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DP15275

**MACROPRUDENTIAL POLICY,
MORTGAGE CYCLES AND
DISTRIBUTIONAL EFFECTS: EVIDENCE
FROM THE UK**

José Luis Peydró, Jagdish Tripathy, Francesc
Rodríguez Tous and Arzu Uluc

INTERNATIONAL MACROECONOMICS AND FINANCE



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Discussion Paper DP15275
Published 10 September 2020
Submitted 08 June 2020

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www.cepr.org

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Abstract

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JEL Classification: E5, G01, G21, G28, G51

Keywords: macroprudential policy, Mortgages, Credit cycles, inequality, House Prices

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Acknowledgements

We are thankful to David Aikman, João Cocco, Andy Haldane, David Jaume, Anil Kashyap, Benjamin Keys, Frederic Malherbe, Ricardo Reis, Sagar Shah, Vahid Saadi, Rhiannon Sowerbutts, Philip Strahan, Misa Tanaka, Neeltje Van Horen, Stijn Van Nieuwerburgh, Matthew Waldron and John O. S. Wilson for insightful comments. We are also grateful to conference participants at the CEPR-Bank of Slovenia-European Central Banking Network Conference on Evaluating the Effectiveness of Macroprudential Policies, EEA, EFA, Annual Workshop of ECB-ESCB at Bank of Greece, 3rd Conference on Financial Stability (Banco de México), Alternative Datasets Conference at Bocconi University, FINEST Workshop at Bristol University, Workshop on Banking and Financial Intermediation at Middlesex University, and seminar participants at the Bank of England, Bundesbank, Cass Business

School, and CPB Netherlands Bureau for Economic Policy Analysis. The views expressed in this article are without exception those of the authors; they are not meant to represent the views of the Bank of England, members of its Financial or Monetary Policy Committee or any other institutions the authors are affiliated to. This project has received funding from the European Research Council (ERC) under the European Union's Horizon 2020 research and innovation programme (grant agreement No 648398). Peydró acknowledges financial support from the ECO2015-68182-P (MINECO/FEDER, UE) grant and the Spanish Ministry of Economy and Competitiveness, through the Severo Ochoa Programme for Centres of Excellence in R&D (SEV-20150563) and a 2018 Leonardo Grant for Researchers and Cultural Creators, BBVA Foundation. José-Luis Peydró is with Imperial College London, UPF-CREI-ICREA-BarcelonaGSE; Francesc Rodriguez-Tous is with Cass Business School; Jagdish Tripathy is with Bank of England (email: jagdish.tripathy@bankofengland.co.uk); Arzu Uluc is with Bank of England.

Macroprudential Policy, Mortgage Cycles and Distributional Effects: Evidence from the UK*

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Abstract

Macroprudential regulators worldwide have introduced regulations to limit household leverage in light of existing evidence which suggests that high leverage is associated with household distress during crisis. We analyse the distributional effects of such a macroprudential policy on mortgage and house price cycles. For identification, we exploit the universe of UK mortgages and a 15%-limit imposed in 2014 on lenders—not households—for high loan-to-income ratio (LTI) mortgages. Despite some regulatory arbitrage (e.g. increases in LTV and average loan size), more-constrained lenders issue fewer high-LTI mortgages. Partial substitution by less-constrained lenders leads to overall credit contraction to low-income borrowers in local-areas more exposed to constrained-lenders, lowering house price growth. Following the Brexit referendum (which led to house-price correction), the 2014-policy strongly implies—via lower pre-correction debt—better house prices and mortgage defaults during an episode of house price correction.

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*We are thankful to David Aikman, João Cocco, Andy Haldane, David Jaume, Anil Kashyap, Benjamin Keys, Frederic Malherbe, Ricardo Reis, Sagar Shah, Vahid Saadi, Rhiannon Sowerbutts, Philip Strahan, Misa Tanaka, Neeltje Van Horen, Stijn Van Nieuwerburgh, Matthew Waldron and John O. S. Wilson for insightful comments. We are also grateful to conference participants at the CEPR-Bank of Slovenia-European Central Banking Network Conference on Evaluating the Effectiveness of Macroprudential Policies, EEA, EFA, Annual Workshop of ECB-ESCB at Bank of Greece, 3rd Conference on Financial Stability (Banco de México), Alternative Datasets Conference at Bocconi University, FINEST Workshop at Bristol University, Workshop on Banking and Financial Intermediation at Middlesex University, and seminar participants at the Bank of England, Bundesbank, Cass Business School, and CPB Netherlands Bureau for Economic Policy Analysis. The views expressed in this article are without exception those of the authors; they are not meant to represent the views of the Bank of England, members of its Financial or Monetary Policy Committee or any other institutions the authors are affiliated to. This project has received funding from the European Research Council (ERC) under the European Union's Horizon 2020 research and innovation programme (grant agreement No 648398). Peydró acknowledges financial support from the ECO2015-68182-P (MINECO/FEDER, UE) grant and the Spanish Ministry of Economy and Competitiveness, through the Severo Ochoa Programme for Centres of Excellence in R&D (SEV-20150563) and a 2018 Leonardo Grant for Researchers and Cultural Creators, BBVA Foundation. José-Luis Peydró is with Imperial College London, UPF-CREI-ICREA-BarcelonaGSE; Francesc Rodriguez-Tous is with Cass Business School; Jagdish Tripathy is with Bank of England (email: jagdish.tripathy@bankofengland.co.uk); Arzu Uluc is with Bank of England.

1 Introduction

Household leverage has received intense academic scrutiny as a source of financial instability, thereby causing long and deep recessions. The empirical literature provides evidence that strong mortgage expansion to households (crucially low-income ones, [Mian and Sufi \(2009\)](#)) was the underlying cause of the 2007-09 US subprime/global financial crisis—as well as previous financial crises in history—with associated high mortgage defaults, house price contractions, and overall negative real effects ([Schularick and Taylor \(2012\)](#); [Favara and Imbs \(2015\)](#); [Mian et al. \(2017\)](#); [Di Maggio and Kermani \(2017\)](#); [Piskorski and Seru \(2018\)](#)). Theoretical literature shows settings in which unrestricted lending to households can lead to housing booms and to negative externalities during financial crises, with strong negative effects on house prices (fire sales) and bank defaults; thereby advocating macroprudential policies to limit pre-crisis household leverage ([Lorenzoni \(2008\)](#); [Corbae and Quintin \(2015\)](#); [Farhi and Werning \(2016\)](#); [Korinek and Simsek \(2016\)](#); [Favilukis et al. \(2017\)](#); [Davila and Korinek \(2017\)](#); [Greenwald \(2017\)](#)). Consistent with this view, policymakers across the world have introduced macroprudential regulations on household leverage, especially after the global financial crisis.¹ Despite the academic and policy importance of macroprudential policy on household leverage, the empirical literature is scant.

In this paper, we analyse the distributional effects of macroprudential policy on mortgage and house price cycles. For empirical identification, we exploit the universe of UK mortgages in conjunction with a 15% cap on the proportion of new mortgages with high loan-to-income ratios ($LTI \geq 4.5$) imposed in 2014 on UK mortgage lenders. These lenders are differentially affected based on their pre-policy share of high LTI loans in their mortgage portfolio.

In brief, our robust results show that more-constrained lenders issue fewer high-LTI mortgages, with stronger effects for low-income households (despite some regulatory arbitrage). Even though the regulation is on lenders (and not households), substitution by less-constrained lenders is only partial and leads to intensive and extensive margin contraction to low-income households in areas more exposed to the more-constrained lenders—with associated dampening of house price growth. Following the Brexit referendum, which led to a house price correction across the UK, the 2014 macroprudential policy leads to lower house price drops and mortgage default rates. The better house price and default rate effects are fully driven by the reduction in household debt due to the policy, and not by endogenous changes in debt orthogonal to the policy.

¹E.g., regulators in Denmark, Hong Kong, Ireland, Israel, Netherlands, Sweden, and the US (among others) have introduced restrictions on household leverage. A complete list of jurisdictions that have regulated household leverage is available in the iMaPP database (IMF). See Online Appendix A for a discussion.

Our main contribution to the literature is to analyse the impact of household leverage regulation on credit and house price cycles, as well as its distributional effects (with respect to household income),² thereby showing costs and benefits of macroprudential policy. To the best of our knowledge, there is currently no paper which studies the effects of changes in macroprudential policy on household leverage during both *boom and correction* periods. This is not only due to the scarcity of administrative data on mortgages, but also due to relative recency of such regulation because of which there have not been negative overall shocks to check the potential benefits of limiting household leverage during a boom. However, the UK offers an ideal setting as it has administrative mortgage data, introduced macroprudential policy on LTI in 2014, and had a drop in house prices stemming from the result of the so-called Brexit referendum held in June 2016.³ Our results show that the policy strongly restricts credit to low-income households in the boom even though the LTI regulation is on banks and not households. However, this distributional effect of the policy on leverage leads to lower mortgage default and substantially lower house price cooling during the correction. In addition, we offer novel findings on regulatory arbitrage.

In the remainder of this Section, we provide a detailed preview of the paper and a detailed discussion of the literature and its contrast with our paper.

Preview of the paper. The macroprudential measure limiting high LTI lending was announced by the UK regulator in June 2014. The measure imposed a 15% limit on the proportion of total number—not total £ amount—of new mortgages with $LTI \geq 4.5$ that a UK lender could issue quarterly starting 2014Q4.⁴ The regulation is applicable on residential mortgages issued by lenders, i.e. mortgages issued to first-time-buyers, home-movers and refinancers where the borrower lives in the underlying property.

We exploit the universe of loan-level data on residential mortgages issued in the UK during 2012Q3-2016Q2 in conjunction with the LTI regulation of June 2014. The loan-level

²We do not interpret this result as a comment on the overall welfare consequences of the regulation since any contraction in lending to specific borrowers resulting from the regulation must be weighed against the negative consequences of a crisis on such borrowers. This is because high household indebtedness is often associated with negative consumption response during economic distress, as has been shown by a range of studies using micro-data from Denmark (Andersen et al. (2016)), the UK (Bunn and Rostom (2014)) and the US (Mian and Sufi (2009) and Baker (2018)). These studies report a negative correlation between pre-crisis levels of household indebtedness and consumption growth during a subsequent crisis. Particularly for the UK, Bunn and Rostom (2014) estimate that *cuts in spending associated with debt...reduced aggregate private consumption by around 2% after 2007, which increased the depth of the recession.*

³See Figure 1 on house prices. Moreover, a growing literature on the effects of the Brexit referendum has shown a slowdown in UK growth and firm investment (Born et al. (2019), Faccini and Palombo (2019)), productivity (Bloom et al. (2019), Broadbent et al. (2019)), syndicated loan markets (Berg et al. (2019)), and stock returns (Davies and Studnicka (2018)) following the referendum.

⁴There was evidence of deteriorating underwriting standards in the UK mortgage market before 2014, with a sharp increase in highly leveraged mortgages. For instance, the share of mortgages with $LTI \geq 4.5$ had been around 10% in 2013; the corresponding figure for the pre-crisis period (2005-2007) was just 6.5%.

data, reported quarterly, includes information on borrower characteristics such as income, age and employment status, on the issuing lender, and on characteristics of the loan itself, such as loan amount, interest rate, interest rate type (fixed or floating), mortgage term, property value and postcode. We merge the loan-level data with lenders' balance-sheet data on credit, size, funding, and liquidity status at a quarterly frequency. In addition, we match mortgage data with house price growth and mortgage non-performance rates (defaults) at a local-area level between 2012Q3 and 2018Q2 to study these variables before and after the Brexit referendum in mid-2016. Therefore, we create a rich dataset that allows us to exploit time, borrower, lender, and location characteristics to study the effects of the policy.

In addition to the loan data before and after the LTI regulation, our research design hinges on the restriction on high LTI lending being binding at a lender level. This provides a variation across UK lenders based on the share of loans with $LTI \geq 4.5$ in their mortgage portfolio a year before the announcement of the regulation (i.e. 2013Q3-2014Q2).⁵ We use these shares to classify UK lenders into two groups of more-constrained and less-constrained lenders.⁶ Constrained lenders issued roughly a third of all mortgages in our sample. The constrained and unconstrained lenders are similar across a range of balance sheet characteristics such as size, liquidity ratio, core funding ratio, return on assets, CET1 ratio, and share of household credit. We use a difference-in-differences specification to make causal inferences on the changes in the mortgage portfolio of the constrained lenders in response to the regulation.

Our results show that the introduction of the regulation affects mortgage lending. First, after the macroprudential regulation and in comparison to unconstrained lenders, constrained lenders cut their share of mortgages with $LTI \geq 4.5$ by 3.9%. Moreover, the fewer high LTI loans issued by constrained lenders after the regulation contain 5.2% fewer loans issued to borrowers in the lowest income-quintile. That is, the policy doubly-affects low-income households through the reduction in high LTI loans (which they are more likely to borrow) *and* the drop in their presence *within* high LTI loans issued by constrained lenders.⁷ Importantly, all the key effects happen *only after* the regulation is implemented and not before its announcement (i.e., parallel trends), which serves as a placebo test.

Second, we analyse credit substitution since the regulation is binding on UK lenders

⁵UK lenders under the purview of the regulation includes banks and building societies which account for 99% and 98% of the total number and total £ value of mortgage lending in the UK, respectively.

⁶Our baseline specification distinguishes more-constrained and less-constrained lenders using an indicator variable. We also refer to these groups as constrained and unconstrained groups, respectively, for brevity.

⁷Results are robust to alternate thresholds for low-income borrowers, for e.g. those based on income-quartiles or terciles. There is no concomitant drop in the share of high LTI loans issued to the lowest age-quintile, which shows that the contraction in high LTI loans by constrained lenders had a stronger effect on low-income households than for younger households.

and not on households: borrowers who could not receive high LTI loans from constrained banks (i.e. low-income borrowers) could either receive a loan with $LTI < 4.5$ from those banks or obtain a high LTI loan from an unconstrained bank. As these two substitution channels are statistically significant at the loan-level, we aggregate mortgage lending by time (quarterly), location (UK local areas equivalent to boroughs) and income-quintiles to study the combined effect of these substitution channels on overall lending to low-income borrowers. We use a triple difference-in-differences specification to check whether the regulation has an overall effect on lending (in total number and total £ value of loans) to borrowers in the lowest income-quantile in the UK, especially in local areas more exposed to constrained banks. Estimated coefficients show that borrowers in the bottom income-quintile in areas with an ex-ante one standard deviation higher exposure to constrained lenders (equivalent to a 6.6% higher share of those banks) experience a 9.0% decline in the total number of mortgages and a 11.4% decline in the total £ value of lending after the introduction of the regulation.⁸ We estimate a lower-bound for these coefficients accounting for unobserved local-area heterogeneity using the techniques described in [Altonji et al. \(2005\)](#) and [Oster \(2017\)](#); the estimated lower-bounds are remarkably close to the actual coefficients.

Third, the contraction in high LTI lending due to the regulation affects house price dynamics. We use house prices from the universe of housing transactions to show that the introduction of the regulation implies a drop in the growth rate of the 25th percentile of house price transactions in local areas more affected by the regulation (i.e., local areas more exposed to constrained lenders *and* with a higher concentration of low-income borrowers).

Fourth, we study the effect of the Brexit referendum (June 2016) which led to a drop in house price growth across the UK. [Figure 1](#) shows that growth in median house price in the UK fell from a peak of 7.5% annually just before the referendum to 2.7% two years after the referendum. We find that local areas more affected by the 2014-policy (i.e., local areas with a higher proportion of constrained lenders and low-income borrowers) have a relatively lower cooling in house price growth after the referendum. The relatively lower decline in house price growth post-Brexit referendum is stronger for lower quantiles of house price transactions. The stronger effect on the lower percentiles of house price transactions is consistent with the ex-ante (to the Brexit referendum) reduction in debt to low-income borrowers.⁹ Moreover, these local areas also experience relatively lower rates of mortgages

⁸These effects are estimated controlling for local area×time, income-band×time and local area×income-band fixed-effects. Therefore, the contraction in lending to low-income borrowers in high exposure areas is unlikely to be driven by demand shocks specific to local areas or low-income borrowers.

⁹The house price trends are estimated in the presence of region×time fixed-effects and, therefore, do not result from regional house price trends unrelated to the regulation; i.e., the house price trends described are between local areas with different exposures to the 2014 regulation in the same region in the UK.

with a payment shortfall after the Brexit referendum.¹⁰ We show that the underlying mechanism linking these results to the policy is the reduction in debt to low-income borrowers in affected local areas driven by the policy. Furthermore, the link between endogenous changes in debt orthogonal to the policy, and house prices and defaults following the referendum is economically and statistically insignificant.

Figure 1: Median house-price growth



The above figure shows the annual growth rate of median house prices in England and Wales. The vertical dashed line denotes the date of the Brexit referendum. The data is from the House Price Statistics for Small Areas (HPSSAs) Dataset 9: Median price paid for administrative geographies.

Finally, we analyse regulatory arbitrage after the introduction of the regulation. First, as the regulation is on the proportion of loans and not the total value, the lower number of high LTI loans issued by the constrained lenders after the regulation are, on average, larger (by 9%) and with higher loan-to-value ratios (by 1.8 pp). The overall LTI distribution of mortgages issued by constrained banks moves significantly to the right after the regulation, with a marked increase in the mass of mortgages with LTI just below 4.5 (i.e. bunching). Moreover, following the counterfactual distribution estimation technique by [Chernozhukov et al. \(2013\)](#) and using the Oaxaca-Blinder decomposition, results suggest that the changes in the LTI distribution are driven by the regulation. The decomposition shows that while the rightward shift of the LTI distribution for constrained banks is almost entirely attributable to the regulation, the more marginal changes in the LTI distribution for unconstrained banks is instead explained by changes in covariates.¹¹

Literature Review. Our main contribution to the literature is to show the impact of household leverage regulation on boom/correction cycles and its distributional effects (with

¹⁰Mortgages with payment shortfalls have, in general, arrears for less than 6 months (not foreclosures, as loans are with full recourse in the UK). Nevertheless, banks must provision for these delinquencies, implying higher NPLs and lower profits and capital. Mortgage non-performance, therefore, is costly for both banks and borrowers (due to worse rating and credit history for the latter).

