also show that assuming preference shocks to utility that are as volatile and persistent as productivity shocks, can improve the ‘fit’ of the complete market economy along a number of dimensions, including the Backus-Smith puzzle. Namely, the correlation between the real exchange rate and relative consumption is equal to -0.22; that between the terms of trade and relative consumption is negative as well, equal to -0.58. Moreover, the terms of trade and real exchange rate are now positively correlated (0.21), although less than in the data. However, while taste shocks in the Arrow Debreu economy also raise the volatility of the real exchange rate, the terms of trade, the import ratio and net exports, volatility is still much lower than in the data. Also, the cross-correlation of consumption becomes slightly negative. The correlation between net exports and GDP, and terms of trade and relative GDP is not as strongly negative as in the data.

The introduction of taste shocks in international business cycle models with complete markets weaken the links between relative consumption and relative marginal utility, thus being functional to generate a low or negative correlation between real exchange rates and relative consumptions. Yet, it is far from obvious that this approach can satisfactorily address the evidence of the low degree of risk sharing in other dimensions. Overall, we take these results as suggesting that the basic mechanism behind the Backus-Smith puzzle requires a combination of supply and demand effects, as implied by imperfect risk pooling.²⁵

5.3 The international transmission of productivity shocks to tradables

In our model, given a value for the distribution margin μ , there are two values of price elasticity and thus of ω that generate a real exchange-rate volatility matching the evidence. In this subsection,

²⁵Another avenue would be to explore richer preferences. Interestingly, however, Chari, Kehoe and McGrattan [2002] show that allowing for habits formation in consumption, which has proved useful in understanding other puzzles (for instance, the equity premium puzzle), cannot account for the Backus-Smith anomaly.

we analyze the difference between these two parameterizations by looking at theoretical impulse responses to a shock to the traded goods sector.

Our experiments consist of shocking the exogenous process for sectoral productivity once by 1 percent at date 0, when both countries are at their symmetric, deterministic steady state, and let productivity be driven by the estimated autoregressive process in (15). Figure 2 draws the responses of the following economic variables: (a) the real exchange rate; (b) the terms of trade; (c) relative consumption; (d) relative aggregate output; (e) the ratio of net exports to output. The two columns in Figure 2 report impulse responses for $\omega = 0.95$ and $\omega = 1.14$, respectively.²⁶

Consider first the impulse responses under the higher ω (first column in the figure). Since for this value of the price elasticity world demand for Home tradables is increasing in its relative price, the increase in the supply of Home traded goods relative to the Foreign goods worsens the Home country's terms of trade. Note that an adverse effect of productivity shocks on the real exchange rate and the terms of trade is predicted by all standard models with product specialization and homothetic preferences (e.g., Lucas [1982] and Backus et al. [1995]).²⁷ The notable feature of our specification with incomplete markets is that a relatively low price elasticity of imports (also owing to the presence of retailing) magnifies the deterioration of the Home terms of trade and real exchange rate, increasing the ensuing negative wealth effect for the domestic household. As a result, consumption abroad rises by more than domestic consumption, while domestic output rises relative to the foreign one. Thus, the real exchange rate, the terms of trade and relative output on the one hand, and relative consumption on the other move in the opposite direction, as the large terms of trade worsening entails an *excessively positive* transmission of the productivity shock in favor of the Foreign country. Note that net exports increase following the rise in productivity, which is consistent with the depreciations of the real exchange rate and the terms of trade.

The response of the economy to an innovation in the productivity of the domestic traded sector is widely different when $\omega = 0.95$. In this case, relative output still rises, but the real exchange rate and the terms of trade now appreciate. Recalls from Section 2 that for a low enough price

²⁶Although not reported in the charts, all variables ultimately return to their steady-state values following this one-time shock, because of the endogeneity of the discount factor. As we mentioned previously, the slow convergence is due to the low value of the parameter ψ required to match the steady state real interest rate.

²⁷This result is seldom highlighted in models with traded and nontraded goods. A possible explanation is that in these models tradables are very often assumed to be perfectly homogeneous across countries, i.e., $\omega \rightarrow \infty$, so that there are no terms of trade fluctuations (see e.g., Stockman and Dellas [1989] and Tesar [1993]). With this specification, a technological advance in the traded-good sector typically brings about an appreciation of the domestic currency owing to an increase in the domestic relative price of nontradables, according to the Balassa-Samuelson hypothesis. Note, however, that these models obviously leave unexplained the terms of trade behavior.

elasticity (low enough ω), world demand for Home tradables will be negatively sloped in the terms of trade, owing to a prevailing negative income effect for the domestic household. An increase in the relative supply of Home tradables will thus require a terms-of-trade appreciation in equilibrium to bring about market clearing. And as the terms of trade improve, Home consumption rises by more than Foreign consumption. As a result, the real exchange rate, the terms of trade and relative consumption are again negatively correlated, but now relative output will move in the same direction as relative consumption, though by a lesser amount. Finally, the positive productivity shock triggers a fall in net exports, which can account for its well-known negative counter-cyclical movements.

To summarize, a productivity shock to the export sector always induces an increase in relative output and (conditional) negative comovements between the real exchange rate, the terms of trade and relative consumption. Depending on the strength of the price-elasticity of imports and thus on the slope of world demand, however, relative consumption can increase or fall in response to a positive shock.

6 Productivity, the real exchange rate and the terms of trade: evidence for the U.S.

In this section we study the comovements between the real exchange rate, the terms of trade, and relative consumption in response to productivity changes in the U.S. economy. Given our focus on time series evidence, we use VAR methods, extending work by Galí [1999] and Christiano, Eichenbaum and Vigfusson [2003] — where technology shocks are identified via long-run restrictions — to an open-economy context. We focus our study on the U.S. economy vis-à-vis an aggregate of other OECD countries.

A number of recent papers have investigated in a closed-economy framework the effects of technology shocks identified using long-run restrictions. This literature uses the basic insight from the standard stochastic growth model that only technology shocks should have a permanent effect on labor productivity to identify economy-wide technology shocks in the data.²⁸ However, since

²⁸See Shapiro and Watson [1988], Francis and Ramey [2003] and Chang and Hong [2002], among others. Some open-economy papers, following Blanchard and Quah [1989], use long-run restrictions derived in the context of the traditional aggregate demand and aggregate supply framework. For instance, Clarida and Galí [1994] identify supply shocks by assuming that demand and monetary shocks do not have long-run effects on relative output levels across countries. While monetary shocks satisfy this assumption in most models, fiscal or preference shocks do not, since they can have long-run effects on output (and hours) in the stochastic growth model.

Galí [1999], several contributions have pointed out that these methods yield results that may be sensitive to assumptions about the particular VAR specification, e.g. the number and kind of variables included and the time series of properties of the variables. In this vein, Christiano, Eichenbaum and Vigfusson [2003] show that the findings in Galí [1999] are turned around when variables like per capita hours worked are treated as a trend stationary process rather than as a difference stationary process, as does the latter author.

Following these insights, we thus examine the effects of technology shocks to the U.S. manufacturing sector (our proxy for traded goods), identified with long run restrictions, on the real exchange rate, the terms of trade, net exports and relative consumption and output, while carrying out several robustness checks. As Chang and Hong [2002] show that using total factor productivity (TFP) instead of labor productivity may affect results for the manufacturing sector, we also assess the robustness of our results to the use of (annual) TFP data. Moreover, the use of TFP provides a further check on the identification strategy, as it amounts to controlling for long-run effects on labor productivity brought about by changes in the long-run capital labor ratio.²⁹ Leaving to the data appendix a more detailed description of data sources, hereafter we briefly describe our approach and discuss the main results.

Over the period 1970 to 2001, we estimate two specifications of the following structural VAR model

$$\begin{bmatrix} \Delta x_t \\ \Delta y_t \end{bmatrix} = \begin{bmatrix} C^{xz}(L) & C^{xm}(L) \\ C^{yz}(L) & C^{ym}(L) \end{bmatrix} \begin{bmatrix} \Delta \varepsilon_t^z \\ \Delta \varepsilon_t^m \end{bmatrix}. \quad (21)$$

Here x_t denotes the variable that is assumed to be affected in the long run only by permanent technology shocks: in our two different specifications, this variable is equal to (the log of) U.S. quarterly manufacturing labor productivity and (the log of) annual manufacturing TFP, respectively, both measured in deviation from labor productivity in an aggregate of other OECD countries. In the quarterly specification y_t is a 5x1 vector of variables, including (the log of) U.S. aggregate GDP and consumption relative to that of a composite of other OECD countries, the U.S. ratio of net export over GDP, (the log of) the U.S. real effective (trade-weighted) exchange rate, and (the log of) the terms of trade (computed as the non-energy imports deflator over the exports deflator). In the annual specification, in order to save degrees of freedom y_t is 3x1. The first two components of the quarterly specification are always included, while the last three are included one by one.³⁰

²⁹For instance, Uhlig [2003] argues that a unit root in labor productivity may result not only from the standard RBC shock to TFP, but also from permanent shocks to the capital-income tax. .

³⁰We also estimated specifications of the model including more U.S. and international variables, like investment,

$C(L)$ is a polynomial in the lag operator; ε_t^z denotes the technology shock to manufacturing, and ε_t^m the other structural, non-technology shocks.³¹ In addition to the usual assumption that the structural shocks are uncorrelated, positing that $C^{xm}(1) = 0$ is enough to identify ε_t^z . This restricts the unit root in the variable x_t to originate solely in the technology shock. Although not necessary for identification, implicit in this benchmark specification is the assumption that all the other variables also have a unit root; this assumption is not rejected by the data over our sample. However, following the suggestions in Christiano, Eichenbaum and Vigfusson [2003], we also estimated specifications of the VAR with those variables, like the real exchange rate, for which the unit root null is not rejected only marginally, in levels. Our main findings below, that a technology improvement leads to a persistent terms-of-trade deterioration and real exchange-rate depreciation, are basically unaltered.³²

Figure 3 shows the effects of the identified technology shocks on the levels of productivity, relative consumption, the real exchange rate, and the terms of trade. The first column is obtained from quarterly data, the second one from annual data. We report error bands for the significance levels of 68 percent and 90 percent (corresponding to the darker and lighter shaded areas, respectively).³³

The first column in Figure 3 shows the impulse responses using Galí’s identification scheme, with x_t equal to (relative) U.S. manufacturing labor productivity. Following a positive technology shock to manufacturing, U.S. total consumption increases gradually but permanently relative to the rest of the world. Moreover, the real exchange rate and the terms of trade strongly appreciate on impact and remain permanently stronger, by an amount that is larger in the case of the real exchange rate, but that for both variables outsizes the increase in productivity. Net exports fall following the positive productivity shock, which is also consistent with the predictions of the model under the negative transmission.

The second column in Figure 3 reports the effects of a technology shock identified as the only shock that permanently affects TFP in U.S. manufacturing. Our findings are broadly robust across different long-run identification schemes. In the annual data VAR also a positive technology shock to the U.S. production of tradables appears to lead to an increase in domestic consumption relative

real wages and hours worked, and different definitions of the terms of trade. Since very similar results to those discussed in the text are obtained, they are not included to save on space. They are available from the authors upon request.

³¹We include up to four lags for quarterly data and one for annual data, based on a BIC criterion and tests of residual serial correlation.

³²These results are not included in the paper to save on space, but they are available upon request.

³³The standard error bands were computed using a bootstrap Monte Carlo procedure with 5000 replications. We thank Yongsung Chang for graciously providing us with his bootstrapping codes.

to the rest of the world, while improving the terms of trade and appreciating the real exchange rate for at least a year. As with quarterly data, the rise in productivity leads to a fall in net exports.³⁴

Finally, we also checked whether these results were robust to using a different identification scheme, namely, assuming that a technology shock is the contemporaneous innovation to relative labor productivity in U.S. manufacturing, while keeping the same order of the variables as in (21) but in levels rather than first-differences. This identification scheme is closer in spirit to the assumption entertained in the calibration that labor productivity is basically an exogenous process. Again, as the results were very similar to those obtained above we do not report them here for the sake of brevity.³⁵

To summarize, U.S. consumption relative to the rest of the world and the real exchange rate move in opposite directions, in sharp contrast with the predictions of the perfect risk-sharing hypothesis. Consistent with the Backus-Smith anomaly, the results in this section indicate that following changes in (relative) labor productivity in the traded goods sector real exchange rates and relative consumption can indeed be negatively correlated. Most interestingly, the appreciation of the real exchange rate, and especially the terms of trade, as well as the fall in net exports in response to a positive technology shock to domestic tradables, is qualitatively consistent with the transmission mechanism at work in our setup under the lower value of the price elasticity, but at odds with the presumption that an increase in the world supply of a good necessarily leads to a fall in its relative price.