¹¹In other words, while constrained banks issue a different LTI post-regulation controlling for borrower characteristics, unconstrained banks issue a similar LTI for those characteristics irrespective of the regulation.

respect to household income), thereby showing costs and benefits of macroprudential policy. In addition, we also provide novel findings on regulatory arbitrage.

To the best of our knowledge, there is currently no paper which studies the effects of changes in macroprudential policy on household leverage during both boom and correction periods. This is not only due to scarcity of administrative data on mortgages (most credit registers are on corporate loans), but also due to relative recency of such regulation because of which there have not been negative overall shocks to check the potential benefits of limiting household leverage during a boom. The UK offers an ideal setting as it has administrative data on the universe of issued mortgages, a macroprudential policy which restricted LTI in 2014, and a cooling in house prices stemming from the result of the Brexit referendum. We find that the policy restricts credit to low-income households during the boom even though the LTI regulation is on banks and not households, an effect which leads to lower mortgage defaults and lower house price drops during the correction. Other papers studying the effects of policies restricting household leverage (e.g. [DeFusco et al. \(forthcoming\)](#), [Acharya et al. \(2019\)](#), and [Bekkum et al. \(2019\)](#) based on regulations introduced post-2008 financial crisis in the US, Ireland and Netherlands, respectively) analyse the effects of the macroprudential policy only around its introduction and not during a correction.¹²

By showing the effects during the correction, we test a key prediction of the theoretical models highlighted in the first page of the Introduction, i.e., the reduction in pre-crisis leverage has an impact on house prices and defaults during a correction. Importantly, the underlying channel is via changes in lending to low-income households, which is important to explain the 2007-09 sub-prime crisis in the US ([Mian and Sufi \(2009\)](#); [Mian and Sufi \(2014\)](#)).

As compared to the regulations in other jurisdictions, measures in the UK set a cap on household leverage at a *lender* level. This feature also distinguishes our study from those by [DeFusco et al. \(forthcoming\)](#), [Acharya et al. \(2019\)](#) and [Bekkum et al. \(2019\)](#). For instance, [DeFusco et al. \(forthcoming\)](#) study the impact of “ability-to-pay” (or affordability) requirements introduced on large mortgages (*jumbo loans*) with debt-service-to-income ratios (DSTI) beyond a certain threshold, with other loans remaining unaffected by the regulation. In contrast to these papers, the key transmission mechanism to lending and house prices in our paper is via change in lender-behaviour (supply) following a macroprudential regulation.

We also contribute to the literature on credit and house prices. Household leverage

¹²[Jiménez et al. \(2017\)](#) study the impact of bank-based macroprudential regulation (countercyclical capital) on corporate loans during a full credit cycle and find no overall effects during boom periods. Our paper finds effects of the regulation during boom periods as well. There are also papers, for instance [Igan and Kang \(2011\)](#) and [Crowe et al. \(2013\)](#), which use aggregate cross-country data to study effects of macroprudential regulation, with the usual limitations in identification.

closely interacts with house prices on the way to affecting the macroeconomy. The literature has emphasized the amplification loop that exists between loosening of lending standards, increase in household leverage and subsequent increases in house prices (for instance [Mayer et al. \(2009\)](#); [Mian and Sufi \(2009\)](#); [Favara and Imbs \(2015\)](#); [Jordà et al. \(2015a\)](#); [Bhutta and Keys \(2016\)](#); [Di Maggio and Kermani \(2017\)](#)). The LTI regulation in the UK, introduced in response to sharp increases in household leverage, resulted in cooling down of house prices in local areas with a higher exposure to lenders more constrained by the regulation and with a higher concentration of low-income borrowers.

Importantly, we show that the macroprudential policy benefits house price dynamics and mortgage defaults after the Brexit referendum (which led to house price cooling across the UK). Lower house price drops can be beneficial given the evidence linking changes in house prices to mortgage default ([Ghent and Kudlyak \(2011\)](#); [Mian and Sufi \(2011\)](#); [Guiso et al. \(2013\)](#); [Fuster and Willen \(2017\)](#); [Bhutta et al. \(2017\)](#)), household consumption ([Campbell and Cocco \(2007\)](#); [Lustig and Van Nieuwerburgh \(2010\)](#); [Mian et al. \(2013\)](#); [Aladangady \(2017\)](#); [DeFusco \(2018\)](#); [Guren et al. \(2018\)](#)), economic activity ([Chaney et al. \(2012\)](#)), and the aggregate demand externality that operates during a financial crisis concurrent with a house price correction ([Mian and Sufi \(2014\)](#)). Thus, our paper is unique in showing the effect of macroprudential policy on smoothing house price cycles, and localized benefits of restricting household leverage during boom periods.

There is also a growing literature which has varyingly attributed the secular rise in income and wealth inequality to taxation (for instance [Piketty and Saez \(2003\)](#)), globalization ([Autor et al. \(2013\)](#)) and automation ([Autor and Dorn \(2013\)](#)). Papers have also studied the redistributive effects of policies, such as those for monetary policy ([Auclert \(2019\)](#)) and for policies aimed at financial access ([Rajan \(2011\)](#); [Agarwal et al. \(2012\)](#)). We contribute to this literature by linking debt, inequality and macroprudential policies. While we find that macroprudential policies aimed at high leverage can disproportionately reduce lending to low-income borrowers (affecting their wealth dynamics since real estate is the key component of household wealth),¹³ they yield benefits—in terms of a lower house price correction and lower mortgage defaults in low-income areas—during a correction.

Regarding regulatory arbitrage, other papers on macroprudential policy have also reported unintended consequences. We find some novel findings given the distinct nature of the 2014-policy (on LTI and on banks). While lenders more-constrained by the regulation reduce their share of mortgages with $LTI \geq 4.5$, they increase the loan-to-value ratios (LTV) and average loan-size of high LTI mortgages. They also disproportionately raise the mass

¹³Property wealth accounts for 60% of total household wealth in the UK. Source: Office of National Statistics, Wealth and Assets Survey 2018.

of mortgages with LTI close to and below the 4.5 threshold, leading to an overall rightward shift in the LTI distribution of mortgages issued post-regulation.

Organization. The rest of the paper is organized as follows. Section 2 provides a background on the 2104 regulation. Section 3 discusses the data and empirical strategy. Section 4 presents the key results in different subsections. Section 5 concludes.

2 Institutional Background

Mortgages are one of the largest asset classes on the UK banks' balance sheets. Therefore, poorly performing mortgages can pose direct risks to the resilience of the UK banking system and to financial stability. Mortgages are also the largest liability on the UK household sector's balance sheet.¹⁴ In the event of a fall in incomes or an increase in interest rates, households may cut back consumption to keep paying their mortgages, which could have indirect effects on the rest of the economy. The macroprudential policy authority in the UK, the Bank of England's Financial Policy Committee (FPC), monitors developments in the housing and mortgage markets in order to mitigate these risks to financial and economic stability.

At its June 2014 meeting, the FPC assessed risks to lenders and the wider economy from an increase in UK household indebtedness. At that time, there were signs that the UK housing market was heating up. There was strong recovery in the housing market, and house prices were rising faster than household income. These developments were associated with a significant increase in the share of high leverage mortgages. For instance, the share of mortgages with loan-to-income ratios (LTI) at or greater than 4.5 rose from 6.5% in the immediate pre-crisis period during 2005-07 to 10% between 2013Q2-2014Q1 (see Online Appendix Figure A.2). Therefore, as insurance against the risk of a loosening in underwriting standards and a significant increase in the number of highly indebted households, the FPC recommended the Prudential Regulation Authority (PRA) and Financial Conduct Authority (FCA) to ensure that mortgage lenders do not extend more than 15% of their total number of new residential mortgages with loan-to-income ratios at or greater than 4.5.^{15,16}

¹⁴Lending to households consistently accounts for roughly half of all the credit issued to the private non-financial sector by banks in the UK (Source: FRED Economic Data). In 2017Q2, mortgages (£1.3 trillion) made up more than 80% of the total stock of household debt (£1.6 trillion) (Bank of England (2017b)).

¹⁵The PRA is responsible for prudential regulation and supervision of banks, building societies, credit unions, insurers and major investment firms in the UK. The FCA is the conduct regulator for financial services firms and financial markets in the UK. Along with the cap on the proportion of mortgages with LTI ≥ 4.5 , the FPC also introduced an affordability test that lenders had to conduct on each mortgage issued by them. See Footnote 22 for a discussion on the effects of the affordability tests.

¹⁶The 4.5 loan-to-income ratio is calibrated to ensure a cap on the mass of borrowers that may face stressed debt-servicing ratios of around 35-40%. This is an inflection point beyond which the proportion of borrowers experiencing repayment difficulties can rise sharply. See Bank of England (2017), page 8 for details.

The PRA and the FCA implemented the FPC’s recommendation as of 1 October 2014. The policy was implemented on a quarterly basis and excludes re-mortgages with no change to the outstanding principal, and lifetime mortgages. The policy applies to mortgage lenders whose annual residential mortgage lending is in excess of £100 million in value and 300 in number of mortgages. According to the policy, a mortgage lender part of a banking group is allowed to allocate all or part of its high LTI allowance to any other regulated entity within that group. When the rules applied as of 1 October 2014, there were 32 banking groups within the scope of the policy. During the period of interest of this study, mortgage lenders inside the scope of the policy account for 99% and 98% percentage of the total number and total £ value of mortgage lending, respectively.¹⁷

3 Data and Empirical Strategy

3.1 Data

We draw upon multiple data sources for our empirical analysis. These are: the PSD001 database which includes the universe of newly issued mortgages in the UK; the PSD007 database which includes the stock of mortgages in the UK; the balance sheet and income statement data from lenders; and the UK land registry for data on housing transactions.

PSD001 (Product Sales Database 001), updated quarterly by the FCA, contains information on the universe of newly issued residential mortgages in the UK.¹⁸ These are the type of mortgages subject to the limit on high loan-to-income (LTI) lending. We exclude from our sample instances of re-mortgaging without a change in the principal of the loan and life-time mortgages since these types of contracts are not in the scope of the policy. We report summary statistics of our sample in Table 1 (A). The data from PSD001 provides a wide range of borrower characteristics, such as borrower’s gross income, age, employment status, and whether the borrower is a home-mover, first-time-buyer or re-mortgagor. Moreover, the database contains information on mortgage characteristics such as loan amount, interest rate, price of the property, LTI, loan-to-value (LTV), term, type of repayment, and location of the property. The data also includes the date of issuance and the issuing bank of the mortgage. Therefore, we can use this data to study changes in the mortgage portfolio of banking groups more constrained by the limit on high LTI lending, as well as whether some borrower cross-sections are disproportionately affected by the regulation.

¹⁷Substitution by mortgage lenders outside the scope of the policy is minimal since they have a small market share and risk coming under the purview of the policy if they cross the £100 million/300 mortgages portfolio threshold.

¹⁸It therefore excludes loans such as second-charged lending, commercial, and buy-to-let mortgages.

PSD007 (Product Sales Database 007) contains information on the stock of outstanding mortgages. This data has been obtained by the Financial Conduct Authority bi-annually starting 2015. The stock data includes information on the performance of each outstanding mortgage in a given period. The performance data is captured in dummies which indicate non-performance because of payment arrears or being under forbearance. The latter refers to mortgages under re-payment difficulties which have received modified terms from lenders in the form of payment suspension, payment reduction, term extension etc. The stock data includes the date of origination, issuing bank, and location of the related property. We use this information to calculate rates of mortgages with payment arrears or mortgages under forbearance across UK local areas over time.

In addition, we have data on lenders' balance sheet position and income statement at a quarterly frequency reported by these institutions to the Bank of England. The data includes lender size, liquidity and core funding from balance sheets; return on assets from income statement; and share of household credit and capital ratio from other regulatory reports.

Finally, the HM Land Registry's Price Paid Data contains information on the price paid, date, and location of all residential housing transactions in England and Wales. We use this data to understand the effect of the policy on house prices, both immediately after its introduction as well as during the correction in housing market dynamics following the Brexit referendum. Therefore, we focus on the period between 2012 (two years before the introduction of the policy) and 2018 (two years after the referendum).

3.2 Empirical Strategy

Our empirical strategy exploits the variation in the exposure to the regulation across lenders for identification. In this section, we first define the criteria for classifying lenders into two groups: more-constrained (alternately "constrained") and less-constrained (alternately "unconstrained") lenders. We study the impact of the regulation on high LTI lending by differentially constrained lenders using loan-level data in Equations 1 and 2. We use lending data aggregated at the quarterly \times local area \times borrower-income level in Equation 3 to study the effects of the regulation on lending to borrowers in the bottom income-quintile in local areas more exposed to constrained lenders. Finally, we study house price growth and mortgage performance in local areas with more constrained lenders *and* more low-income borrowers using Equation 4 and Equation 5, respectively.

Constrained and unconstrained lenders. The 2014 regulation set a 15% limit on the proportion of new residential mortgages that lenders could issue every quarter with

loan-to-income ratios (LTI) at or greater than 4.5. We use the lenders’ quarterly shares of high LTI loans in the run up to its announcement (from 2013Q3 to 2014Q2) to identify lenders more likely to be constrained by the regulation. There are 5 lenders, which issue a third of all mortgages in our sample, whose average level of high LTI loans is greater than (or very close to) 15%. We refer to these five lenders as “constrained” lenders. The remaining lenders are well within the 15% limit set by the regulation and hence less constrained by the regulation.¹⁹ We refer to these lenders as “unconstrained” lenders. Apart from the shares of high LTI lending, constrained lenders do not significantly differ from unconstrained lenders in other key characteristics, such as size, liquidity, and capital ratio, as we show in Table A.3 in the Online Appendix.

We exploit the differential exposure to the regulation for constrained and unconstrained lenders as a quasi-experiment. Figure 2 shows the time frame. We use data on lending 2-years prior to and 2-years after the announcement of the regulation in 2014Q2, i.e. all the mortgages issued during 2012Q3-2016Q2. We drop all mortgages issued during 2014Q3 from our sample since the regulation became binding only starting 2014Q4. We restrict our study to mortgages issued until 2016Q2 when the UK held the so-called Brexit referendum.

Share of high LTI mortgages. The baseline specification described in Equation 1 is a standard difference-in-differences specification which uses loan-level data to test whether mortgages issued by constrained lenders are less likely to be high LTI loans ($LTI \geq 4.5$). The treatment variable in this specification is $Post_t \times Constrained_b$ where $Post_t$ and $Constrained_b$ are dummies which indicate periods after the implementation of the regulation (2014Q4 onwards) and the constrained lenders, respectively. We regress a dummy $\mathbb{D}(LTI \geq 4.5)_{i,l,b,t}$ indicating whether a loan issued to borrower i in local area l by bank b in quarter t has $LTI \geq 4.5$ on $Post_t$ interacted with $Constrained_b$, while controlling for borrower and loan characteristics ($X_{i,l,b,t}$), bank \times time controls ($X_{b,t}$), local area \times time fixed-effects ($f_{l,t}$), and local area \times bank fixed-effects ($f_{l,b}$).²⁰ Local area l refers to the geographic location of the underlying property at the Local Administrative Unit Level 1 (LAU1) level based on the postcode reported in the loan-level PSD001 dataset.²¹

The coefficient β on $Post_t \times Constrained_b$ in this specification reflects the change in the share of high LTI loans in the portfolio of constrained banks in $Post_t$ periods.²² The sample

¹⁹We cannot share the lender *specific* shares of high LTI loans because of data confidentiality requirements.

²⁰Borrower controls include borrower type, employment status, age, and income. Loan controls include type of rate, term, type of repayment, loan value, and property value. Bank controls include size, liquidity and core funding ratios, return on assets, CET1 ratio, and share of household lending. These variables are described in Online Appendix B.

²¹The Office of National Statistics provides an overview of the administrative geographies for the United Kingdom at [link](#). LAU1 is the second lowest level of administrative geography, followed by electoral wards. There are 415 LAU1 local areas in the UK with an average population of around 158,000 residents (2016).

²²The estimated coefficient does not confound the effects of ‘affordability tests’ implemented along with

period for loan-level regressions is from 2012Q3 to 2016Q2. As mentioned before, we exclude loans issued in 2014Q3, but the results are robust to their inclusion.

$$\mathbb{D}(\text{LTI} \geq 4.5)_{i,l,b,t} = \beta \cdot \text{Post}_t \times \text{Constrained}_b + \gamma_1 \cdot X_{b,t} + \gamma_2 \cdot X_{i,l,b,t} + f_{l,t} + f_{l,b} + \varepsilon_{i,l,b,t} \quad (1)$$

Composition of high LTI mortgages. We use an analogue of the above specification to test whether the average size, LTV, and borrowers’ income of the high LTI loans change differentially for the constrained banks after the regulation. For these regressions, the dependant variables ($\log(\text{Loan Size})_{i,l,b,t}$, $\text{LTV}_{i,l,b,t}$, and $\log(\text{Borrower Income})_{i,l,b,t}$) are regressed on $\text{Post}_t \times \text{Constrained}_b$ in the sub-sample of high LTI loans while including all the controls described in Equation 1.