It is worth stressing that we do not expect our empirical results for the U.S. — a very large, rich and relatively closed economy — to readily generalize to smaller and/or more open economies. Our theoretical model suggests that the wealth effects underlying the negative transmission of productivity shocks are rather unlikely in such economies. In this sense we see our findings as complementing the cross-sectional evidence provided by Acemoglu and Ventura [2003], showing that terms of trade changes are negatively related to output growth driven by capital accumulation. To reconcile the difference in our results it should be kept in mind that, first, we analyze time series for one country with quite distinctive features, as opposed to using cross-sectional data techniques.

³⁴Notice that the terms of trade appreciation cannot be easily rationalized with the well-known bias in measured import prices that arises from a lack of adjustment to an increasing number of imported goods (see Feenstra [1994]). As shown by Ruhl [2003], this bias is negatively correlated with the level of imports, as the (mis)measured price index fails to fall as much as the correct price index, thus biasing results against finding a terms of trade appreciation.

³⁵Consistently with our results, Alquist and Chinn [2002] find, with cointegrating techniques, that each percentage point increase in the U.S.-Euro area economy-wide labor productivity differential results in a 5-percentage-point real appreciation of the dollar in the long run.

Second, we adopt a setting that directly tackles the problem of disentangling changes in technology from other factors, like those affecting demand.

Our model links the terms of trade appreciation in response to productivity shocks to price elasticities, in turn reflecting market structure and basic features of the economy, such as Home bias in consumption. In their empirical contribution based on panel data techniques, Debaere and Lee [2003] attribute terms of trade appreciation to quality and variety upgrading effects, proxied by variables such as income per capita and spending on research and development. While the creation of new and better goods varieties obviously runs against a fall in a country's terms of trade, the question is whether productivity improvements (that reduce marginal costs of production and/or marginal costs of creation of new varieties) could lead to both the creation of new varieties and an appreciation of the terms of trade. Some theoretical work in this area suggest that the answer in standard trade and macro models with endogenous creation of new goods is negative or ambiguous (see Acemoglu and Ventura [2003] and Corsetti, Martin and Pesenti [2004]).³⁶

7 Concluding remarks

Many contributions to the literature have stressed that movements in the terms of trade in response to country-specific shocks may provide risk insurance to countries specialized in different types of goods. In this paper, we have reconsidered the link between exchange rate volatility and international consumption risk sharing, using a standard model with incomplete asset markets, where a low price elasticity of tradables arises from trade barriers in the form of distribution services. In numerical exercises conducted under a plausible parameterization of our world economy, we find that the international transmission of productivity shocks envisioned in our model can actually account both for the high volatility of international prices and for the (unconditional) negative link between the real exchange rate and relative consumption observed in the data.

We complement our theoretical analysis with suggestive evidence supporting the prediction of a negative conditional correlation between relative consumption and international relative prices. Following a permanent positive shock to U.S. labor productivity in manufacturing (our measure of tradable goods), domestic output and consumption increase relative to the rest of the world, but both the terms of trade and the real exchange rate appreciate. Consistent with our model, productivity improvements do not lead to a drop in the international relative price of domestic tradables. This result is reasonably robust to the definition of the terms of trade and the use of

³⁶A theoretical attempt to build a model encompassing a discussion of both elasticities and creation of new goods is provided by Ruhl [2003].

TFP instead of labor productivity.

Our analysis suggests that large equilibrium terms-of-trade movements, reflecting trade frictions, may be much less effective in providing insurance against production risk — and can even be counterproductive, in the sense of amplifying cross-country wedges in wealth stemming from asymmetric productivity shocks. In other words, international relative prices may move in ways that runs counter to efficient risk sharing. Given the relevance of this issue to our understanding of the international transmission of supply shocks and the mechanism of international risk-sharing, further empirical and theoretical work would prove extremely helpful.

References

- [1] Acemoglu, Daron, and Jaume Ventura [2003]. “The World Income Distribution,” *Quarterly Journal of Economics*, forthcoming.
- [2] Alquist, Ron, and Menzie David Chinn [2002]. “The Euro and the Productivity Puzzle: An Alternative Interpretation,” mimeo, University of Michigan.
- [3] Anderson, James, and Eric van Wincoop [2004]. “Trade costs”, *Journal of Economic Literature*, forthcoming.
- [4] Backus, David K., Patrick J. Kehoe, and Finn E. Kydland [1992]. “International Real Business Cycles,” *Journal of Political Economy* 101, 745-75.
- [5] Backus, David K., Patrick J. Kehoe, and Finn E. Kydland [1995]. “International Business Cycles: Theory and Evidence,” in Thomas F. Cooley (ed.) *Frontiers of Business Cycle Research*, Princeton University Press, Princeton, 331-56.
- [6] Backus, David K., and Gregor W. Smith, [1993]. “Consumption and Real Exchange Rates in Dynamic Economies with Non-traded Goods,” *Journal of International Economics* 35, 297-316.
- [7] Baxter, Marianne, and Mario J. Crucini [1995]. “Business Cycles and the Asset Structure of Foreign Trade,” *International Economic Review* 36, 821-54.
- [8] Betts Caroline and Patrick Kehoe [2001], Real Exchange Rate Movements and the Relative Price of Non Traded Goods, Federal Reserve Banks of Minneapolis, Research Department Staff Report
- [9] Blanchard Olivier J. and Danny Quah [1989]. “The Dynamic Effects of Aggregate Demand and Supply Disturbances,” *American Economic Review* 19, 655-73.
- [10] Burstein, Ariel T., João Neves, and Sergio Rebelo [2003]. “Distribution Costs and Real Exchange Rate Dynamics During Exchange-Rate-Based Stabilizations,” *Journal of Monetary Economics*, September, 1189-1214.
- [11] Burstein, Ariel T., João Neves, and Sergio Rebelo [2004]. “Investment Prices and Exchange Rates: Some Basic Facts,” National Bureau of Economic Research Working Paper no. 10238.
- [12] Chang, Yongsung, and Jay H. Hong [2002]. “On the Employment Effect of Technology: Evidence from U.S. Manufacturing for 1958-1996,” mimeo, University of Pennsylvania.

- [13] Chari, V.V., Patrick J. Kehoe, and Ellen McGrattan [2002]. "Can Sticky Prices Generate Volatile and Persistent Real Exchange Rates?," *Review of Economic Studies* 69, 633-63.
- [14] Christiano, Lawrence J., Martin Eichenbaum and Robert Vigfusson [2003], "What Happens After a Technology Shock?," mimeo, Northwestern University.
- [15] Clarida, Richard, and Jordi Galí [1994]. "Sources of Real Exchange Rate Fluctuations: How Important Are Nominal Shocks?" *Carnegie-Rochester Series in Public Policy* 41, 1-56
- [16] Cole, Harold L., and Maurice Obstfeld [1991]. "Commodity Trade and International Risk Sharing: How Much Do Financial Markets Matter?" *Journal of Monetary Economics* 28, 3-24.
- [17] Cooley, Thomas F. and Edward C. Prescott [1995]. "Economic Growth and Business Cycles," in Thomas F. Cooley (ed.) *Frontiers of Business Cycle Research*, Princeton University Press, Princeton, 1-38.
- [18] Corsetti, Giancarlo, and Paolo Pesenti [2001a]. "Welfare and Macroeconomic Interdependence," *Quarterly Journal of Economics*, May.
- [19] Corsetti Giancarlo and Luca Dedola [2002], "Macroeconomics of International Price Discrimination," European Central Bank Working Paper No. 176; *Journal of International Economics*, forthcoming.
- [20] Corsetti, Giancarlo, Philippe Martin and Paolo Pesenti [2004]. "Productivity, aggregate demand and real exchange rate: lessons from a general equilibrium analysis of the 'Home Market Effect'." mimeo, European University Institute
- [21] Costello, Donna [1993]. "A Cross-country, Cross-industry Comparison of Productivity Growth," *Journal of Political Economy* 101 (2), 207-22.
- [22] Debaere, Peter, and Hongshik Lee [2003]. "The Real-Side Determinants of Countries' Terms of Trade: A Panel Data Analysis," mimeo, University of Texas, Austin.
- [23] Engel, Charles [1999]. "Accounting for Real Exchange Rate Changes," *Journal of Political Economy* 107, 507-38.
- [24] Engel, Charles [2000]. "Comments on Obstfeld and Rogoff's "The Six Major Puzzles in International Macroeconomics: Is There a Common Cause?"," National Bureau of Economic Research Working Paper no. 7818.

- [25] Feenstra, Robert [1994]. “New Product Varieties and the Measurement of International Prices,” *American Economic Review* 84, 157-177.
- [26] Francis, Neville, and Valerie A. Ramey [2001]. “Is the Technology-Driven Real Business Cycle Hypothesis Dead? Shocks and Aggregate Fluctuations Revisited,” mimeo, University of California, San Diego.
- [27] Galí, Jordi [1999]. “Technology, Employment and the Business Cycle: Do Technology Shocks Explain Aggregate Fluctuations?” *American Economic Review* 89, 249-71.
- [28] Heathcote, Jonathan and Fabrizio Perri [2002]. “Financial Autarky and International Business Cycles,” *Journal of Monetary Economics* 49, 601-27.
- [29] Hooper, Peter, Karen Johnson and Jaime Marquez [2000]. “Trade elasticities for the G-7 countries.” *Princeton Studies in International Economics* No. 87.
- [30] King, Robert G. and Mark W. Watson [1998]. “The Solution of Singular Linear Difference Systems Under Rational Expectations,” *International Economic Review* 39, 1015-26.
- [31] Kollman, Robert [1995]. “Consumption, Real Exchange Rates and the Structure of International Asset Markets,” *Journal of International Money and Finance* 55, 29-57.
- [32] Lewis, Karen K. [1996]. “What Can Explain the Apparent Lack of International Consumption Risk Sharing?”, *Journal of Political Economy* 104 (2), 267-97.
- [33] Lewis, Karen K. [1999]. “Trying to Explain the Home Bias in Equity and Consumption,” *Journal of Economic Literature* 37, 571-608.
- [34] Lucas, Robert E., Jr. [1982]. “Interest Rates and Currency Prices in a Two-Country World,” *Journal of Monetary Economics* 32, 335-59.
- [35] Mendoza, Enrique [1991]. “Real Business Cycles in a Small Open Economy,” *American Economic Review* 81(4), 797-818.
- [36] Obstfeld, Maurice [1990]. “Intertemporal Dependence, Impatience, and Dynamics,” *Journal of Monetary Economics* 26, 45-75.
- [37] Obstfeld, Maurice, and Kenneth Rogoff [2001]. “The Six Major Puzzles in International Macroeconomics: Is There a Common Cause?” in Ben Bernanke and Kenneth Rogoff (eds.), *NBER Macroeconomics Annual 2000*, Cambridge, MA: MIT Press.

- [38] Ravn, Morten [2001]. "Consumption Dynamics and Real Exchange Rates," mimeo, London Business School.
- [39] Ruhl, Kim J. [2003]. "Solving the Elasticity Puzzle in International Economics," mimeo, University of Minnesota.
- [40] Schmitt-Grohé, Stephanie and Martín Uribe [2001]. "Closing Small Open Economy Models," *Journal of International Economics*, forthcoming.
- [41] Shapiro, Matthew D., and Mark Watson [1988]. "Sources of Business Cycle Fluctuations," *NBER Macroeconomics Annual 1998*, Cambridge, MA: MIT Press.
- [42] Stockman, Alan C. [1998]. "New Evidence Connecting Exchange Rates to Business Cycles," Federal Reserve Bank of Richmond *Economic Quarterly* 84, 73-89.
- [43] Stockman, Alan C., and Harris Dellas [1989]. "International Portfolio Nondiversification and Exchange Rate Variability," *Journal of International Economics* 26 (3-4), 271-90.
- [44] Stockman, Alan C., and Linda Tesar [1995]. "Tastes and Technology in a Two-Country Model of the Business Cycle: Explaining International Comovements," *American Economic Review* 83, 473-86.
- [45] Taylor, John [1993]. *Macroeconomic Policy in a World Economy: From Economic Design to Practical Operation*, New York, NY: Norton.
- [46] Tesar, Linda [1993]. "International Risk Sharing and Non-Traded Goods," *Journal of International Economics* 35 (1-2), 69-89.
- [47] Uhlig, H. [2003], "Do technology shocks lead to a fall in total hours worked?", *Journal of the European Economic Association*, Vol. 2, No. 2-3, pp. 361-371.
- [48] Whalley, John [1985]. *Trade Liberalization Among Major World Trading Areas*, Cambridge, MA: MIT Press.