Redistribution of high LTI loans across borrowers. We use Equation 2 to investigate the impact of the regulation on borrower income and age in the sub-sample of mortgages with $\text{LTI} \geq 4.5$. We classify mortgages into income- and age-quintiles based on the distribution of reported income and age of borrowers of all mortgages issued during 2012Q3-2014Q2 (i.e. before the introduction of the regulation) and assign them a dummy $\mathbb{D}(\text{Inc}/\text{Age} = j)$, where $j \in \{1, 2, 3, 4, 5\}$. For each individual i , we use information on the location of the underlying property to position them in their corresponding quintile j based on the income- or age-distribution of all mortgages issued at a regional level. We use a localized distribution, rather than say the national distribution, to account for regional differences across the UK.²³

The estimated $\beta_{\text{Inc}/\text{Age},j}$ reflects the change in the share of the borrower-quintile j (in income or age) in high LTI mortgages issued by constrained lenders in Post_t periods.

$$\mathbb{D}(\text{Inc}/\text{Age}=j)_{i,l,b,t} = \beta_{\text{Inc}/\text{Age},j} \cdot \text{Post}_t \times \text{Constrained}_b + \gamma_1 \cdot X_{b,t} + \gamma_2 \cdot X_{i,l,b,t} + f_{l,t} + f_{l,b} + \varepsilon_{i,l,b,t} \quad (2)$$

To test the parallel-trends assumption underlying equations Equation 1 and Equation 2, we regress the dependant variables on Constrained_b interacted with the full set of time-dummies to test whether the estimated coefficients in periods *before* the introduction of the

the lender-level restriction on high LTI lending. As per these tests, UK lenders must assess the affordability of each mortgage issued to ensure that borrowers could still afford the mortgage if the Bank rate were to increase by 300 basis points (stress rate) within 5 years of the origination of the mortgage. However, these tests had already been introduced by the FCA in 2012 and UK lenders were using stress tests of 250-300 basis points when the 2014 regulation were introduced. See [Bank of England \(2014\), Box 5](#) for details.

²³The income and age quintile classification is based on all mortgages issued at the Local Administrative Unit 2 (LAU2) level. There are 40 LAU2 regions in the UK. LAU2 is at a higher level than LAU1, the level at which we aggregate lending data in Section 4.3. The main reason for using LAU2 instead of the more granular LAU1 level to classify borrowers into their income- or age-quintiles is the presence of very few mortgages for some local areas at the LAU1 level in Pre_t periods to generate reliable quintile thresholds.

regulation are 0.

Extensive and intensive margin. We analyse credit substitution since the regulation is binding on UK lenders and not on households: borrowers who could not receive high LTI loans from more-constrained banks (i.e. low-income borrowers) could either receive a loan with $LTI < 4.5$ from those banks or obtain a loan from a less-constrained bank. To study the net effect of these two substitution channels, we aggregate lending data at a quarterly \times local area \times borrower-income level, and use a triple difference-in differences specification (Equation 3) to check whether the regulation has an overall effect on lending to low-income borrowers at the extensive (number of mortgages) and intensive (total £ value of mortgages) margins in local areas more exposed to the constrained lenders after the regulation.

First, we calculate the share of mortgages issued by the constrained lenders a year before the introduction of the regulation—i.e. during 2013Q3-2014Q2—in a given local area ($ConstrShare_t$, in percentages). Next, we aggregate lending by volume and value at a quarterly \times local area \times income-band (5 quintiles) level. We estimate Equation 3 regressing the total number (extensive margin) or total £ value (intensive margin) of mortgages aggregated by quarter \times local area \times income-quintile on $Post_t \times ConstrShare_t \times \mathbb{D}(Inc=1)$, where $ConstrShare_t$ is as described earlier and $\mathbb{D}(Inc=1)$ is a dummy variable which indicates lending to low-income borrowers (i.e. borrowers in the bottom income-quintile).

$$\log(\mathbb{L}_{l,j,t}) = \beta_{Inc} \cdot Post_t \times ConstrShare_t \times \mathbb{D}(Inc=1) + f_{l,j} + f_{j,t} + f_{l,t} + \varepsilon_{l,j,t} \quad (3)$$

The coefficient β_{Inc} reflects the impact of the regulation on lending to borrowers in the bottom income-quintile in local areas with a higher exposure to constrained banks. We saturate the specification with local area \times time ($f_{l,t}$), local area \times income-quintile ($f_{l,j}$), and income-quintile \times time ($f_{j,t}$) fixed-effects. This allows for strong identification since we control for factors that may impact borrowers in a specific income group or specific local areas at a given time, such as a demand shock common to low-income groups or local areas in $Post_t$ periods. We report standard errors clustered at the local-area level.

To test the parallel-trends assumption underlying Equation 3, we regress the lending data on $ConstrShare_t \times \mathbb{D}(Inc=1)$ interacted with the full set of time-dummies to test whether exposure to the constrained lenders explains lending to low-income borrowers only in periods after the introduction of the regulation.

House price growth. We use the specification described in Equation 4 to analyse the effect of the policy on house prices. The dependent variable is the annual change in the logarithm of house prices in local area l at *half-year* t . Local house prices are based on the 25th, 50th (median), and 75th percentile of all housing transactions in local area l at time t . We use half-years, rather than quarters, in order to reduce noise since some local areas have

a limited number of transactions in a given quarter. We study the period from 2012H2 to 2018H1.

$$\begin{aligned} \Delta \log(P_{l,t}) = & \beta_{pol} \cdot \text{Post}_t \times \text{Constr}_l \times \text{LowInc}_l + \beta_{ref} \cdot \text{BrexitRef}_t \times \text{Constr}_l \times \text{LowInc}_l \\ & + \gamma_1 \cdot \text{Post}_t \times \text{Constr}_l + \gamma_2 \cdot \text{Post}_t \times \text{LowInc}_l + \gamma_3 \cdot \text{BrexitRef}_t \times \text{Constr}_l \quad (4) \\ & + \gamma_4 \cdot \text{BrexitRef}_t \times \text{LowInc}_l + f_l + f_{r,t} + \varepsilon_{l,t} \end{aligned}$$

We regress the dependent variable on two triple-interaction terms. The first is $\text{Post}_t \times \text{Constr}_l \times \text{LowInc}_l$, where Post_t is a dummy variable which indicates periods from 2014H2 onwards, Constr_l indicates local areas where the market share of constrained lenders before the policy (i.e. ConstrShare_l described in Equation 3) is greater than the median share, and LowInc_l indicates local areas where the share of mortgages received by low-income borrowers is above the 75th percentile.²⁴ The interaction between Constr_l and LowInc_l captures local areas more affected by the regulation, i.e. those with a higher share of constrained lenders *and* a substantial number of low-income borrowers. Figure 6 shows the spatial distribution of the affected local areas in a map of the UK; the affected local areas are not concentrated in a specific region of the UK. The coefficient β_{pol} reflects whether house price growth changes in local areas more affected by the regulation.

The second triple-interaction is $\text{BrexitRef}_t \times \text{Constr}_l \times \text{LowInc}_l$, where BrexitRef_t indicates periods after the Brexit referendum in the UK, i.e. from 2016H2 onwards. The coefficient β_{ref} reflects whether house price growth is different in local areas more affected by the policy after the Brexit referendum.

We add local area (f_l) and region×time ($f_{r,t}$) fixed-effects in the specification. Local area fixed-effects control for time-invariant heterogeneity across local areas that could affect house prices. Region×time fixed-effects control for shocks to the housing market specific to the regions in the UK—i.e., we compare local areas with different exposures to constrained lenders within a particular region. This means that our results cannot be driven by, for instance, a differential impact of the referendum on housing markets in London vs. Wales.

Mortgage performance. We also test whether there is a change in mortgage non-performance rates in affected local areas (those with ex-ante high exposure to constrained lenders and higher concentration of low-income borrowers) after the Brexit referendum.

We use the *stock* data on mortgages to create a database of the proportion of non-

²⁴Results reported in Table 7 and Table 8 are significant when we classify local areas in the top tercile or the top 2-quantile of the share of mortgages received by low-income borrowers as LowInc_l . Online Appendix Figure A.3 shows that there is sufficient cross-variation between share of mortgages issued by constrained UK lenders and the share of mortgages issued to low-income borrowers across local areas—a factor essential for identification using the specification described in Equation 4.

performing mortgages in a given local area in the UK over time. A mortgage is classified as being non-performing in case it has a payment shortfall, it is under forbearance, or if it has a possession order.²⁵ These three criteria form three disjoint sets which together form the set of distressed or non-performing loans.

Table 1 (A) shows the summary statistics of the proportion of mortgages with a payment shortfall, under forbearance, and with a possession order at a local-area level over time. UK local areas have roughly 1.8% of mortgages with a payment shortfall at any given point in time. The total share of non-performing loans is 4.8%. These summary statistics are based on such shares for six periods for each local area—2015Q2 (i.e., end of June 2015), 2015Q4 (i.e. end 2015), 2016Q2, 2016Q4, 2017Q2 and 2017Q4.²⁶ Thus we have 3 observations before the Brexit referendum result and 3 observations post-referendum for the 415 local areas in our sample.

We study the evolution of mortgage non-performance rates in the form of payment shortfalls and forbearance using a modified version of Equation 4 (described below) since, unlike the data for house price growth, we do not have data on mortgage performance at a local-area level before the introduction of the mortgage regulation.

$$NP_{l,t} = \beta_{ref}^{np} \cdot \text{BrexitRef}_t \times \text{Constr}_l \times \text{LowInc}_l + \gamma_1 \cdot \text{BrexitRef}_t \times \text{Constr}_l + \gamma_2 \cdot \text{BrexitRef}_t \times \text{LowInc}_l + f_l + f_{r,t} + \varepsilon_{l,t} \quad (5)$$

$NP_{l,t}$ is the mortgage non-performance rate (payment shortfall or forbearance) in local area l at time t . The rest of the variables are as defined in Equation 4. β_{ref}^{np} reflects the differential rates of mortgage non-performance in the affected local areas (i.e. $\text{Constr}_l \times \text{LowInc}_l$) after the Brexit referendum.

The debt-reduction channel. We finally test whether the channel through which the policy affects house price growth and mortgage performance is the reduction in mortgage lending to low-income borrowers. To do so, we calculate the change in the total lending to low-income borrowers between 2014Q2 (announcement of the policy) and 2016Q2 (Brexit referendum) for each local area, and normalise this difference by the total amount of lending in 2014Q2: $\Delta\text{Low-Income Debt}_l = (\text{Credit to borrowers in the first income quintile in 2016Q2}$

²⁵Forbearance refers to mortgages that have received modified terms from lenders after a long period of remaining under arrears. A mortgage under forbearance may have received one of 7 different types of modified terms—payment shortfall arrangement; payment shortfall capitalisation; switch to interest only; payment suspended; reduced payments; term extension; and, other forbearance. In June 2016, mortgages with a payment shortfall are, on an average, under arrears for 6 months; mortgages under forbearance have been under arrears for, on an average, 22 months.

²⁶PSD007 data—or the stock of mortgages in the UK—has been reported half-yearly since 2015. PSD007 is based on regulatory reports submitted to the conducts’ authority in the UK (the FCA) by individual lenders.

- Credit to borrowers in the first income quintile in 2014Q2) / Credit to borrowers in 2014Q2. We regress this variable on the interaction $\text{Constr}_l \times \text{LowInc}_l$, as shown in Equation 6.

$$\Delta \text{Low-Income Debt}_l = \alpha + \beta_{debt} \cdot \text{Constr}_l \times \text{LowInc}_l + \varepsilon_l \quad (6)$$

We take the estimated model—which has a *negative* and significant β_{debt} , explicitly showing that areas more affected by the policy have a stronger reduction in credit to low-income borrowers as suggested by previous specifications—and create two new variables: the *predicted* and the *residual* $\Delta \text{Low-Income Debt}$. In doing so, we bisect the variation in credit to low-income borrowers into two components: the change associated to the policy, and the endogenous change orthogonal to the policy. We use these two variables to explain house price growth after the Brexit referendum by using Equation 7 (we use an equivalent approach for mortgage performance). We multiply these two variables by -1 to make the sign of the coefficients more readily comparable to the previous two specifications. β_{ref}^{hpred} reflects the effect of the debt-reduction to low-income borrowers due to the policy on house price growth after the Brexit referendum.

$$\begin{aligned} \Delta \log(P_{l,t}) = & \beta_{ref}^{hpred} \cdot \text{BrexitRef}_t \times \text{Predicted} (-\Delta \text{Low-Income Debt}_l) \\ & + \beta_{ref}^{hres} \cdot \text{BrexitRef}_t \times \text{Residual} (-\Delta \text{Low-Income Debt}_l) + f_l + f_{r,t} + \varepsilon_{l,t} \end{aligned} \quad (7)$$

4 Empirical Results

4.1 Share of high LTI mortgages

Table 2 presents estimates of the coefficient β described in Equation 1 using loan-level data. The negative and significant estimate of the coefficient β suggests that constrained lenders reduce mortgages with high loan-to-income ratios (LTI) in the aftermath of the regulation. The coefficient is highly robust to the inclusion of different set of controls. In Column 1, we use the full sample with no controls. In Columns 2-6, we restrict the sample to those observations for which we have all controls in order to not have different samples with the inclusion of controls.²⁷ The coefficient is highly significant and stable as we add borrower and loan-controls (Column 3), bank \times time controls (Column 4), local area \times time FE (Column 5), and local area \times bank FE (Column 6). In Column 6, which includes all the controls mentioned earlier, the coefficient indicates that the share of high LTI mortgages in the portfolio of constrained lenders drops by roughly 3.9% in the aftermath of the regulation.

²⁷The sample size decreases from 2.9 million to 1.8 million when all controls are included. This drop is primarily due to missing observations for loan-level controls described in Online Appendix A.1.

We include Post_t and Constrained_b as standalone variables when they are not absorbed by fixed-effects; estimated coefficients for either variable is positive and smaller in magnitude than the coefficient estimated on $\text{Post}_t \times \text{Constrained}_b$.

Figure 3 plots the coefficients obtained by regressing the dependant variable on the dummy for constrained lenders interacted with the full set of time-dummies (except 2014Q2, the reference time period) and the controls described in Equation 1. The resulting coefficients are negative and significant *only* in periods after the introduction of the regulation. The coefficients in Pre_t periods are quantitatively and statistically not different from 0, consistent with the parallel-trends assumption underlying Equation 1.

4.2 Composition of high LTI mortgages

Next, we investigate whether constrained lenders change the composition of their high LTI mortgages in response to the regulation. First, we test whether there is a change in the size and loan-to-value ratio (LTV), aspects not restricted by the 2014 regulation, of high LTI mortgages. Then, we check whether there is a change in the income or age composition of borrowers receiving high LTI mortgages. This latter point shows whether the contraction in high LTI lending by constrained lenders might disproportionately affect some borrower cross-sections.

Size, LTV, and income of high LTI mortgages. We use the regression specification described in Equation 1 with dependent variables $\log(\text{Loan size})$, LTV , and $\log(\text{Borrower Income})$ in a sub-sample of mortgages with $\text{LTI} \geq 4.5$ (around 150 thousand mortgages). The estimate β shows changes in these alternate dependant variables in the portfolio of high LTI mortgages issued by the constrained lenders after the regulation.

As we see in Table 3 (A), average loan size of high LTI mortgages issued by constrained lenders is roughly 9% larger in periods after the introduction of the regulation. In a similar vein, constrained lenders issue high LTI mortgages with 1.8 percent point higher LTV ratio (Table 3 (B), Column 6) in Post_t periods. Thus, while constrained lenders issue fewer high LTI mortgages after the introduction of the regulation, those loans are larger in size and have a higher LTV. We further elaborate on the effect of the regulation on the overall changes in the LTI distribution of mortgages issued by constrained lenders in Section 4.5.

Table 3 (C) further shows that average borrower incomes are 11.5% higher in high LTI mortgages issued by constrained lenders in Post_t periods. These results lead us to investigate whether the contraction in high LTI lending by constrained lenders has a disproportionate impact on any specific borrower cross-sections such as low-income and/or younger borrowers.

The results presented so far use a specification which distinguishes more-constrained and

less-constrained lenders using an indicator variable. Results are robust to directly interacting the lender-specific share of high LTI loans in the total mortgage portfolio with Post_t as an alternate, continuous explanatory variable (see Online Appendix E).

Redistribution of high LTI loans across borrowers. We regress dummies indicating low-income or young borrowers (as described in Section 3) on the set of controls described in Equation 2 in the sub-sample of loans with $\text{LTI} \geq 4.5$. The coefficient β_j (on $\text{Post}_t \times \text{Constrained}_b$, where j refers to low-income or young borrowers) reflects the change in the share of the borrower-type j in the sub-sample of high LTI loans issued by constrained lenders in Post_t periods.

Table 4 (A) highlights a stark change in the distribution of borrower-incomes in high LTI loans issued by constrained lenders in Post_t periods. Column 1 shows that constrained lenders decrease the share of high LTI loans issued to borrowers in the lowest income-quintile by 5.2%. This decline is matched by an increase in the proportion of high LTI loans obtained by borrowers in the highest income-quintile (Column 5). Interestingly, Table 4 (B) shows that there was no change in the age composition of high LTI mortgages issued by constrained lenders in Post_t periods. This shows that the contraction in high LTI loans by constrained lenders had a stronger effect on low-income borrowers. They are affected by both the reduction in the share of high LTI loans (which they are more likely to borrow) and the drop in their share *within* high LTI loans issued by constrained lenders after the regulation. This result is also consistent with the increase in the average size of high LTI loans issued by constrained lenders reported in Table 3 (A).

However, low-income borrowers receiving fewer high LTI loans from constrained lenders may not reflect an aggregate effect. For instance, such borrowers may receive smaller loans (with $\text{LTI} < 4.5$) from the constrained lenders or obtain mortgages with high LTI from unconstrained lenders. We find some evidence of the former channel in the results reported in Table 5 (A). We regress the dummy indicating low-income borrowers ($\mathbb{D}(\text{Inc}=1)$) on $\text{Post}_t \times \text{Constrained}_b$ in different sub-samples of LTI groups. While the share of lowest income-quintile borrowers drops in the sample of high LTI loans issued by constrained lenders after the regulation (Table 5 (A), Column 5),²⁸ it expands in the rest of the LTI groups (Columns 1-4). In particular, the expansion is stronger in the sample of mortgages with LTI just below the 4.5 threshold (Table 5 (A), Column 4).²⁹

²⁸This is the same regression as in Table 4 (A), Column 1.