A Data Sources

This appendix describes the data used in this paper. The complete dataset is available from the authors upon request and covers the period 1970 to 2001, unless otherwise stated.

To calibrate the process of the shocks for the Home country labor productivity in tradables and nontradables we use the annual BLS series “Index of output per hour in manufacturing” and “Index of output per hour in private services,” respectively. For the Foreign country we use an aggregation of the index of manufacturing output and output in services divided by sectoral total employment for an aggregate of OECD countries (Canada, Japan, EU-15) obtained from the OECD STAN sectoral database.

U.S. GDP, consumption and investment are annual chain-weighted 1996 dollar NIPA series from the BEA. World GDP, consumption and investment are annual constant 1995 PPP dollar series for Japan, Canada and EU-15 from the OECD Quarterly National Accounts. The U.S. labor input is the “Index of total hours in the non-farm business sector” from the BLS, while world labor input is aggregate employment for Japan, Canada and EU-15 from the OECD.

The series for U.S. imports and exports at current and constant prices are annual NIPA series from the BEA. The series for the U.S. real exchange rate is a trade-weighted measure of the real value of the dollar computed by J.P. Morgan vis-à-vis the main U.S. trading partners; the series for the U.S. (ex-oil) terms of trade is the ratio of the NIPA (non-oil) import price deflator over the export price deflator from the BEA. The relative price of nontradables in terms of tradables is computed as the ratio of the services CPI over the commodities CPI. Again all this are annual series.

In the estimation of the VAR models for the series on world labor productivity (quarterly) and total factor productivity (annual) we use the ratio between aggregate GDP and labor input for Japan, Canada and EU-15, and the index of TFP in the aggregate OECD countries from the OECD, respectively. In the quarterly VAR, the series for GDP, consumption, net exports, real exchange rate and terms of trade are the quarterly counterpart of the annual series described above.

Table 1: Correlations between real exchange rates and relative consumptions^a

| Country | Correlation | | | |
|---------------------|---------------|--------------|------------------|--------------|
| | HP-Filtered | | First-Difference | |
| | U.S. | OECD | U.S. | OECD |
| Australia | -0.01 | 0.05 | -0.09 | -0.13 |
| Austria | -0.35 | -0.54 | -0.20 | -0.30 |
| Belgium | -0.12 | 0.15 | -0.11 | 0.19 |
| Canada | -0.41 | -0.10 | -0.20 | 0.02 |
| Denmark | -0.16 | -0.27 | -0.20 | -0.21 |
| E.U. | -0.30 | -0.10 | -0.23 | -0.04 |
| Finland | -0.27 | -0.64 | -0.40 | -0.55 |
| France | -0.18 | 0.12 | -0.21 | -0.01 |
| Germany | -0.27 | -0.17 | -0.13 | 0.01 |
| Italy | -0.26 | -0.51 | -0.27 | -0.31 |
| Japan | 0.09 | 0.27 | 0.04 | 0.08 |
| South Korea | -0.73 | -0.50 | -0.79 | -0.63 |
| Mexico | -0.73 | -0.77 | -0.68 | -0.74 |
| Netherlands | -0.41 | -0.20 | -0.30 | -0.19 |
| New Zealand | -0.25 | -0.37 | -0.27 | -0.28 |
| Portugal | -0.56 | -0.73 | -0.48 | -0.67 |
| Sweden | -0.52 | -0.39 | -0.34 | -0.29 |
| Spain | -0.60 | -0.66 | -0.41 | -0.38 |
| Switzerland | 0.16 | 0.53 | 0.09 | 0.32 |
| Turkey | -0.31 | -0.25 | -0.34 | -0.17 |
| U.K. | -0.47 | -0.08 | -0.40 | -0.04 |
| U.S. | N/A | -0.30 | N/A | -0.31 |
| Median ^b | -0.30 | -0.27 | -0.27 | -0.21 |
| | (-0.12,-0.56) | (0.12,-0.54) | (-0.11,-0.41) | (0.02,-0.55) |

^aConsumption and bilateral and effective real exchange rates are annual series from the OECD Main Economic Indicators dataset, from 1973 to 2001. ^bIn parenthesis the cross-sectional 68 percent confidence interval.

Table 2. Parameter values

Benchmark Model

Preferences and Technology

| | |
|-------------------------------------|-------------------------------------|
| Risk aversion | $\sigma = 2$ |
| Consumption share | $\alpha = 0.34$ |
| Elasticity of substitution between: | |
| Home and Foreign traded goods | $\frac{1}{1-\rho} = \{0.95, 1.14\}$ |
| traded and non-traded goods | $\frac{1}{1-\phi} = 0.74$ |
| Share of Home Traded goods | $a_H = 0.72$ |
| Share of non-traded goods | $a_N = 0.45$ |
| Elasticity of the discount factor | |
| with respect to C and L | $\psi = 0.08$ |
| Distribution Margin | $\eta = 1.09$ |
| Labor Share in Tradables | $\xi = 0.61$ |
| Labor Share in Nontradables | $\zeta = 0.56$ |
| Depreciation Rate | $\delta = 0.10$ |

Productivity Shocks

$$\lambda = \begin{bmatrix} 0.820 & 0.001 & 0.068 & 0.180 \\ 0.001 & 0.820 & 0.180 & 0.068 \\ -0.003 & 0.033 & 0.961 & -0.011 \\ 0.033 & -0.003 & -0.011 & 0.961 \end{bmatrix}$$

Variance-Covariance Matrix (in percent)

$$V(u) = \begin{bmatrix} 0.057 & 0.028 & 0.010 & 0.008 \\ 0.028 & 0.057 & 0.008 & 0.010 \\ 0.010 & 0.008 & 0.008 & 0 \\ 0.008 & 0.010 & 0 & 0.008 \end{bmatrix}$$

Table 3. Exchange rates and prices in the theoretical economies^a

| Statistics | Data | Variations on the benchmark economy | | | | | | | | | | | |
|---|-------|-------------------------------------|-----------------|----------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|---------------|------------------|-------------------|
| | | Benchmark Economy | | Arrow-Debreu Economy | No Spillover | | CES Investment | | No Distribution | | | Taste Shocks (I) | Taste Shocks (II) |
| | | $\omega = 0.95$ | $\omega = 1.14$ | $\omega = 0.95$ | $\omega = 0.93$ | $\omega = 1.17$ | $\omega = 0.50$ | $\omega = 0.75$ | $\omega = 0.27$ | $\omega = 0.44$ | $\omega = 10$ | $\omega = 0.95$ | $\omega = 0.82$ |
| <i>Standard deviation relative to GDP</i> | | | | | | | | | | | | | |
| Real exchange rate | 3.28 | 3.28 | 3.28 | 0.79 | 3.28 | 3.28 | 3.28 | 3.28 | 3.28 | 3.28 | 1.06 | 1.26 | 3.28 |
| Terms of trade | 1.79 | 2.84 | 4.47 | 0.78 | 3.18 | 4.20 | 2.53 | 4.41 | 2.68 | 4.92 | 0.26 | 1.38 | 3.39 |
| <i>absolute</i> | | | | | | | | | | | | | |
| Relative price of nontradables | 1.73 | 1.41 | 1.13 | 0.91 | 1.44 | 1.25 | 1.10 | 0.83 | 2.83 | 2.07 | 1.95 | 0.95 | 1.41 |
| <i>Cross-correlations</i> | | | | | | | | | | | | | |
| Between real exchange rate and | | | | | | | | | | | | | |
| Relative GDPs | -0.23 | -0.94 | 0.93 | -0.14 | -0.99 | 0.84 | -0.79 | 0.89 | -0.57 | 0.68 | -0.54 | -0.56 | -0.87 |
| Relative consumptions | -0.45 | -0.48 | -0.45 | 0.98 | -0.65 | -0.34 | -0.11 | 0.18 | -0.22 | -0.81 | 0.84 | -0.34 | -0.73 |
| Net exports | 0.39 | 0.95 | 0.95 | -0.64 | 0.94 | 0.95 | 0.92 | 0.97 | 0.98 | 0.99 | 0.63 | 0.69 | 0.94 |
| Terms of trade | 0.60 | 0.98 | 0.97 | 0.02 | 0.96 | 0.96 | 0.99 | 0.99 | 0.99 | 0.99 | 0.53 | 0.71 | 0.96 |
| Between terms of trade and | | | | | | | | | | | | | |
| Relative GDPs | -0.20 | -0.98 | 0.98 | 0.89 | -0.94 | 0.95 | -0.82 | 0.86 | -0.52 | 0.70 | 0.25 | -0.19 | -0.82 |
| Relative consumptions | -0.53 | -0.63 | -0.66 | 0.18 | -0.81 | -0.56 | -0.19 | 0.06 | -0.21 | -0.85 | 0.88 | -0.64 | -0.86 |
| Net exports | 0.43 | 0.98 | 0.99 | 0.73 | 0.99 | 0.99 | 0.95 | 0.99 | 0.98 | 0.99 | 0.99 | 0.92 | 0.99 |

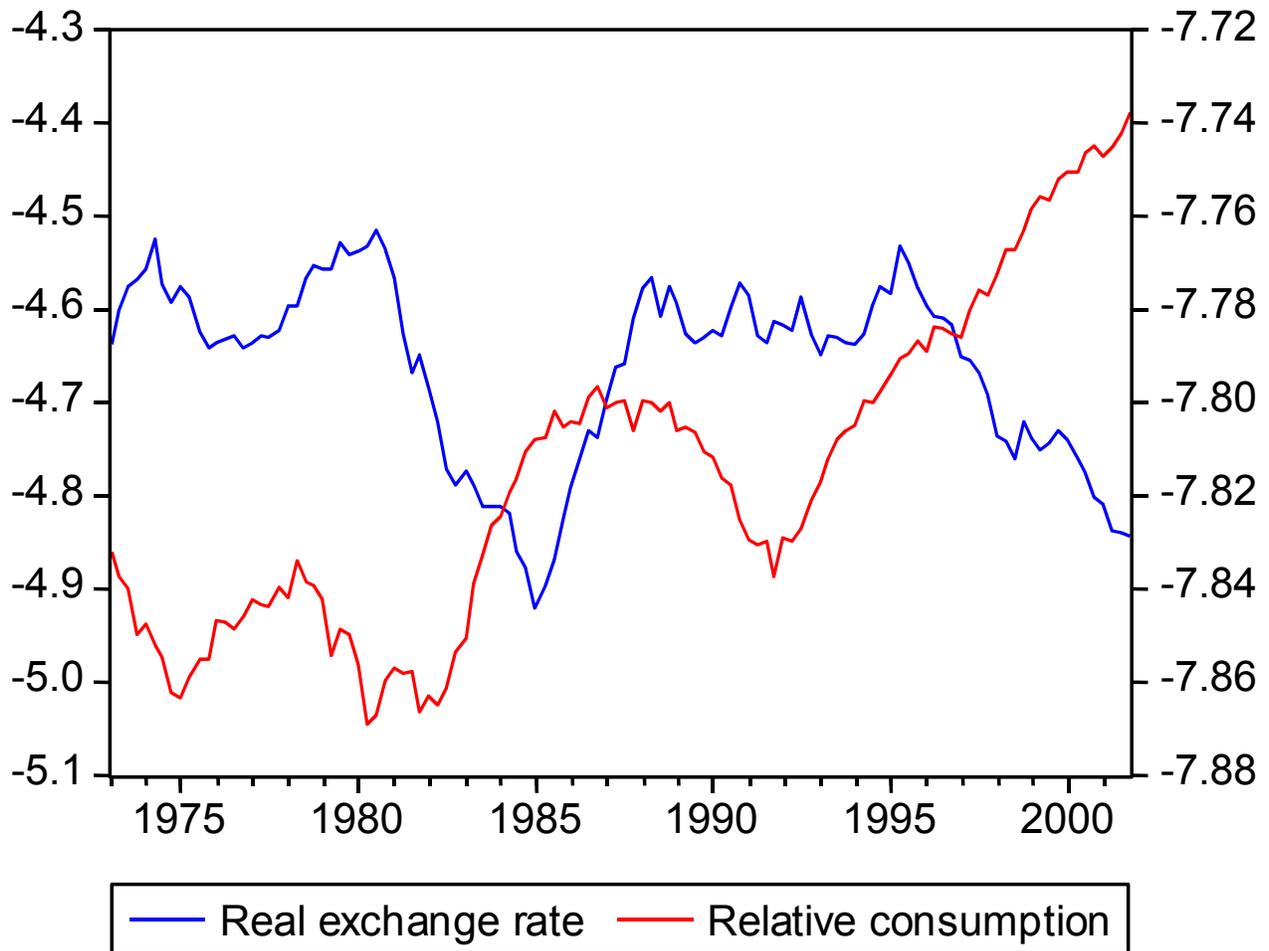
^a $\omega = \frac{1}{1-\rho}$ denotes the elasticity of substitution between Home and Foreign traded goods. The data reported under the heading "Data" are those of the U.S. vis-à-vis the rest of the OECD countries. Simulation results of the Arrow-Debreu economy and the benchmark economy with taste shocks are reported under Taste Shocks (I) and (II), respectively. See the text for a description of the variations on the benchmark economy.