²⁹The coefficients in Table 5 (A) do not suggest that the increase in the share of loans with $\text{LTI} < 4.5$ going to low-income borrowers perfectly substitute for the drop in high LTI lending by constrained lenders. The reduction in high LTI lending to low-income borrowers by constrained lenders is driven by both the reduction in the share of such borrowers in their high LTI portfolio, *and* the drop in the size of this portfolio itself. We address the net effect on lending to low-income borrowers in Section 4.3.

Table 5 (B) further shows that unconstrained lenders also substitute for the contraction in high LTI lending to low-income borrowers by constrained lenders. We regress the dummy indicating low-income borrowers ($\mathbb{D}(\text{Inc}=1)$) on $\text{Post}_t \times \text{Unconstrained}_b \times \text{ConstrShare}_l$ in different sub-samples of LTI groups. Unconstrained_b is a dummy for the unconstrained lenders and ConstrShare_l is the market share of *constrained* lenders in the mortgage market in local area l . Column (5) in Table 5 (B) reports a positive and statistically significant coefficient on $\text{Post}_t \times \text{Unconstrained}_b \times \text{ConstrShare}_l$ in the sub-sample of loans with $\text{LTI} \geq 4.5$. This shows that there is an increase in the proportion of loans obtained by low-income borrowers in high LTI loans issued by unconstrained lenders in Post_t periods in areas more exposed to *constrained* lenders.

4.3 Lending to low-income borrowers

What is the net effect on lending to low-income borrowers of the contraction in high LTI lending by constrained lenders and any possible substitution for this contraction by the two channels described above? To address this question, we aggregate lending by both constrained and unconstrained lenders at the local area \times borrower income-quintile \times quarterly level. We use this aggregate data and the saturated regression specification described in Equation 3 to test whether low-income borrowers in local areas more exposed to the constrained lenders experience any change in lending at the extensive (number of loans) or the intensive (value of loans) margin after the regulation.

The coefficient β_{Inc} estimated for the total volume of lending in Table 6 (A) is negative and highly significant.³⁰ This suggests that there is a contraction in the volume of loans to borrowers in the lowest income-quintile in local areas more exposed to constrained banks in Post_t periods. Similarly, the estimated coefficient for the total value (or £ amount) of lending is -0.0173 (Table 6 (B)) and highly significant. The estimated coefficients are highly robust; a lower-bound for these coefficients calculated accounting for unobserved local-area heterogeneity—and using the technique described in Altonji et al. (2005) and Oster (2017)—is remarkably close to the actual coefficients. The estimated coefficients show that borrowers in the bottom income-quintile in local areas with a one standard deviation higher pre-shock share of constrained lenders experience a 9.0% drop in the number of mortgages and a 11.4% drop in the £ amount of mortgages in Post_t periods.

Figure 5 shows coefficients from regressing the aggregated lending data on $\text{ConstrShare}_l \times \mathbb{D}(\text{Inc}=1)$ interacted with the full set of time-dummies. The estimated coefficients are zero in periods *before* 2014Q3 and show a contraction in the amount of lending to low-income

³⁰The lending data is aggregated at the level of 3-digit administrative units (LAU1). See Section 3.2, particularly *Extensive and intensive margin*, for details.

borrowers only in periods *after* 2014 Q3. These time-varying coefficients are consistent with the parallel-trends assumption underlying Equation 3.

4.4 House prices and mortgage defaults in affected local areas

We explore whether the constraint on high LTI mortgage lending has an effect on house prices and mortgage performance. Macroprudential policy is a tool to “lean against the wind” in order to smooth the credit cycle. We investigate this hypothesis by studying whether the policy leads to lower house price growth during a boom and, consequently, lower house price drops during a correction. The Brexit referendum in June 2016 provides the correction in the form of house price cooling across the UK following the vote. We use this shock to test whether the policy had an effect on house prices and mortgage performance in local areas more affected by the regulation, potentially alleviating the effects of the correction in house prices following the referendum.

4.4.1 House price growth

The coefficients of interest— β_{pol} and β_{ref} in Equation 4—correspond to the triple interactions between $Constr_t$, $LowInc_t$, and the dummies for the post-policy ($Post_t$) and post-referendum ($BrexitRef_t$) periods. These interactions explore whether house prices of local areas with a high presence of constrained banks and low-income borrowers behave differently after the introduction of the policy and the Brexit referendum. To shut down other potential channels, we saturate the specification with local authority (f_l) and region \times time ($f_{r,t}$) fixed effects.³¹

Figure 1 shows the actual aggregate trends for annual (median) house price growth for England and Wales during 2012Q3-2018Q1, the period relevant to our study. There is an increasing trend since 2013Q3, peaking at around 7.5% in the quarter before the EU referendum. After the referendum, however, the trend completely reverses and, in just one quarter, growth falls by 2 percentage points and continues to fall afterwards.

Table 7 reports results based on Equation 4. In columns 1-3, the dependent variable is the annual change in the 25th percentile of house price transactions at a local-area level. In Column 1, we use only time fixed-effects. The coefficient β_{pol} is negative, as well as statistically and economically significant. After the policy is introduced, local areas with high shares of constrained banks and low-income borrowers (“affected local areas”) experience a 1.5 percentage points lower house price growth. The average price growth during this period is 6.8%, so the effect is sizable. This result is consistent with our previous evidence that the

³¹These regions are: East, East Midlands, London, North East, North West, South East, South West, Wales, West Midlands, and Yorkshire and the Humber.

limit to high-LTI mortgages leads to a decline in mortgage supply, particularly to low-income borrowers. This supply shock particularly affects the lower-end of housing transaction prices.

The coefficient β_{ref} , on the other hand, is positive, and statistically and economically significant. Post-referendum and in the middle of a correction in house price growth, affected local areas experience a higher growth in house prices. The magnitude of the effect— $\beta_{pol} + \beta_{ref}$ —is 1.34% and statistically significant (p-value-6.9%). This result is consistent with the objective of such macroprudential interventions: restricting leverage during a boom leads to lower drops in house price growth following a correction.

The coefficients of the triple interactions are fairly stable as we add local area and region \times time fixed-effects in the second and third columns. In the most saturated specification (Column 3), the coefficients imply a 1.35 percentage points lower growth post-policy, and a 1.26 percentage points higher growth after the referendum.

We test the timing of the changes by plotting the coefficient of the interaction between $Constr_t$ and $LowInc_t$ over time. To do so, we use the specification in Column 3 but interact the two variables with time (half-year) dummies. The coefficients are plotted in Figure 7. Before the introduction of the policy (first vertical dotted line) the coefficients are centred around 0, which alleviates pre-trend concerns. After the policy is introduced, the coefficient becomes negative until the referendum (second vertical dotted line). After the referendum, the coefficients are positive.

We compare these results to those based on the middle- and upper-end of the distribution of house price transactions. We expect estimates of β_{ref} based on these alternate house prices to be weaker—consistent with the effect of the policy on lending to low-income borrowers and hence the lower-end of house prices. This analysis could also rule out the possibility that the shock to house price growth after the referendum is correlated with other shocks in affected local areas: if that were the case, estimates of β_{ref} would be similar irrespective of house prices being drawn from the lower, mid or upper-end of housing transactions.

Columns 4-6 of Table 7 show the results when using the change in the median price as our dependent variable. The magnitude of the coefficients is significantly lower, and in particular the effect of the policy is insignificant. In Columns 7-9 of Table 7, we use the 75th percentile of the distribution of house prices to calculate the house price growth in each local area. The coefficient for the triple interaction during the period after the referendum is positive, although substantially smaller in size and not statistically significant.

4.4.2 Mortgage performance

Columns 1-3 in Table 8 show that the coefficient on $Ref_t \times Constr_t \times LowInc_t$ using the data on proportion of mortgages with a payment shortfall at a local-area level is negative

and significant. The estimated coefficient suggests that local areas with a higher share of constrained lenders *and* a higher share of low-income borrowers experienced a 0.064% lower rate of payment shortfall in the aftermath of the Brexit referendum. This is approximately a 4% decline from the average payment shortfall rate across UK local areas.

These results show that the affected local areas (i.e. local areas with greater presence of both constrained lenders and low-income borrowers, which experience a contraction in lending as a result of the regulation) not only experience a relatively lower drop in house price growth post-Brexit referendum, but also relatively lower rates of mortgage non-performance rates. Figure 8 (A) shows that the estimated coefficient on the dummy for affected local areas ($\text{Constr}_l \times \text{LowInc}_l$) interacted with half-yearly time-dummies is significant *only* in periods after the Brexit-referendum, i.e. after 2016Q2.

However, we observe this lower mortgage non-performance rate primarily in terms of payment shortfall; the result showing changes in rates of mortgages under forbearance in affected local areas is less robust (Columns 4-6 in Table 8). Figure 8 (B) shows that the proportion of mortgages under forbearance in affected local areas in periods after the Brexit-referendum are not statistically different from the proportion before 2016Q2. As noted earlier, mortgages undergo forbearance only after remaining under arrears for a long duration. It is possible that our sample on mortgage non-performance (until 2018Q2) is not long enough to pick up changes in forbearance rates in affected local areas after the Brexit referendum.³²

4.4.3 Pre-correction debt-reduction channel and post-correction outcomes

In Table 9, we show that the channel through which the policy has an effect on house prices and mortgage performance is the debt of low-income borrowers. In Panel A, we show that a stronger reduction in credit to low-income borrowers due to the policy leads to higher house price growth during the correction period: the coefficient is positive and statistically significant. Economically, the coefficient shows that an additional 1 percentage point decline in credit to low-income borrowers (normalised by total credit in the local area) is associated with an additional 1 percentage point of growth in house prices. The coefficient is slightly smaller in absolute value when using the median house price at a local-area level (columns 3 and 4), and falls to around a third when using the 75th percentile of house prices. These differences are consistent with the policy affecting debt to low-income borrowers. Differently, the coefficient on the residual from Equation 6 is both statistically and economically insignificant; i.e., the endogenous changes in household debt unrelated to the policy do not affect house growth post-referendum.

³²This also applies to mortgages with a possession order; default rates using possession orders do not undergo a relative change in affected local areas after the Brexit referendum.

In Panel B of Table 9, we show that the same channel also underlies the reduction in mortgage non-performance rates post-referendum. A higher reduction in credit to low-income borrowers driven by the policy reduces payment shortfalls and forbearance rates. Economically, a 1 percentage point reduction in low-income credit is associated with a 4bp drop in the share of mortgages with payment shortfall and a 6bp drop in the share of mortgages under forbearance. Again, contrary to the change in debt due to the policy, the coefficient of the residual change in debt is not statistically or economically significant.

4.5 Overall changes in the LTI distribution of lenders

We now discuss the effect of the regulation on the overall LTI distribution of mortgages issued by constrained and unconstrained lenders. Figure 9 (A) shows the LTI distribution of the mortgages in periods before (2012Q3 - 2014Q2) and after (2014Q4 - 2016Q2) the introduction of the regulation in colours blue and red for constrained (left panel) and unconstrained banks, respectively. The LTI distribution of mortgages issued by the constrained banks shows a clear shift to the right in $Post_t$ periods, with an increase in the mass of mortgages with LTI close to and below 4.5. We use a counterfactual distribution estimation technique to provide a breakdown of the changes in the LTI distribution between Pre_t and $Post_t$ periods attributable to the regulation and to changes in the underlying distribution of covariates plausibly unrelated to the regulation.

We use the technique described in Chernozhukov et al. (2013) to estimate a counterfactual LTI distribution for $Post_t$ periods for the two lender groups. As per the technique, we take the distribution of covariates (bank \times time, borrower, and loan controls described in Online Appendix Table A.1) in periods after the introduction of the regulation ($Post_t$ periods) and estimate a counterfactual LTI distribution based on the likelihood of obtaining a LTI for a given combination of covariates in periods before the introduction of the regulation (Pre_t periods). We use the estimated counterfactual to provide a breakdown of the changes in quantiles of the LTI distribution between Pre_t and $Post_t$ periods attributable to changes in covariates and to the regulation, the so-called Oaxaca-Blinder decomposition (Figure 9 (B)).

Figures 9 (B) (i) and (ii) show the Oaxaca-Blinder decomposition for the constrained and unconstrained banks, respectively. The decomposition is based on a breakdown of the total change (solid line) in the quantiles of the LTI distribution into two driving forces—changes in underlying covariates (dotted line) and the effect of the regulation (dashed line).³³ Figure 9 (B) (i) shows that changes in LTI quantiles for the constrained banks in $Post_t$ periods

³³A simple way to understand the decomposition is that the effect of changes in covariates is equal to the differences between the estimated counterfactual and the Pre_t period LTI distribution; differences between the $Post_t$ period and the estimated counterfactual LTI distribution show the effect of the regulation.

are largely attributable to the regulation; the changes driven by differences in underlying covariates is quantitatively a lot smaller, and consistently so across the different quantiles. On the other hand, changes in the quantiles of the LTI distribution for unconstrained banks in Figure 9 (B) (ii) are largely explained by the changes in underlying covariates. The effect of the regulation on the LTI quantiles of unconstrained banks is estimated to be close to 0.

5 Conclusion

We analyse the distributional effects of macroprudential policy on mortgage and house price cycles, thereby showing costs and benefits of macroprudential policy. To the best of our knowledge, this is the first paper which studies the effects of changes in macroprudential policy on household leverage during both boom and correction periods. For empirical identification, we exploit the universe of UK mortgages in conjunction with a 15% cap on the proportion of new mortgages with high loan-to-income ratios ($LTI \geq 4.5$) imposed in 2014 on UK mortgage lenders. These lenders are differentially affected based on their pre-policy share of high LTI loans in their mortgage portfolio.

Our robust results show that more-constrained lenders issue fewer high-LTI mortgages. There is also regulatory arbitrage post-regulation as more-constrained lenders increase the loan-to-value ratios (LTV) and average loan size of high LTI mortgages, and disproportionately raise the mass of mortgages with LTI close to and below the 4.5 threshold. The regulation has a stronger effect on lending to low-income households who are affected by both the reduction in high LTI loans (which they are more likely to borrow) and the drop in their share *within* high LTI loans issued by constrained lenders after the regulation.

Even though the regulation is on lenders (and not households), substitution by less-constrained lenders is partial and leads to overall credit contraction to low-income households in areas more exposed to the more-constrained lenders. This causes dampening of house price growth in local areas with a higher share of these constrained lenders *and* a higher share of low-income borrowers in periods after the introduction of the regulation, but before the Brexit referendum. Following the Brexit referendum, that led to a house price correction across the UK, the 2014 macroprudential policy leads to significantly lower house price drops and relatively lower mortgage default rates during the correction. The better house price and default rate effects are fully driven by the reduction in household debt due to the policy, not by endogenous changes in debt orthogonal to the policy.

References

- Acharya, Viral V, Katharina Bergant, Matteo Crosignani, Tim Eisert, and Fergal McCann (2019). “The Anatomy of the Transmission of Macroprudential Policies: Evidence from Ireland”. *Working Paper*.
- Agarwal, Sumit, Efraim Benmelech, Nittai Bergman, and Amit Seru (2012). “Did the Community Reinvestment Act (CRA) lead to risky lending?” *Working Paper*.
- Aladangady, Aditya (2017). “Housing wealth and consumption: Evidence from geographically-linked microdata”. *American Economic Review* 107.11.
- Altonji, Joseph G, Todd E Elder, and Christopher R Taber (2005). “Selection on observed and unobserved variables: Assessing the effectiveness of Catholic schools”. *Journal of Political Economy* 113.1, pp. 151–184.
- Andersen, Asger Lau, Charlotte Duus, and Thais Lærkholm Jensen (2016). “Household debt and spending during the financial crisis: Evidence from Danish micro data”. *European Economic Review* 89, pp. 96–115.
- Auclert, Adrien (2019). “Monetary policy and the redistribution channel”. *American Economic Review* 109.6.
- Autor, David H and David Dorn (2013). “The growth of low-skill service jobs and the polarization of the US labor market”. *American Economic Review* 103.5.
- Autor, David H, David Dorn, and Gordon H Hanson (2013). “The China syndrome: Local labor market effects of import competition in the United States”. *American Economic Review* 103.6.
- Baker, Scott R. (2018). “Debt and the Response to Household Income Shocks: Validation and Application of Linked Financial Account Data”. *Journal of Political Economy* 126.4.
- Bank of England (2014). “Financial Stability Report (June)”.
- (2017). “Financial Stability Report (June)”.
- (2017b). “Financial Stability Report (November)”.
- Bekkum, Sjoerd van, Marc Gabarro, Rustom M Irani, and José-Luis Peydró (2019). “Macroprudential policy and household leverage: Evidence from administrative household-level data”. *Working Paper*.
- Berg, Tobias, Anthony Saunders, Larissa Schafer, and Sascha Steffen (2019). “Brexit and the contraction of syndicated lending”. *Working Paper*.
- Bhutta, Neil and Benjamin J Keys (2016). “Interest rates and equity extraction during the housing boom”. *American Economic Review* 106.7.
- Bhutta, Neil, Jane Dokko, and Hui Shan (2017). “Consumer ruthlessness and mortgage default during the 2007 to 2009 housing bust”. *Journal of Finance* 72.6.