Table 4. Business cycle statistics in the theoretical economies^a

| Statistics | Data | Variations on the benchmark economy | | | | | | | | | | | |
|---|-------|-------------------------------------|-----------------|----------------------|-----------------|-----------------|----------------|-----------------|-----------------|-----------------|---------------|------------------|-------------------|
| | | Benchmark Economy | | Arrow-Debreu Economy | No Spillover | | CES Investment | | No Distribution | | | Taste Shocks (I) | Taste Shocks (II) |
| | | $\omega = 0.95$ | $\omega = 1.14$ | $\omega = 0.95$ | $\omega = 0.93$ | $\omega = 1.17$ | $\omega = 0.5$ | $\omega = 0.75$ | $\omega = 0.27$ | $\omega = 0.44$ | $\omega = 10$ | $\omega = 0.95$ | $\omega = 0.82$ |
| <i>Standard deviation relative to GDP</i> | | | | | | | | | | | | | |
| Consumption | 0.92 | 0.54 | 0.55 | 0.49 | 0.56 | 0.46 | 0.50 | 0.49 | 0.47 | 0.67 | 0.44 | 0.59 | 0.68 |
| Investment | 4.25 | 3.13 | 3.14 | 3.14 | 3.28 | 3.28 | 3.15 | 2.62 | 3.15 | 3.17 | 3.28 | 2.93 | 2.91 |
| Employment | 1.09 | 0.49 | 0.51 | 0.49 | 0.54 | 0.55 | 0.45 | 0.39 | 0.49 | 0.51 | 0.54 | 1.00 | 1.00 |
| <i>absolute</i> | | | | | | | | | | | | | |
| Import ratio | 4.13 | 2.26 | 4.33 | 0.62 | 2.67 | 4.47 | 1.02 | 2.59 | 1.23 | 3.67 | 4.70 | 1.20 | 2.60 |
| Net exports over GDP | 0.63 | 0.25 | 0.38 | 0.04 | 0.30 | 0.38 | 0.06 | 0.11 | 0.23 | 0.41 | 0.33 | 0.16 | 0.32 |
| <i>Cross-correlations</i> | | | | | | | | | | | | | |
| Between foreign and domestic | | | | | | | | | | | | | |
| GDP | 0.49 | 0.42 | 0.40 | 0.43 | 0.52 | 0.50 | 0.40 | 0.49 | 0.43 | 0.40 | 0.28 | 0.31 | 0.29 |
| Consumption | 0.32 | 0.10 | 0.10 | 0.37 | -0.22 | 0.12 | 0.33 | 0.54 | 0.45 | -0.25 | 0.48 | -0.001 | -0.23 |
| Investment | 0.08 | 0.35 | 0.31 | 0.34 | 0.51 | 0.49 | 0.04 | 0.59 | 0.33 | 0.29 | 0.08 | 0.28 | 0.28 |
| Employment | 0.32 | 0.27 | 0.19 | 0.28 | 0.53 | 0.46 | 0.13 | 0.62 | 0.29 | 0.21 | -0.02 | 0.08 | 0.05 |
| Between net exports and GDP | -0.51 | -0.51 | 0.52 | 0.34 | -0.45 | 0.48 | -0.49 | 0.38 | -0.32 | 0.32 | -0.15 | -0.23 | -0.46 |

^a $\omega = \frac{1}{1-\rho}$ denotes the elasticity of substitution between Home and Foreign traded goods. The data reported under the heading "Data" are those of the U.S. vis-à-vis the rest of the OECD countries. Simulation results of the Arrow-Debreu economy and the benchmark economy with taste shocks are reported under Taste Shocks (I) and (II), respectively. See the text for a description of the variations on the benchmark economy.

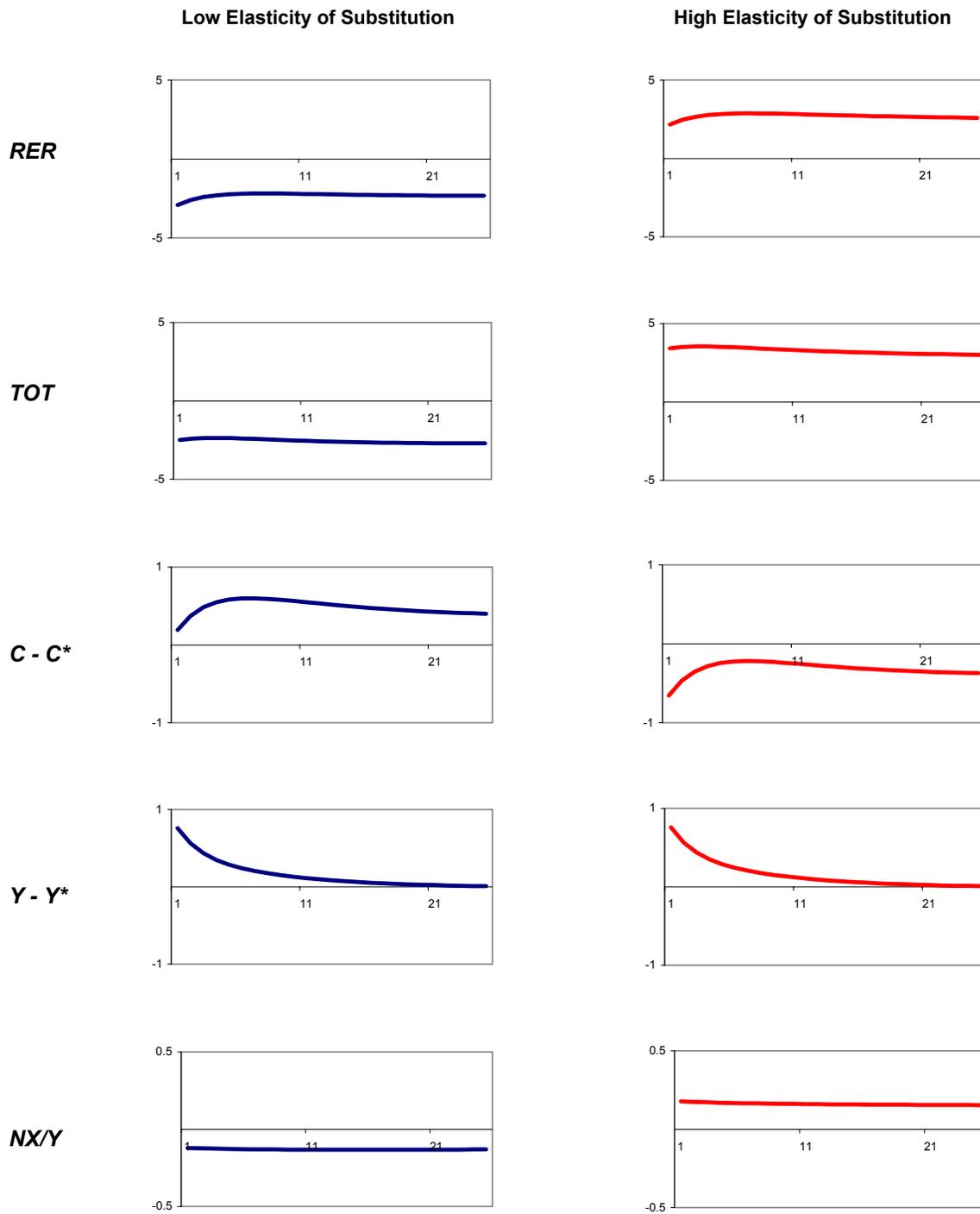
Figure 1 U.S. Real exchange rate and relative consumption