- Bloom, Nicholas, Philip Bunn, Scarlet Chen, Paul Mizen, Pawel Smietanka, and Gregory Thwaites (2019). “The impact of Brexit on UK firms”. *Working Paper*.
- Born, B, GJ Muller, M Schularick, and P Sedlacek (2019). “The costs of economic nationalism: Evidence from the Brexit experiment”. *Economic Journal* 129.623.
- Broadbent, Ben, Federico Di Pace, Thomas Drechsel, Richard Harrison, and Silvana Tenreyro (2019). “The Brexit vote, productivity growth and macroeconomic adjustments in the United Kingdom”. *Working Paper*.
- Bunn, Philip and May Rostom (2014). “Household debt and spending”. *Working Paper*.
- Campbell, John Y and João F Cocco (2007). “How do house prices affect consumption? Evidence from micro data”. *Journal of Monetary Economics* 54.3.
- Chaney, Thomas, David Sraer, and David Thesmar (2012). “The collateral channel: How real estate shocks affect corporate investment”. *American Economic Review* 102.6.
- Chernozhukov, Victor, Iván Fernandez-Val, and Blaise Melly (2013). “Inference on counterfactual distributions”. *Econometrica* 81.6.
- Corbae, Dean and Erwan Quintin (2015). “Leverage and the foreclosure crisis”. *Journal of Political Economy* 123.1.
- Crowe, Christopher, Giovanni Dell’Ariccia, Deniz Igan, and Pau Rabanal (2013). “How to deal with real estate booms: Lessons from country experiences”. *Journal of Financial Stability* 9.3.
- Davies, Ronald B. and Zuzanna Studnicka (2018). “The heterogeneous impact of Brexit: Early indications from the FTSE”. *European Economic Review* 110.C.
- Davila, Eduardo and Anton Korinek (2017). “Pecuniary externalities in economies with financial frictions (vol 1, pg 1, 2017)”. *Review of Economic Studies* 84.4.
- DeFusco, Anthony, Stephanie Johnson, and John Mondragon (forthcoming). “Regulating household leverage”. *Review of Economic Studies*.
- DeFusco, Anthony A (2018). “Homeowner borrowing and housing collateral: New evidence from expiring price controls”. *Journal of Finance* 73.2.
- Di Maggio, Marco and Amir Kermani (2017). “Credit-induced boom and bust”. *Review of Financial Studies* 30.11.
- Faccini, Renato and Edoardo Palombo (2019). “News uncertainty in Brexit UK”. *Working Paper*.
- Farhi, Emmanuel and Iván Werning (2016). “A theory of macroprudential policies in the presence of nominal rigidities”. *Econometrica* 84.5.
- Favara, Giovanni and Jean Imbs (2015). “Credit supply and the price of housing”. *American Economic Review* 105.3.

- Favilukis, Jack, Sydney C Ludvigson, and Stijn Van Nieuwerburgh (2017). “The macroeconomic effects of housing wealth, housing finance, and limited risk sharing in general equilibrium”. *Journal of Political Economy* 125.1.
- Fuster, Andreas and Paul S Willen (2017). “Payment size, negative equity, and mortgage default”. *American Economic Journal: Economic Policy* 9.4.
- Ghent, Andra C and Marianna Kudlyak (2011). “Recourse and residential mortgage default: Evidence from US states”. *Review of Financial Studies* 24.9.
- Greenwald, Daniel (2017). “The mortgage credit channel of macroeconomic transmission”.
- Guiso, Luigi, Paola Sapienza, and Luigi Zingales (2013). “The determinants of attitudes toward strategic default on mortgages”. *Journal of Finance* 68.4.
- Guren, Adam M, Alisdair McKay, Emi Nakamura, and Jón Steinsson (2018). “Housing wealth effects: The long view”. *Working Paper*.
- Igan, Deniz O and Heedon Kang (2011). “Do loan-to-value and debt-to-income limits work? Evidence from Korea”. *Working Paper*.
- Jiménez, Gabriel, Steven Ongena, José-Luis Peydró, and Jesús Saurina (2017). “Macroprudential policy, countercyclical bank capital buffers, and credit supply: Evidence from the Spanish dynamic provisioning experiments”. *Journal of Political Economy* 125.6.
- Jordà, Òscar, Moritz Schularick, and Alan M Taylor (2015a). “Betting the house”. *Journal of International Economics* 96.
- Korinek, Anton and Alp Simsek (2016). “Liquidity trap and excessive leverage”. *American Economic Review* 106.3.
- Lorenzoni, Guido (2008). “Inefficient credit booms”. *Review of Economic Studies* 75.3.
- Lustig, Hanno and Stijn Van Nieuwerburgh (2010). “How much does household collateral constrain regional risk sharing?” *Review of Economic Dynamics* 13.2.
- Mayer, Christopher, Karen Pence, and Shane M. Sherlund (2009). “The Rise in Mortgage Defaults”. *Journal of Economic Perspectives* 23.1.
- Mian, Atif and Amir Sufi (2009). “The consequences of mortgage credit expansion: Evidence from the US mortgage default crisis”. *Quarterly Journal of Economics* 124.4.
- (2011). “House prices, home equity-based borrowing, and the US household leverage crisis”. *American Economic Review* 101.5.
- (2014). *House of debt: How they (and you) caused the Great Recession, and how we can prevent it from happening again*. University of Chicago Press.
- Mian, Atif, Kamallesh Rao, and Amir Sufi (2013). “Household balance sheets, consumption, and the economic slump”. *Quarterly Journal of Economics* 128.4.
- Mian, Atif, Amir Sufi, and Emil Verner (2017). “Household debt and business cycles worldwide”. *Quarterly Journal of Economics* 132.4.

- Oster, Emily (2017). “Unobservable selection and coefficient stability: Theory and evidence”. *Journal of Business & Economic Statistics*, pp. 1–18.
- Piketty, Thomas and Emmanuel Saez (2003). “Income inequality in the United States, 1913–1998”. *Quarterly Journal of economics* 118.1.
- Piskorski, Tomasz and Amit Seru (2018). “Debt relief and slow recovery: A decade after Lehman”. *Working Paper*.
- Rajan, Raghuram G (2011). *Fault lines: How hidden fractures still threaten the world economy*. Princeton University press.
- Schularick, Moritz and Alan M Taylor (2012). “Credit booms gone bust: Monetary policy, leverage cycles, and financial crises, 1870-2008”. *American Economic Review* 102.2.

Table 1 (A): Summary statistics of loan level data

| Total sample of mortgages | | | | | | |
|------------------------------|-----------|---------|---------|---------|---------|---------|
| Variable | count | mean | sd | p25 | p50 | p75 |
| Loan Value | 1,849,952 | 165,138 | 119,625 | 90,000 | 136,093 | 202,995 |
| Property Value | 1,849,952 | 270,612 | 215,928 | 145,000 | 210,000 | 318,000 |
| Mortgage Term | 1,849,952 | 23.44 | 7.75 | 18.00 | 25.00 | 30.00 |
| LTI | 1,849,952 | 3.00 | 1.10 | 2.23 | 3.05 | 3.82 |
| LTV | 1,849,952 | 66.0 | 22.3 | 51.9 | 73.1 | 84.9 |
| Borrower Income | 1,849,952 | 58,535 | 45,396 | 33,000 | 46,706 | 68,001 |
| Borrower Age | 1,849,952 | 38.21 | 9.80 | 30.00 | 37.00 | 45.00 |
| Sample of high LTI mortgages | | | | | | |
| Variable | count | mean | sd | p25 | p50 | p75 |
| Loan Value | 149,709 | 235,786 | 140,768 | 142,400 | 200,000 | 285,000 |
| Property Value | 149,709 | 362,452 | 259,608 | 202,000 | 287,000 | 427,000 |
| Mortgage Term | 149,709 | 28.16 | 6.16 | 25.00 | 30.00 | 34.00 |
| LTI | 149,709 | 4.85 | 0.30 | 4.63 | 4.77 | 4.99 |
| LTV | 149,709 | 69.8 | 15.5 | 60.3 | 73.5 | 80.6 |
| Borrower Income | 149,709 | 48,803 | 31,646 | 29,308 | 41,014 | 58,765 |
| Borrower Age | 149,709 | 35.00 | 8.24 | 29.00 | 34.00 | 40.00 |

The above shows summary statistics based on loan- and borrower-level data of residential mortgages issued in the UK during 2012Q3-2016Q2 that were under the purview of the 2014 regulation.

Table 1 (B): Summary statistics of local-area level data

| House price growth | | | | | | |
|---------------------------------|-------|-------|-------|-------|-------|-------|
| Variable | count | Mean | SD | p25 | p50 | p75 |
| $\Delta \log(\text{p25 price})$ | 4,176 | 0.051 | 0.047 | 0.021 | 0.048 | 0.079 |
| $\Delta \log(\text{p50 price})$ | 4,176 | 0.051 | 0.046 | 0.021 | 0.048 | 0.079 |
| $\Delta \log(\text{p75 price})$ | 4,176 | 0.050 | 0.050 | 0.017 | 0.047 | 0.082 |
| Mortgage non-performance rates | | | | | | |
| Variable: Mortgages... | count | Mean | SD | p25 | p50 | p75 |
| with payment shortfall (%) | 2,361 | 1.84 | 0.74 | 1.30 | 1.71 | 2.28 |
| under forbearance (%) | 2,361 | 2.69 | 0.94 | 1.98 | 2.55 | 3.26 |
| with possession order (%) | 2,361 | 0.45 | 0.28 | 0.23 | 0.39 | 0.64 |

The above shows summary statistics based on data available for house price growth and mortgage distress rates for local areas in the UK (LAU1 (lower local administrative units), or equivalently boroughs or counties). House price growth is calculated as the annual change in the logarithm of the corresponding price level (25th, 50th, or 75th percentile) paid in local area l .

Table 2: Loan level: Share of loans with high loan-to-income ratios ($LTI \geq 4.5$)

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
|-------------------------------|-----------------------|---|------------------------|------------------------|------------------------|------------------------|
| | | ← Dummy for high LTI loans ($\mathbb{D}(LTI \geq 4.5)$) → | | | | |
| $Post_t \times Constrained_b$ | -0.0295** (0.0118) | -0.0474*** (0.0116) | -0.0394*** (0.0099) | -0.0513*** (0.0093) | -0.0566*** (0.0089) | -0.0388*** (0.0076) |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank×time Controls | | | | Yes | Yes | Yes |
| Local Area×Time FE | | | | | Yes | Yes |
| Local Area×Bank FE | | | | | | Yes |
| Observations | 2,907,529 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 |
| R ² | 0.005 | 0.012 | 0.014 | 0.066 | 0.103 | 0.121 |

The dependent variable in this table is $\mathbb{D}(LTI \geq 4.5)_{i,l,b,t}$, a binary variable that equals 1 if the mortgage granted by lender b to borrower i in quarter t in local area l has a loan-to-income ratio (LTI) equal or above 4.5, and 0 otherwise. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. Column (1) includes all observations; Columns (2)-(6) include only observations for which the full set of control variables contain non-missing values. $Post_t$ is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. $Constrained_b$ is a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2), 0 otherwise. All regressions are estimated using ordinary least squares. Borrower, loan, and bank controls, as well as fixed-effects are either included (“Yes”) or not included. Online Appendix Section B shows the definition of the main controls. $Post_t$ and $Constrained_b$ are included as standalone variables when they are not absorbed by fixed-effects; the coefficients are reported in Table A.5. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 3: Loan level: Size, LTV and borrower income of high LTI loans

| (A) Size | | | | | | |
|---|---|----------|----------|----------|-----------|-----------|
| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
| | $\leftarrow \log(\text{Loan Size}) \rightarrow$ | | | | | |
| Post _{<i>t</i>} ×Constrained _{<i>b</i>} | 0.116* | 0.208*** | 0.230*** | 0.217*** | 0.0929*** | 0.0898*** |
| | (0.0624) | (0.0562) | (0.0412) | (0.0368) | (0.0214) | (0.0198) |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank×Time Controls | | | | Yes | Yes | Yes |
| Local Area×Time FE | | | | | Yes | Yes |
| Local Area×Bank FE | | | | | | Yes |
| Observations | 252,547 | 149,709 | 149,709 | 149,709 | 149,709 | 149,709 |
| R ² | 0.077 | 0.059 | 0.088 | 0.315 | 0.603 | 0.627 |
| (B) LTV | | | | | | |
| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
| | $\leftarrow \text{LTV} \rightarrow$ | | | | | |
| Post _{<i>t</i>} ×Constrained _{<i>b</i>} | 0.0552 | 0.667 | -0.892 | 0.788 | 2.408*** | 1.796*** |
| | (1.102) | (1.371) | (0.953) | (0.746) | (0.661) | (0.567) |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank×Time Controls | | | | Yes | Yes | Yes |
| Local Area×Time FE | | | | | Yes | Yes |
| Local Area×Bank FE | | | | | | Yes |
| Observations | 252,542 | 149,709 | 149,709 | 149,709 | 149,709 | 149,709 |
| R ² | 0.021 | 0.021 | 0.037 | 0.277 | 0.349 | 0.376 |
| (C) Borrower income | | | | | | |
| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
| | $\leftarrow \log(\text{Borrower Income}) \rightarrow$ | | | | | |
| Post _{<i>t</i>} ×Constrained _{<i>b</i>} | 0.115* | 0.223*** | 0.250*** | 0.241*** | 0.116*** | 0.115*** |
| | (0.0687) | (0.0601) | (0.0438) | (0.0396) | (0.0236) | (0.0223) |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank×Time Controls | | | | Yes | Yes | Yes |
| Local Area×Time FE | | | | | Yes | Yes |
| Local Area×Bank FE | | | | | | Yes |
| Observations | 252,547 | 149,709 | 149,709 | 149,709 | 149,709 | 149,709 |
| R ² | 0.081 | 0.059 | 0.092 | 0.306 | 0.581 | 0.607 |

The dependent variables are: in Panel A ($\log(\text{Loan Value}_{i,l,b,t})$), the (log of the) value of the mortgage granted by lender b to borrower i in quarter t in local area l ; in Panel B ($\text{LTV}_{i,l,b,t}$), the loan-to-value ratio (in percentage points) of the mortgage granted by lender b to borrower i in quarter t in local area l ; and in Panel C ($\log(\text{Income}_{i,l,b,t})$), the (log of the) annual gross income of borrower(s) i associated to a mortgage granted by lender b in quarter t in local area l . The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. The sample is restricted to mortgages with a loan-to-income ratio (LTI) ≥ 4.5 . Column (1) includes all observations; Columns (2)-(6) include only observations for which the full set of control variables contain non-missing values. Post_{*t*} is a binary variable that equals 1 for quarters after 2014Q3, 0 before. Constrained_{*b*} is a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2), 0 otherwise. All regressions are estimated using ordinary least squares. Borrower, loan, and bank controls, as well as fixed-effects are either included (“Yes”) or not included. Online Appendix Section B shows the definition of the main controls. Post_{*t*} and Constrained_{*b*} are included as standalone variables when they are not absorbed by fixed-effects; we do not report their coefficients. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 4: Loan level: Income and age quintiles in high LTI loans

| (A) Income Quintiles | | | | | |
|---|---|----------------------|---------------------|-----------------------|-----------------------|
| VARIABLES | (1) | (2) | (3) | (4) | (5) |
| | ← Dummy for income quintiles ($\mathbb{D}(\text{Inc} = j)$) → | | | | |
| | I | II | III | IV | V |
| $\text{Post}_t \times \text{Constrained}_b$ | -0.0520*** (0.0153) | -0.0154* (0.0089) | -0.0133 (0.0093) | 0.0235*** (0.0086) | 0.0572*** (0.0150) |
| Borrower, Loan, Bank Controls | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Bank FE | Yes | Yes | Yes | Yes | Yes |
| Observations | 149,709 | 149,709 | 149,709 | 149,709 | 149,709 |
| R ² | 0.366 | 0.102 | 0.110 | 0.168 | 0.236 |

| (B) Age Quintiles | | | | | |
|---|--|---------------------|---------------------|--------------------|--------------------|
| VARIABLES | (1) | (2) | (3) | (4) | (5) |
| | ← Dummy for age quintiles ($\mathbb{D}(\text{Age} = j)$) → | | | | |
| | I | II | III | IV | V |
| $\text{Post}_t \times \text{Constrained}_b$ | -0.0058 (0.0072) | -0.0012 (0.0069) | -0.0013 (0.0056) | 0.0039 (0.0054) | 0.0044 (0.0039) |
| Borrower, Loan, Bank Controls | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Bank FE | Yes | Yes | Yes | Yes | Yes |
| Observations | 149,709 | 149,709 | 149,709 | 149,709 | 149,709 |
| R ² | 0.343 | 0.121 | 0.103 | 0.174 | 0.389 |

The dependent variables are binary variables indicating whether borrower(s) i associated to a mortgage granted by lender b in quarter t in local area l belongs to a particular quintile (1) or not (0). The dependent variables in Panel A (B) focus on income (age) quintiles: in each column, the dependent variable is a binary variable equal to 1 if the income (age) of the borrower(s) is in that particular quintile (with I being the bottom and V the top quintile) of the income (age) distribution. The quintiles are constructed using the sample from the pre-policy period. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3, and it is restricted to mortgages with a loan-to-income ratio (LTI) ≥ 4.5 . Post_t is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. Constrained_b is a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages before the announcement of the policy (2013Q3-2014Q2), 0 otherwise. All regressions are estimated using ordinary least squares. All regressions include borrower, loan, and bank controls, and local area \times quarter and local area \times bank fixed-effects. Online Appendix Section B shows the definition of the main controls. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 5: Loan level: Substitution effects in lending to low-income borrowers

(A) Constrained lenders across LTI groups

| VARIABLES | (1) | (2) | (3) | (4) | (5) |
|---|--|--------------------------|--------------------------|-------------------------|------------------------|
| | ← Dummy for low-income borrowers ($\mathbb{D}(\text{Inc}=\text{I})$) → | | | | |
| Sub-sample | LTI \in (0,2.09) | LTI \in [2.09,2.87) | LTI \in [2.87,3.57) | LTI \in [3.57,4.5) | LTI \geq 4.5 |
| $\text{Post}_t \times \text{Constrained}_b$ | 0.0136** (0.0056) | 0.0236*** (0.0072) | 0.0244*** (0.0083) | 0.0321*** (0.0079) | -0.0520*** (0.0153) |
| Borrower, Loan, Bank Controls | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Bank FE | Yes | Yes | Yes | Yes | Yes |
| Observations | 404,337 | 417,771 | 427,980 | 445,284 | 149,709 |
| R^2 | 0.187 | 0.213 | 0.220 | 0.258 | 0.366 |