The real exchange rate is eP^*/P , where the nominal exchange rate e is the U.S. dollar price of a basket of OECD currencies, P^* is an aggregate of OECD CPIs, and P is the U.S. CPI. See the Appendix for the sources.

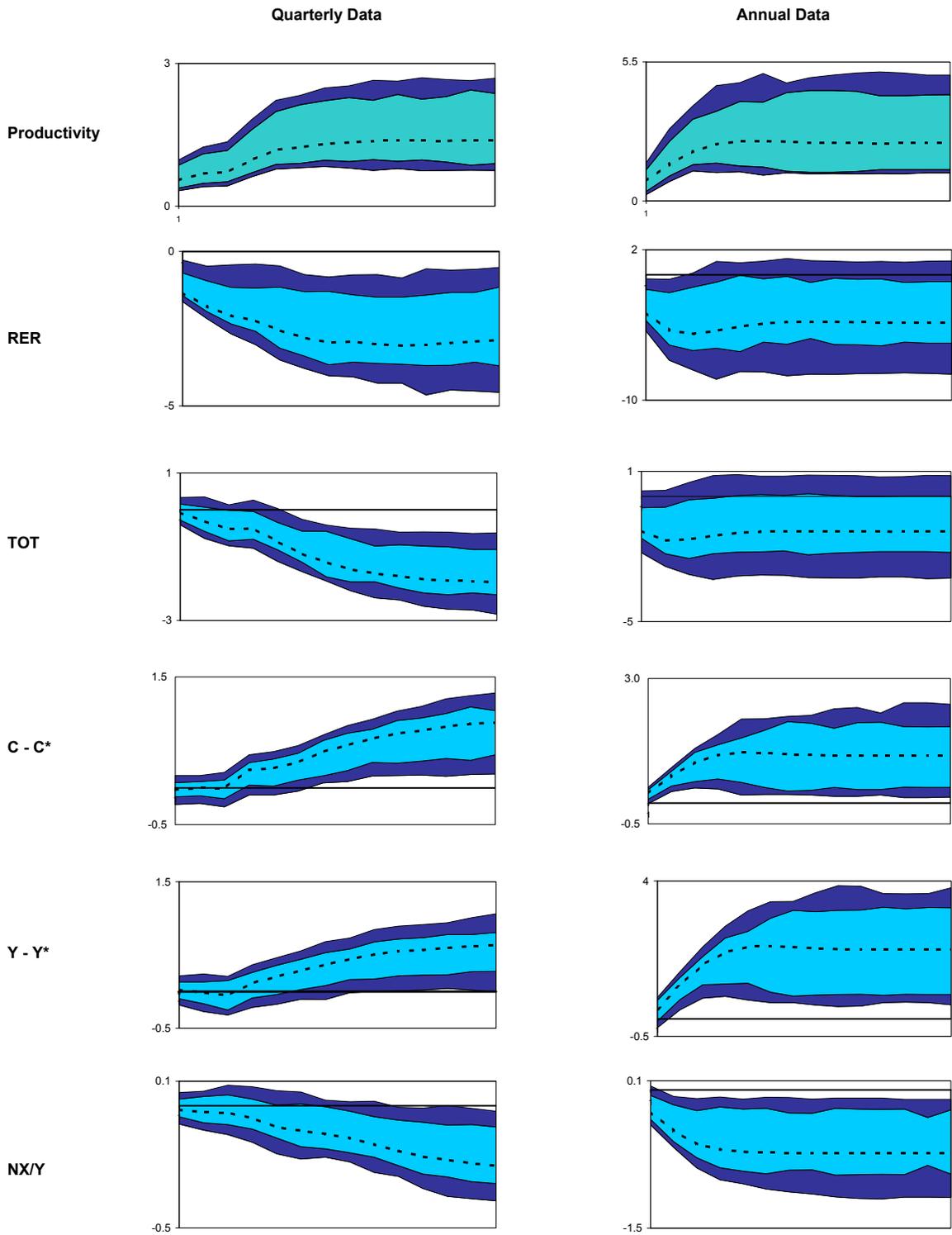
Figure 2

Theoretical Responses to a Technology Shock in the Traded-Goods Sector



All series are in percent.

Figure 3
Impulse Responses to a Technology Shock in the Traded-Goods Sector



The first column describes the responses from a 6-variable VAR, using quarterly data. The variables are labor productivity, the real exchange rate, the terms of trade, relative consumption, relative output, and net exports. The second column shows the responses from a 4-variable VAR, using annual data. The variables are TFP, relative consumption, relative output, and, alternatively, the real exchange rate, the terms of trade, and net exports. All series are in percent.