(B) Unconstrained lenders across LTI groups in high-exposure local areas

| VARIABLES | (1) | (2) | (3) | (4) | (5) |
|--|--|----------------------------|---------------------------|---------------------------|--------------------------|
| | ← Dummy for low-income borrowers ($\mathbb{D}(\text{Inc}=\text{I})$) → | | | | |
| Sub-sample | LTI \in (0,2.09) | LTI \in [2.09,2.87) | LTI \in [2.87,3.57) | LTI \in [3.57,4.5) | LTI \geq 4.5 |
| $\text{Post}_t \times \text{Unconst.}_b \times \text{ConstrShare}_l$ | -0.000417** (0.000204) | -0.000831*** (0.000257) | -0.000810** (0.000355) | -0.00102*** (0.000306) | 0.00147*** (0.000499) |
| Borrower, Loan, Bank Controls | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | Yes | Yes | Yes | Yes | Yes |
| Local Area \times Bank FE | Yes | Yes | Yes | Yes | Yes |
| Observations | 404,337 | 417,771 | 427,980 | 445,284 | 149,709 |
| R^2 | 0.187 | 0.213 | 0.220 | 0.258 | 0.366 |

The dependent variable, $\mathbb{D}(\text{Inc}=\text{I})$, is a binary variable indicating whether the income of borrower(s) i associated to the mortgage granted by lender b in quarter t in local area l is in the bottom quintile of the income distribution among borrowers. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. We calculate the quartiles of the loan-to-income distribution for $\text{LTI} < 4.5$ in the pre-policy period. In each column (1)-(4), we restrict the sample to mortgages with an LTI inside that particular quartile. In Column (5), we restrict the sample to mortgages with $\text{LTI} \geq 4.5$. Post_t is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. Constrained_b (Unconst._b) is a binary variable that equals 1 if lender b is at or very close to (below) the limit of 15% share of high LTI mortgages before the announcement of the policy (2013Q3-2014Q2), 0 otherwise. ConstrShare_l is a variable that equals the total share of lending by constrained lenders in local area l in the years before the announcement of the policy. All regressions are estimated using ordinary least squares. All regressions include borrower, loan, and bank controls, local area \times quarter and local area \times bank fixed-effects. Online Appendix Section B shows the definition of the main controls. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 6: Local-area–income level: Overall lending to low-income borrowers

| (A) Total number of loans (extensive margin) | | | | | |
|--|--|-------------------------|-------------------------|-------------------------|-------------------------|
| VARIABLES | (1) | (2) | (3) | (4) | (5) |
| | $\leftarrow \log(\# \text{ of loans}) \rightarrow$ | | | | |
| $\text{Post}_t \times \text{ConstrShare}_l \times \mathbb{D}(\text{Inc}=\text{I})_j$ | -0.0125*** (0.00335) | -0.0135*** (0.00271) | -0.0137*** (0.00279) | -0.0137*** (0.00277) | -0.0137*** (0.00479) |
| Lower-level interactions | Yes | Yes | Yes | - | - |
| Local Area \times Income Quintile FE | | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | | | Yes | Yes | Yes |
| Time \times Income Quintile FE | | | | Yes | Yes |
| Observations | 31,043 | 31,043 | 31,043 | 31,043 | 31,043 |
| R ² | 0.029 | 0.869 | 0.896 | 0.902 | 0.902 |

| (B) Total £ value of loans (intensive margin) | | | | | |
|--|---|-------------------------|-------------------------|-------------------------|-------------------------|
| VARIABLES | (1) | (2) | (3) | (4) | (5) |
| | $\leftarrow \log(\pounds \text{ of loans}) \rightarrow$ | | | | |
| $\text{Post}_t \times \text{ConstrShare}_l \times \mathbb{D}(\text{Inc}=\text{I})_j$ | -0.0157*** (0.00351) | -0.0171*** (0.00257) | -0.0173*** (0.00265) | -0.0173*** (0.00262) | -0.0173*** (0.00462) |
| Lower-level interactions | Yes | Yes | Yes | - | - |
| Local Area \times Income Quintile FE | | Yes | Yes | Yes | Yes |
| Local Area \times Time FE | | | Yes | Yes | Yes |
| Time \times Income Quintile FE | | | | Yes | Yes |
| Observations | 31,043 | 31,043 | 31,043 | 31,043 | 31,043 |
| R ² | 0.141 | 0.838 | 0.857 | 0.861 | 0.861 |

The dependent variable in Panel A is the logarithm of the total number of mortgages granted in quarter t in local area l to borrowers in the income quintile j . The dependent variable in Panel B is the logarithm of the total value (in GBP) of mortgages granted in quarter t in local area l to borrowers in the income quintile j . The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. Post_t is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. ConstrShare_l is a variable that equals the total share of lending by constrained lenders in local area l in the years before the announcement of the policy. $\mathbb{D}(\text{Inc}=\text{I})_j$ is a binary variable that equals 1 for the bottom income quintile, 0 for the rest. All regressions are estimated using ordinary least squares. Fixed-effects and lower-level interactions are either included (“Y”), spanned by other fixed-effects (“-”), or not included. Robust standard errors clustered at local-area (columns (1) to (4)) and at LAU2 (column 5) level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 7: Local-area level: House price growth in affected local areas

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|---|---------------------|---------------------|---------------------|--------------------|--------------------|--------------------|------------------|------------------|------------------|
| | 25th percentile | | | 50th percentile | | | 75th percentile | | |
| BrexitRef _{<i>t</i>} × Constr _{<i>l</i>} × LowInc _{<i>l</i>} | 0.028*** (0.008) | 0.027*** (0.007) | 0.026*** (0.007) | 0.017** (0.008) | 0.018** (0.007) | 0.017** (0.007) | 0.004 (0.007) | 0.004 (0.007) | 0.002 (0.007) |
| Post _{<i>t</i>} × Constr _{<i>l</i>} × LowInc _{<i>l</i>} | -0.015** (0.007) | -0.012* (0.007) | -0.014* (0.007) | -0.003 (0.007) | -0.001 (0.007) | -0.001 (0.007) | 0.009 (0.008) | 0.012 (0.007) | 0.010 (0.007) |
| Lower-level interactions | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Time FE | Yes | - | - | Yes | - | - | Yes | - | - |
| Region × Time FE | | Yes | Yes | | Yes | Yes | | Yes | Yes |
| Local Area FE | | | Yes | | | Yes | | | Yes |
| Observations | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 |
| R ² | 0.327 | 0.389 | 0.441 | 0.336 | 0.407 | 0.453 | 0.237 | 0.301 | 0.347 |

The dependent variable for columns (1)-(3) is the change in the logarithm of the 25th percentile price paid at local area l . The dependent variable for columns (4)-(6) is the change in the logarithm of the median price paid at local area l . The dependent variable for columns (7)-(9) is the change in the logarithm of the 75th percentile price paid at local area l . The 25th, 50th, and 75th percentiles are calculated using all house transactions during a six-month period t . The change is calculated with respect to the same period in the year before. The sample spans from 2012H2 to 2018H1. Post _{t} is a binary variable equal to 1 in the period 2014H2 - 2018H1, i.e. after the introduction of the policy, 0 before. BrexitRef _{t} is a binary variable equal to 1 in the period 2016H2 - 2018H1, i.e. after the Brexit referendum, 0 before. Constr _{l} is a binary variable equal to 1 if the market share in mortgage lending by constrained lenders in the local area l is above the median before the announcement of the policy, 0 otherwise. LowInc _{l} is a binary variable equal to 1 if the local area is in the top quartile in terms of share of low income borrowers before the announcement of the policy, 0 otherwise. All regressions are estimated using ordinary least squares. Fixed-effects and lower level interactions are included (“Yes”), spanned by other fixed-effects (“-”), or not included. The full set of coefficients for lower level interactions are reported in Table A.6. Robust standard errors are clustered at local-area level and are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 8: Local-area level: Mortgage non-performance rates in affected local areas

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
|--|--------------------|---------------------|---------------------|----------------------|---------------------|---------------------|
| | Payment shortfall | | | Forbearance | | |
| BrexitRef _t × Constr _l × LowInc _l | -0.057* (0.031) | -0.072** (0.029) | -0.064** (0.029) | -0.146*** (0.052) | -0.102** (0.045) | -0.096** (0.047) |
| Lower-level interactions | Yes | Yes | Yes | Yes | Yes | Yes |
| Time FE | Yes | - | - | Yes | - | - |
| Region × Time FE | | Yes | Yes | | Yes | Yes |
| Local Area FE | | | Yes | | | Yes |
| Observations | 2,361 | 2,361 | 2,361 | 2,361 | 2,361 | 2,361 |
| R ² | 0.661 | 0.667 | 0.982 | 0.627 | 0.634 | 0.975 |

The dependent variable in (1)-(3) is the percentage of mortgages with payment shortfall in local area l at the end of the half-year t . The dependent variable in (4)-(6) is the percentage of mortgages under forbearance in local area l at the end of half-year t . The sample spans from 2015H1 to 2017H2. BrexitRef_t is a binary variable equal to 1 in the period 2016H2 - 2017H2, i.e. after the Brexit referendum, 0 before. Constr_l is a binary variable equal to 1 if the market share in mortgage lending by constrained lenders in the local area l is above the median before the announcement of the policy, 0 otherwise. LowInc_l is a binary variable equal to 1 if the local area is in the top quartile in terms of share of low income borrowers before the announcement of the policy, 0 otherwise. All regressions are estimated using ordinary least squares. Fixed-effects and lower level interactions are included (“Yes”), spanned by other fixed-effects (“-”), or not included. Robust standard errors are clustered at local-area level and reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table 9: Local-area level: Changes in debt due to policy vs. orthogonal to policy

(A) House price growth

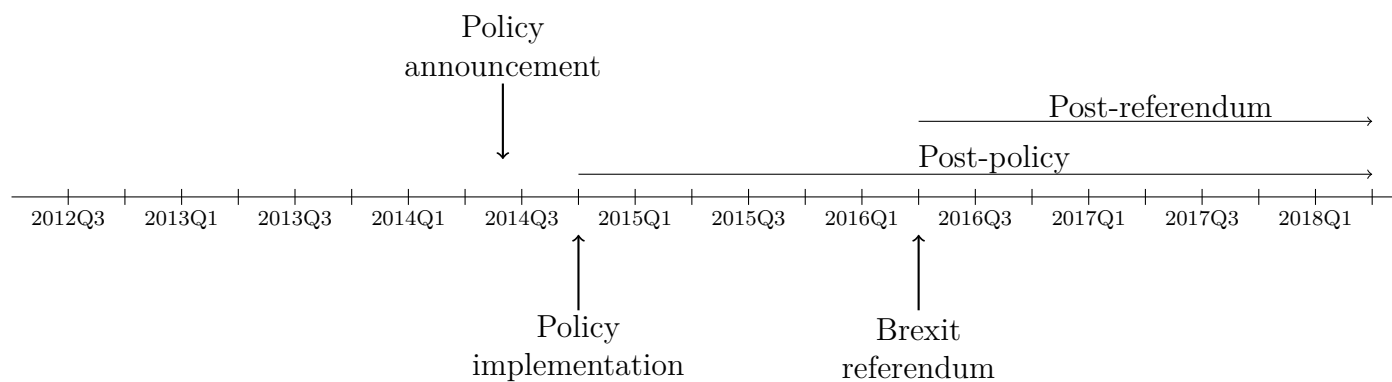
| VARIABLES | (1) 25th percentile | (2) | (3) 50th percentile | (4) | (5) 75th percentile | (6) |
|---|------------------------|-----------------------|------------------------|----------------------|------------------------|---------------------|
| BrexitRef _t × Predicted ($-\Delta$ Low-Income Debt _t) | 0.0099*** (0.0024) | 0.0097*** (0.0024) | 0.0090*** (0.0023) | 0.0089** (0.0023) | 0.0038* (0.0021) | 0.0036* (0.0021) |
| BrexitRef _t × Residual ($-\Delta$ Low-Income Debt _t) | -0.0001 (0.0009) | -0.0001 (0.0008) | -0.0006 (0.0008) | -0.0006 (0.0008) | -0.0009 (0.0008) | -0.0009 (0.0008) |
| Region × Time FE | Yes | Yes | Yes | Yes | Yes | Yes |
| Local Area FE | | Yes | | Yes | | Yes |
| Observations | 2,776 | 2,776 | 2,776 | 2,776 | 2,776 | 2,776 |
| R ² | 0.366 | 0.457 | 0.371 | 0.456 | 0.244 | 0.326 |

(B) Mortgage non-performance rates

| VARIABLES | (1) Payment shortfall | (2) | (3) Forbearance | (4) |
|---|--------------------------|------------------------|------------------------|------------------------|
| BrexitRef _t × Predicted ($-\Delta$ Low-Income Debt _t) | -0.0406*** (0.0086) | -0.0375*** (0.0089) | -0.0607*** (0.0123) | -0.0587*** (0.0129) |
| BrexitRef _t × Residual ($-\Delta$ Low-Income Debt _t) | -0.0011 (0.0037) | -0.0006 (0.0036) | -0.0010 (0.0054) | -0.0033 (0.0060) |
| Region × Time FE | Yes | Yes | Yes | Yes |
| Local Area FE | | Yes | | Yes |
| Observations | 2,361 | 2,361 | 2,361 | 2,361 |
| R ² | 0.634 | 0.983 | 0.615 | 0.979 |

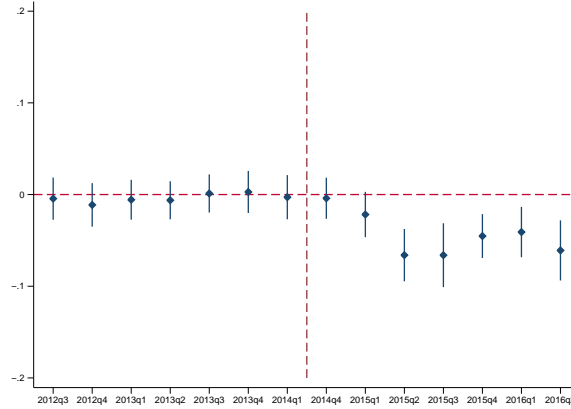
In Panel A, the dependent variable is the change in the logarithm of the 25th percentile (columns 1-2), median (c. 3-4), or 75th percentile (c. 5-6) price paid at local area l . The percentiles are calculated using all house transactions during a six-month period t . The change is calculated with respect to the same period in the year before. The sample spans from 2012H2 to 2018H1. In Panel B, the dependent variable is the percentage of mortgages with payment shortfall (c. 1-2) and the percentage of mortgages under forbearance (c. 3-4) in local area l at the end of half-year t . The sample spans from 2015H1 to 2017H2. BrexitRef_t is a binary variable equal to 1 from 2016H2 onwards, i.e. after the Brexit referendum, 0 before. Predicted Δ Low-Income Debt_t and Residual Δ Low-Income Debt_t are constructed from an OLS regression where we estimate Δ Low-Income Debt_t (difference in total lending to low-income borrowers from 2014Q2 to 2016Q2 divided by total lending in 2014Q2 at local area l) on Constr_t × LowInc_t; we multiply these two controls by -1 for better comparison with the previous tables. All regressions are estimated using ordinary least squares. Fixed-effects are included (“Yes”) or not included. Robust standard errors are clustered at local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Figure 2: Timeline



The above figure shows the important dates for our analysis. The macroprudential policy was announced in 26th of June 2014, and it was implemented from 1st of October 2014. The Brexit referendum in the UK took place on the 23rd of June 2016. For the analysis on the impact of the macroprudential policy on credit availability, we focus on all mortgages issued between 2012Q3 and 2016Q2. We drop the mortgages issued in 2014Q3 since the policy had already been announced but not yet implemented during that quarter; we stop in 2016Q2 to avoid the possibility that our results on credit availability reflect some of the effects of the Brexit referendum. For the analysis on house price growth (non-performing mortgages), we study the period between 2012Q4-2018Q2 (2015Q2-2017Q4). We study the period after the Brexit referendum to understand whether the changes in credit availability have beneficial effects on house price dynamics and mortgage defaults following the referendum.

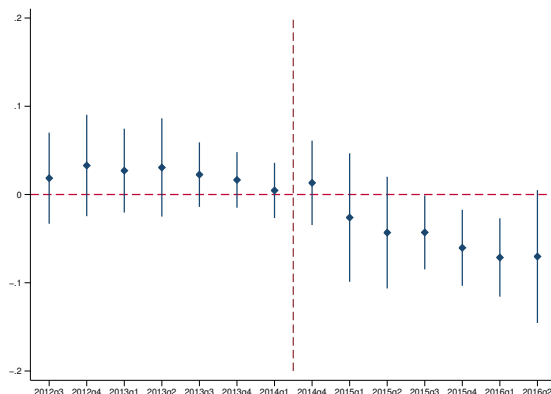
Figure 3: Loan level: Share of loans with high loan-to-income ratios ($LTI \geq 4.5$)



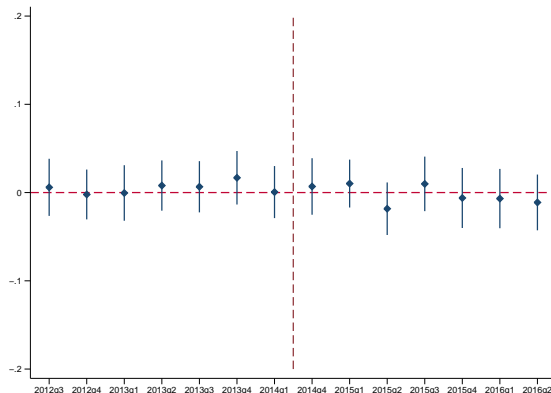
The figure above shows the estimated coefficients of the interaction of Constrained_b , a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2), 0 otherwise, and the full set of time(quarter)-dummies (excluding 2014Q2, the reference time-period). The dependent variable is $\mathbb{D}(LTI \geq 4.5)_{i,l,b,t}$, which is a binary variable that equals 1 if the mortgage granted by lender b to borrower i in quarter t in local area l has a loan-to-income ratio (LTI) equal or above 4.5. The bands represent the 95% confidence interval based on robust standard errors double-clustered at the bank \times quarter and local-area level. The vertical dashed line shows the date of the introduction of the policy. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. The regression is estimated using ordinary least squares, and includes borrower, loan, and bank controls, as well as Local Area \times Quarter and Local Area \times Bank fixed-effects. The definition of the main independent variables can be found in Online Appendix Section B.

Figure 4: Loan level: Share of low-income and young borrowers in high LTI loans

(A) Low-income borrowers (regression results in Table 4 (A), column 1)



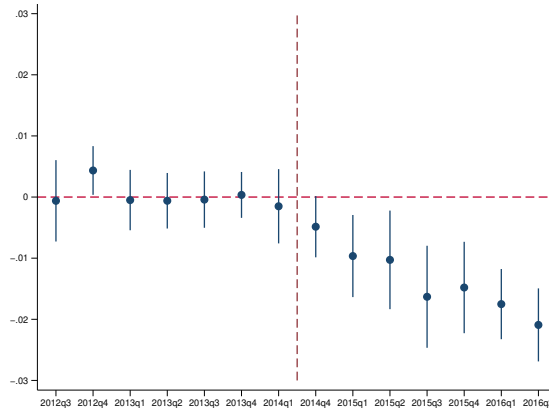
(B) Young borrowers (regression results in Table 4 (B), column 1)



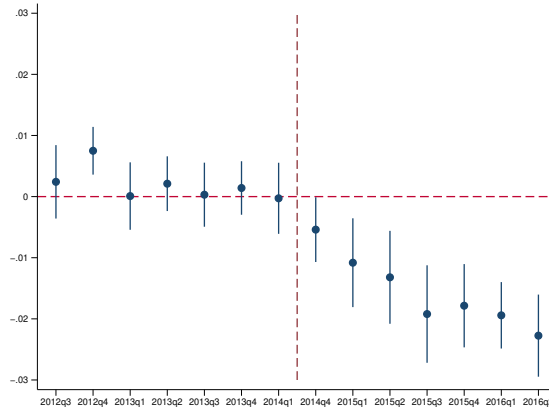
The figures above show the estimated coefficients of the interaction of Constrained_b , a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2), 0 otherwise, and the full set of time(quarter)-dummies (excluding 2014Q2, the reference time-period). The dependent variable in (A) is $\mathbb{D}(\text{Inc}=\text{I})$, which is a binary variable that equals 1 if the income of the borrower is in the bottom quintile of the income distribution, 0 otherwise. The dependent variable in (B) is $\mathbb{D}(\text{Age}=\text{I})$, which is a binary variable that equals 1 if the age of the main borrower is in the bottom quintile of the age distribution, 0 otherwise. The bands represent the 95% confidence interval based on robust standard errors double-clustered at the bank \times quarter and local-area level. The vertical dashed line shows the date of the introduction of the policy. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3, and it only includes mortgages with an $\text{LTI} \geq 4.5$. The regressions are estimated using ordinary least squares, and include borrower, loan, and bank controls, as well as Local Area \times Quarter and Local Area \times Bank fixed-effects. The definition of the main independent variables can be found in Online Appendix Section B.

Figure 5: Local-area–income level: Overall lending to low-income borrowers

(A) Extensive margin effects (regression results in Table 6 (A), column 4)

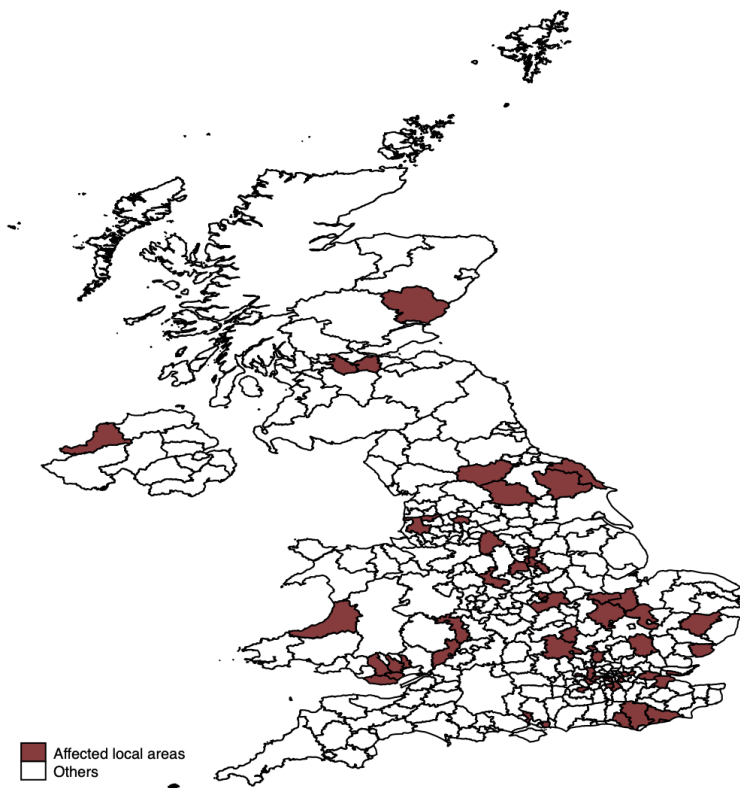


(B) Intensive margin effects (regression results in Table 6 (B), column 4)



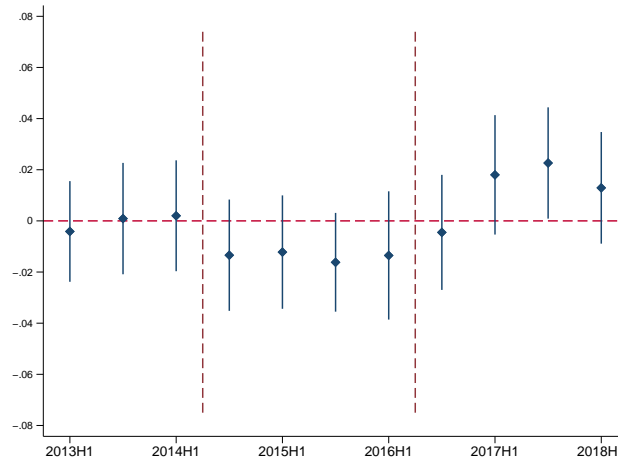
The figures above show the estimated coefficients of the interaction of ConstrShare_l (a variable that equals the total share of lending by constrained lenders in local area l in the years before the announcement of the policy), $\mathbb{D}(\text{Inc}=\text{I})$ (a binary variable that equals 1 for the bottom income quintile, 0 for the rest), and the full set of time(quarter)-dummies (excluding 2014Q2, the reference time-period). The dependent variable in Panel A is the logarithm of the total number of mortgages granted in quarter t in local area l to borrowers in the income quintile j . The dependent variable in Panel B is the logarithm of the total value (in GBP) of mortgages granted in quarter t in local area l to borrowers in the income quintile j . The bands represent the 95% confidence interval based on robust standard errors clustered at local-area level. The vertical dashed line shows the date of the introduction of the policy. The sample period is from 2012Q3 to 2016Q2. The regressions are estimated using ordinary least squares and include Local Area \times Time, Local Area \times Income Quintile, and Time \times Income Quintile fixed-effects.

Figure 6: Local-areas more affected by the policy



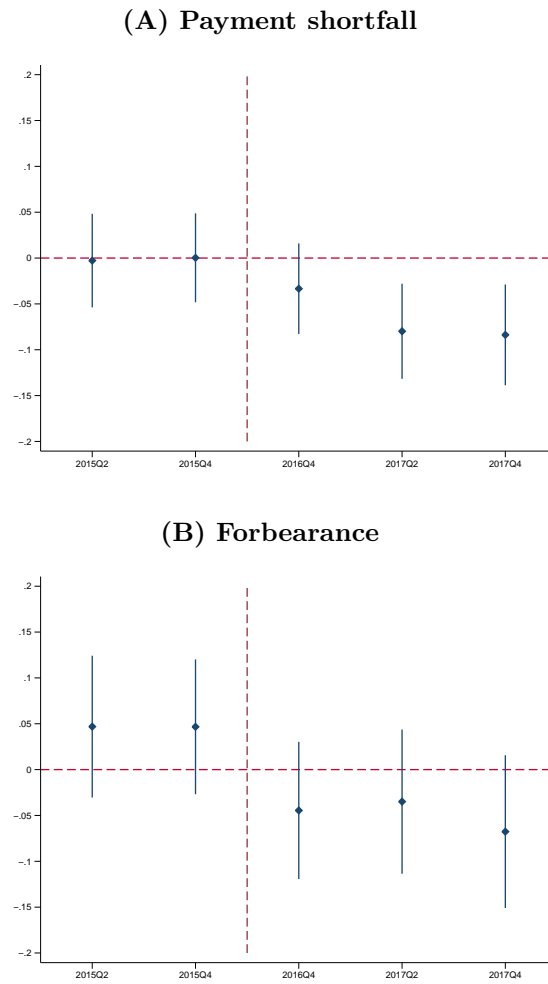
The figure above shows the spatial distribution of the local areas more affected by the policy. These are the local areas with an ex-ante higher exposure to constrained lenders *and* a higher proportion of low-income borrowers. The affected local areas are indicated by the interaction between $Constr_t$, which indicates local areas where the share of constrained lenders before the policy is greater than the median share, and $LowInc_t$, which indicates local areas where the share of loans received by low-income borrowers is above the 75th percentile. UK local areas (at the official Local Administrative Unit Level 1) are equivalent to counties/boroughs.

Figure 7: Local-area level: House price growth in affected local areas



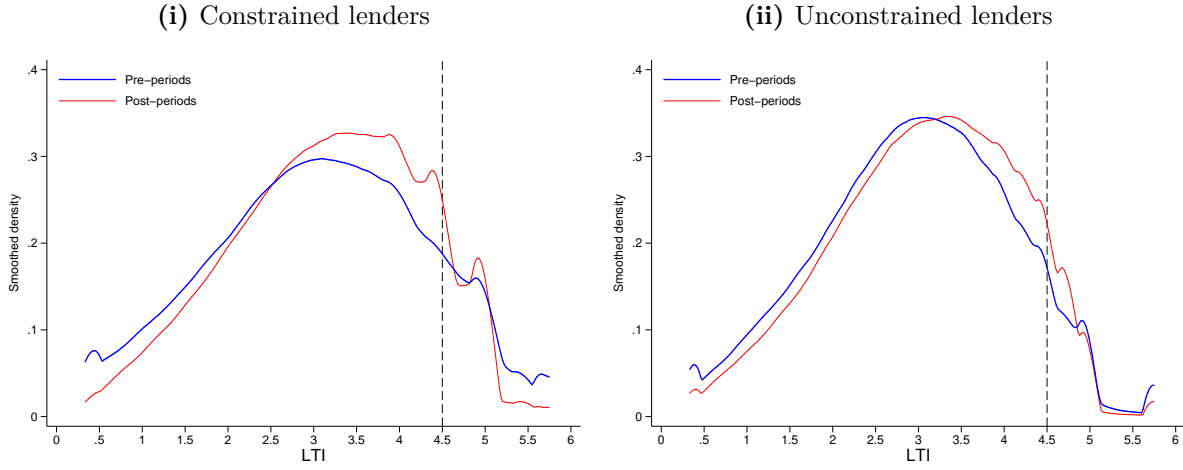
The figures above show the estimated coefficients of the interaction of $Constr_l$ (a binary variable equal to 1 if the market share in mortgage lending by constrained lenders in the local authority l is above the median before the announcement of the policy, 0 otherwise), $LowInc_l$ (binary variable equal to 1 if the local authority is in the top quartile in terms of share of low income borrowers before the announcement of the policy, 0 otherwise), and the full set of time(half-year)-dummies (excluding 2012H2, the reference time-period). The dependent variable is the annual change in the logarithm of the 25th percentile house price in local authority l in half-year t . The bands represent the 90% confidence interval based on robust standard errors clustered at local-area level. The vertical dashed lines show the date of the introduction of the policy and the Brexit referendum. The sample period is from 2012H2 to 2018H1. The regression is estimated using ordinary least squares and includes Local Area and Region \times Time fixed-effects.

Figure 8: Local-area level: Mortgage non-performance in affected local areas



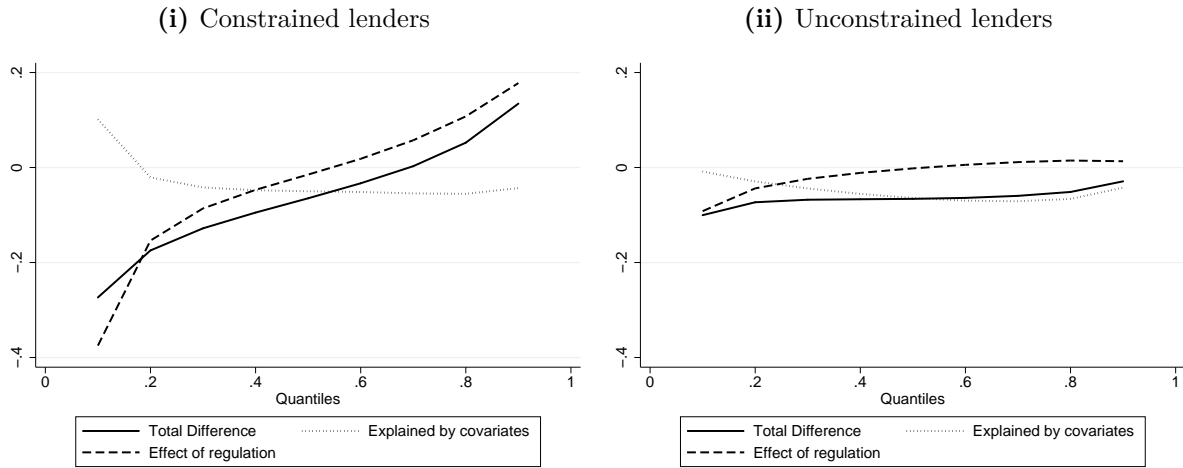
The figures above show the estimated coefficients of the interaction of $Constr_l$ (a binary variable equal to 1 if the market share in mortgage lending by constrained lenders in the local authority l is above the median before the announcement of the policy, 0 otherwise), $LowInc_l$ (binary variable equal to 1 if the local authority is in the top quartile in terms of share of low income borrowers before the announcement of the policy, 0 otherwise), and the full set of time(half-year)-dummies (Q2 refers to the end of June of the corresponding year; Q4 refers to the end of the corresponding year). 2016Q2 is excluded as the reference time-period. The dependent variable in Panel (A) is the percentage of mortgages with payment shortfall in local area l at the end of the half-year t . The dependent variable in Panel (B) is the percentage of mortgages under forbearance in local area l at the end of the half-year t . The bands represent the 95% confidence interval based on robust standard errors clustered at local-area level. The vertical dashed line shows the date of the introduction of the Brexit referendum. The sample period is from 2015Q2 to 2017Q4. The regressions are estimated using ordinary least squares and include Local Area and Region \times Time fixed-effects.

Figure 9 (A): Distribution of LTI for constrained and unconstrained lenders



The above figures show a histogram of the loan-to-income ratios of mortgages issued by the constrained (left panel) and unconstrained lenders (right panel) in periods before (2012Q3-2014Q2; Pre-periods) and after (2014Q4-2016Q2; Post-periods) the introduction of the limit on high LTI lending by UK regulators. The underlying data is the universe of residential mortgages issued in the UK from PSD001.

Figure 9 (B): Oaxaca-Blinder decomposition for constrained and unconstrained lenders



The above figure shows the Oaxaca-Blinder decomposition for the change in the LTI distribution of mortgages issued by constrained (left panel) and unconstrained lenders (right panel) in Pre_t and $Post_t$ periods. The decomposition is based on a counterfactual distribution estimated using the technique described in Chernozhukov et al. (2013). The technique uses a set of covariates for all mortgages issued in $Post_t$ periods and estimates a counterfactual distribution based on the likelihood of obtaining a certain LTI for the set of covariates in Pre_t periods. The difference between the quantiles of the $Post_t$ period LTI distribution and the counterfactual distribution is attributable to the regulation (dashed line); the difference between the quantiles of the counterfactual and Pre_t period LTI distribution is explained by changes in the distribution of covariates in the mortgages issued by the lender in Pre_t and $Post_t$ periods (dotted line).

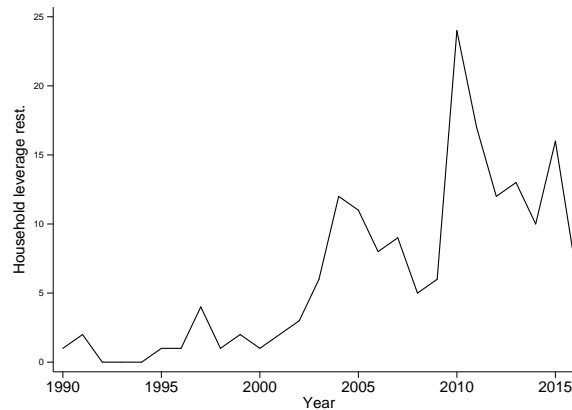
INTERNET APPENDIX

Macroprudential Policy, Mortgage Cycles and Distributional Effects: Evidence from the UK

A Restrictions on household leverage introduced during 1990-2016

Figure A.1 shows the number of regulations introduced annually world-wide during 1990-2016 aimed at restricting household leverage by setting limits on loan-to-value (LTV) and debt-service-to-income (DSTI) ratios. The figure shows that most interventions to restrict household leverage have taken place after the 2008 financial crisis. The restriction on high LTI lending introduced in the UK was one of 10 regulations on household leverage introduced worldwide in 2014.

Figure A.1: Number of regulations restricting household leverage during 1990-2016



The above figure shows the number of regulations aimed at restricting household leverage introduced world-wide during 1990-2016. The underlying data used for the figure is obtained from the Integrated Macroprudential Policy (iMaPP) database collated by the IMF ([link](#)). The figure for each year sums the number of instances of regulations restricting loan-to-value and debt-service-to-income ratios, indicated by two separate variables in the database.

B Description of loan level and bank-time level controls

Table A.1: Loan level controls

| Variable | Description |
|----------------|---|
| Borrower-type | Includes <i>first-time-buyers, home-movers, refinancers</i> |
| Ratetype | <i>Fixed</i> or <i>floating</i> |
| Dealtype | Duration of initial teaser rates |
| Employment | Employment status of the primary borrower |
| Age | Age of the primary borrower |
| Mortgage term | Duration of payment period |
| Repayment | Type of repayment, for e.g. <i>interest only</i> |
| Loan value | Total loan amount issued to the borrower |
| Property value | £ value of the property evaluated by the lender |

Source: PSD001

Table A.2: Bank-time level controls

| Variable | Description |
|---------------------------|--|
| Size | Logarithm of total assets |
| Liquidity ratio | Cash and government bonds over total assets |
| Core funding ratio | Retail deposits over total assets |
| Return on assets | Profits over total assets |
| CET1 ratio | Common Equity Tier 1 capital over risk-weighted assets |
| Share of household credit | Household lending over total lending |

Source: Statistical and regulatory data reported by financial institutions to the Bank of England (BT/PL/AL forms) and PRA.

Table A.3: Summary statistics for constrained and unconstrained lenders

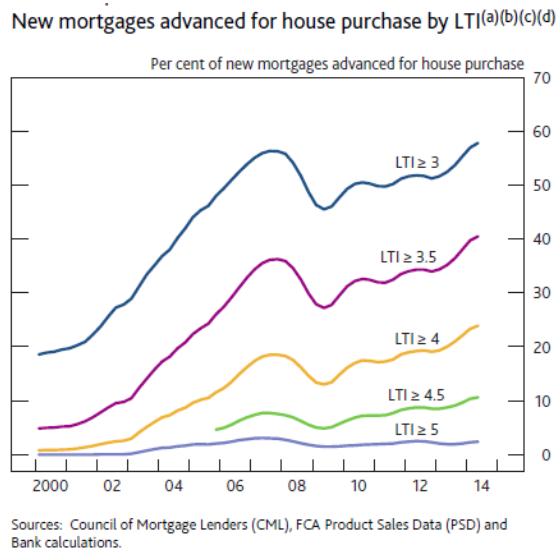
| Statistic | $\mu_{Unconst.}$ | $\mu_{Constrained}$ | t | p |
|-------------------------------|------------------|---------------------|--------|------|
| Size | 9.94 | 10.54 | -1.07 | 0.29 |
| Liquidity (%) | 9.95 | 9.33 | 0.47 | 0.64 |
| Core funding (%) | 73.3 | 66.3 | 1.31 | 0.19 |
| Return on assets (%) | 0.14 | 0.12 | 0.38 | 0.7 |
| CET1 Ratio (%) | 16.14 | 18.39 | -1.4 | 0.17 |
| Share of household credit (%) | 71.88 | 65.36 | 1.04 | 0.3 |
| Share of high LTI loans | 0.05 | 0.18 | -10.32 | 0.00 |

The above table shows the difference in the averages of the bank-level control variables for constrained and unconstrained lenders based on data at a quarterly level for the period 2013Q3 - 2014Q2. Column 1 shows the averages for unconstrained banks; column 2 shows the averages for constrained banks; column 3 provides the t-statistic for the null hypothesis of equal means, and column 4 provides the corresponding p-value. A constrained lender is defined as a lender that is at or very close to the 15% limit on the share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2). The definition of these variables can be found in Online Appendix Table A.2.

C Lending standards prior to the introduction of the 2014 regulation

Figure A.2 shares evidence on changes in lending standards in the UK mortgage market in the run up to the introduction of the 2014 restrictions on high LTI lending. The figure plots the share of mortgages above a given level of LTI, such as the share of mortgages with $LTI \geq 4$. The share of mortgages with these relatively high levels of borrower leverage had crept up back to the pre-crisis levels by early-2014.

Figure A.2: Share of different LTI groups in new mortgages issued before the introduction of the limit on high LTI lending



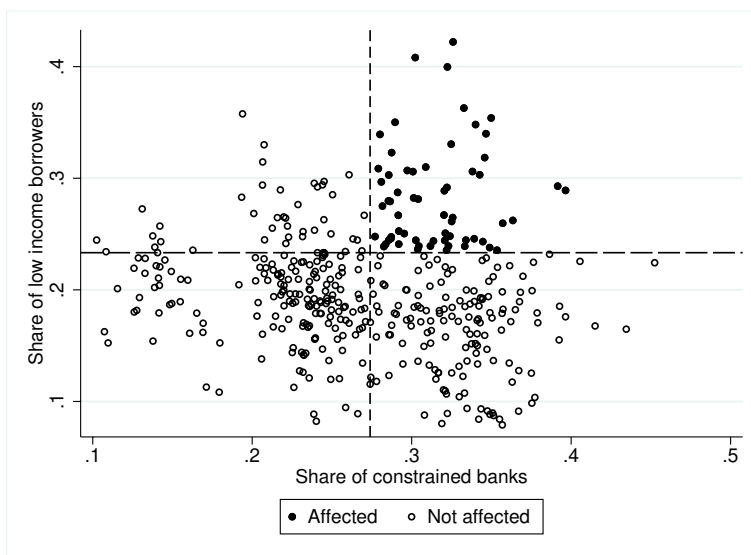
The above figure shows the share of different LTI groups in the new mortgages issued at a quarterly level between 2013Q2-2014Q1. During this period, around 10% of lending for house purchase was extended at an LTI at or above four and a half times income; this compares to 6.5% in the immediate pre-crisis period, 2005-07.

D Cross-variation across UK local-areas of exposure to constrained lenders and share of low-income borrowers

Section 4.4 show that (1) local-areas more affected by the regulation experienced a cooling down of house prices after the regulation, and (2) experienced relatively lower house price declines and mortgage non-performance rates during the house price correction in the UK following the so-called Brexit referendum. “Affected” local-areas are those with a higher share of constrained lenders and a higher share of low-income borrowers.

Figure A.3 shows the cross-variation between the share of constrained lenders and the share of mortgages (in Pre_t) periods received by low-income borrowers. Figure shows that there is a sufficient cross-variation between the two variables for identification using our empirical strategy; the “affected” local-areas are reflected as shaded dots in the figure.

Figure A.3: Cross-variation across UK local-areas of exposure to constrained lenders and low-income borrowers



The above figure shows the cross-variation between the share of constrained banking groups and the share of borrowers in the bottom income quintile across local areas in the United Kingdom. Share of borrowers in the bottom income quintile is based on thresholds defined using the income distribution of mortgages issues at a more aggregated NUTS2 level in periods before the introduction of the regulation (see Section 4.3). The above figure shows that the share of constrained banks and the share of low-income borrowers is not highly correlated. We study house price growth and mortgage arrears in local areas more exposed to constrained banks and with greater share of low-income borrowers in the aftermath of the Brexit referendum as compared to other local areas in Section 4.4.

E Supplementary results

Table A.4: Results with a continuous explanatory variable

(A) Share of loans with loan-to-income ratios ≥ 4.5 ; robustness result for Table 2

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
|--|--|----------------------|-----------------------|-----------------------|-----------------------|-----------------------|
| | ← Dummy for high LTI loans ($\mathbb{D}(\text{LTI} \geq 4.5)$) → | | | | | |
| Post _{<i>t</i>} × Share _{<i>b</i>} | -0.414*** (0.0935) | -0.347*** (0.109) | -0.329*** (0.0969) | -0.451*** (0.0828) | -0.473*** (0.0748) | -0.476*** (0.0806) |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank × time Controls | | | | Yes | Yes | Yes |
| Local Area × Time FE | | | | | Yes | Yes |
| Local Area × Bank FE | | | | | | Yes |
| Observations | 2,907,529 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 |
| R ² | 0.013 | 0.015 | 0.016 | 0.068 | 0.106 | 0.121 |

(B) Composition of high LTI loans; robustness result for Table 3 and Table 4

| VARIABLES | (1) | (2) | (3) | (4) | (5) |
|--|---------------------|-------------------|---------------------|----------------------------|----------------------------|
| | log(Loan Size) | LTV | log(Income) | $\mathbb{D}(\text{Inc}=I)$ | $\mathbb{D}(\text{Age}=I)$ |
| Post _{<i>t</i>} × Share _{<i>b</i>} | 0.862*** (0.231) | 10.80† (6.766) | 1.049*** (0.261) | -0.308* (0.161) | -0.0307 (0.0709) |
| Controls and FE | Yes | Yes | Yes | Yes | Yes |
| Observations | 149,709 | 149,705 | 149,709 | 149,709 | 149,709 |
| R ² | 0.627 | 0.376 | 0.607 | 0.366 | 0.578 |

The dependent variable in Panel (A) is $\mathbb{D}(\text{LTI} \geq 4.5)_{i,l,b,t}$, a binary variable that equals 1 if the mortgage granted by lender b to borrower i in quarter t in local area l has an LTI ≥ 4.5 , 0 otherwise. The dependent variables in Panel (B) are the following: (1) the (log of the) mortgage granted by lender b to borrower i in quarter t in local area l ; (2) the LTV of the mortgage granted by lender b to borrower i in quarter t in local area l ; (3) the logarithm of the annual gross income of borrower(s) i in the mortgage granted by lender b in quarter t in local area l ; (4) a binary variable equal to 1 if the income of borrower i is in the bottom quintile of the distribution, 0 otherwise; and (5) a binary variable equal to 1 if the age of borrower i is in the bottom quintile of the distribution, 0 otherwise. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. Panel (A) includes all mortgages. Panel (B) include only LTI ≥ 4.5 . Post_{*t*} is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. Share_{*b*} is the share of high LTI mortgages held by bank b in the quarters before the announcement of the policy. All regressions are estimated using ordinary least squares. Borrower, loan, and bank controls, as well as fixed-effects are either included (“Yes”) or not included. Online Appendix Section B shows the definition of the main controls. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level; †: significant at 11%.

Continuous explanatory variable. The baseline specification (Equation 1) is based on classifying UK lenders into two groups, “constrained” and “unconstrained”, based on their share of high LTI loans in all mortgages issued during 2013Q3-2014Q2. In table A.4 we show that the results are robust to using a continuous explanatory variable, i.e. Share_{*b*}, instead of using a dummy for constrained lenders as done for Table 2, Table 3 and Table 4. Table A.4 (A) shows that this alternate treatment variable also highlights a contraction in the share of high LTI loans issued by the more-constrained lenders in Post_{*t*} periods. Table A.4 (B)

shows increases in the loan size, LTV and borrower income, and a drop in the proportion of low-income borrowers in the fewer high LTI loans issued by the more-constrained lenders as estimated on the alternate continuous treatment variable.

Coefficients on low-level interactions. In the main text, we do not report the coefficients on the all levels and interaction of the main independent variables due to lack of space. We reproduce here the full set of levels from Table 2 (in Table A.5) and Table 7 (in Table A.6).

Table A.5: Loan level: Share of loans with high loan-to-income ratios ($LTI \geq 4.5$)

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) |
|--|---|------------------------|------------------------|------------------------|------------------------|------------------------|
| | ← Dummy for high LTI loans ($\mathbb{D}(LTI \geq 4.5)$) → | | | | | |
| Post _t × Constrained _b | -0.0295** (0.0118) | -0.0474*** (0.0116) | -0.0394*** (0.0099) | -0.0513*** (0.0093) | -0.0566*** (0.0089) | -0.0388*** (0.0076) |
| Constrained _b | 0.0599*** (0.00934) | 0.0836*** (0.00827) | 0.0779*** (0.00759) | 0.0706*** (0.00690) | 0.0647*** (0.00620) | |
| Post _t | 0.00500 (0.00671) | 0.0219*** (0.00661) | 0.0130** (0.00518) | 0.0215*** (0.00745) | | |
| Borrower and Loan Controls | | | Yes | Yes | Yes | Yes |
| Bank × time Controls | | | | Yes | Yes | Yes |
| Local Area × Time FE | | | | | Yes | Yes |
| Local Area × Bank FE | | | | | | Yes |
| Observations | 2,907,529 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 | 1,849,952 |
| R ² | 0.005 | 0.012 | 0.014 | 0.066 | 0.103 | 0.121 |

The dependent variable in this table is $\mathbb{D}(LTI \geq 4.5)_{i,l,b,t}$, a binary variable that equals 1 if the mortgage granted by lender b to borrower i in quarter t in local area l has a loan-to-income ratio (LTI) equal or above 4.5, and 0 otherwise. The sample period is from 2012Q3 to 2016Q2, excluding 2014Q3. Column (1) includes all observations; Columns (2)-(6) include only observations for which the full set of control variables contain non-missing values. Post_t is a binary variable that equals 1 for quarters after 2014Q3, 0 for quarters before. Constrained_b is a binary variable that equals 1 if lender b is at or very close to the limit of 15% share of high LTI mortgages in the four quarters before the announcement of the policy (2013Q3-2014Q2), 0 otherwise. All regressions are estimated using ordinary least squares. Borrower, loan, and bank controls, as well as fixed-effects are either included (“Yes”) or not included. Online Appendix Section B shows the definition of the main controls. Robust standard errors double-clustered at bank-quarter and local-area level are reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.

Table A.6: Local-area level: House price growth

| VARIABLES | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
|--|----------------------|---------------------|---------------------|----------------------|--------------------|--------------------|----------------------|-------------------|-------------------|
| | 25th percentile | | | 50th percentile | | | 75th percentile | | |
| $Post_t \times Const_l \times LowInc_l$ | -0.015** (0.007) | -0.012* (0.007) | -0.014* (0.007) | -0.003 (0.007) | -0.001 (0.007) | -0.001 (0.007) | 0.009 (0.008) | 0.012 (0.007) | 0.010 (0.007) |
| $BrexitRef_t \times Const_l \times LowInc_l$ | 0.028*** (0.008) | 0.027*** (0.007) | 0.026*** (0.007) | 0.017** (0.008) | 0.018** (0.007) | 0.017** (0.007) | 0.004 (0.007) | 0.004 (0.007) | 0.002 (0.007) |
| $Const_l$ | -0.011*** (0.003) | -0.007** (0.003) | | -0.005 (0.003) | -0.002 (0.004) | | 0.002 (0.003) | 0.004 (0.004) | |
| $LowInc_l$ | -0.005 (0.004) | -0.002 (0.004) | | -0.003 (0.003) | -0.000 (0.003) | | -0.003 (0.004) | -0.000 (0.004) | |
| $Post_t \times Const_l$ | 0.038*** (0.004) | 0.016*** (0.005) | 0.017*** (0.005) | 0.027*** (0.004) | 0.007 (0.005) | 0.008 (0.005) | 0.009** (0.004) | -0.007 (0.007) | -0.007 (0.007) |
| $Post_t \times LowInc_l$ | 0.007 (0.005) | 0.003 (0.005) | 0.004 (0.005) | 0.001 (0.004) | -0.003 (0.004) | -0.003 (0.004) | -0.000 (0.005) | -0.005 (0.005) | -0.004 (0.005) |
| $BrexitRef_t \times Const_l$ | -0.039*** (0.004) | -0.008 (0.005) | -0.008 (0.005) | -0.038*** (0.004) | -0.008 (0.005) | -0.008 (0.005) | -0.025*** (0.004) | 0.002 (0.007) | 0.002 (0.007) |
| $BrexitRef_t \times LowInc_l$ | -0.004 (0.005) | -0.005 (0.005) | -0.004 (0.005) | 0.004 (0.005) | 0.003 (0.005) | 0.004 (0.005) | 0.005 (0.005) | 0.005 (0.004) | 0.005 (0.005) |
| $Const_l \times LowInc_l$ | -0.001 (0.005) | -0.002 (0.005) | | -0.003 (0.004) | -0.005 (0.004) | | -0.005 (0.005) | -0.007 (0.005) | |
| Time FE | Yes | - | - | Yes | - | - | Yes | - | - |
| Region \times Time FE | | Yes | Yes | | Yes | Yes | | Yes | Yes |
| Local Area FE | | | Yes | | | Yes | | | Yes |
| Observations | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 | 4,176 |
| R ² | 0.327 | 0.389 | 0.441 | 0.336 | 0.407 | 0.453 | 0.237 | 0.301 | 0.347 |

The dependent variable for columns (1)-(3) is the change in the logarithm of the 25th percentile price paid at local area l . The dependent variable for columns (4)-(6) is the change in the logarithm of the median price paid at local area l . The dependent variable for columns (7)-(9) is the change in the logarithm of the 75th percentile price paid at local area l . The 25th, 50th, and 75th percentiles are calculated using all house transactions during a six-month period t . The change is calculated with respect to the same period in the year before. The sample spans from 2012H2 to 2018H1. $Post_t$ is a binary variable equal to 1 in the period 2014H2 - 2018H1, i.e. after the introduction of the policy, 0 before. $BrexitRef_t$ is a binary variable equal to 1 in the period 2016H2 - 2018H1, i.e. after the Brexit referendum, 0 before. $Const_l$ is a binary variable equal to 1 if the market share in mortgage lending by constrained lenders in the local area l is above the median before the announcement of the policy, 0 otherwise. $LowInc_l$ is a binary variable equal to 1 if the local area is in the top quartile in terms of share of low income borrowers before the announcement of the policy, 0 otherwise. All regressions are estimated using ordinary least squares. Fixed-effects are included ("Yes"), spanned by other fixed-effects ("-"), or not included. Robust standard errors are clustered at local-area level and reported in parentheses. ***: Significant at 1% level; **: significant at 5% level; *: significant at 10% level.